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CMS统计分析及相关工具简介

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23/01/2024

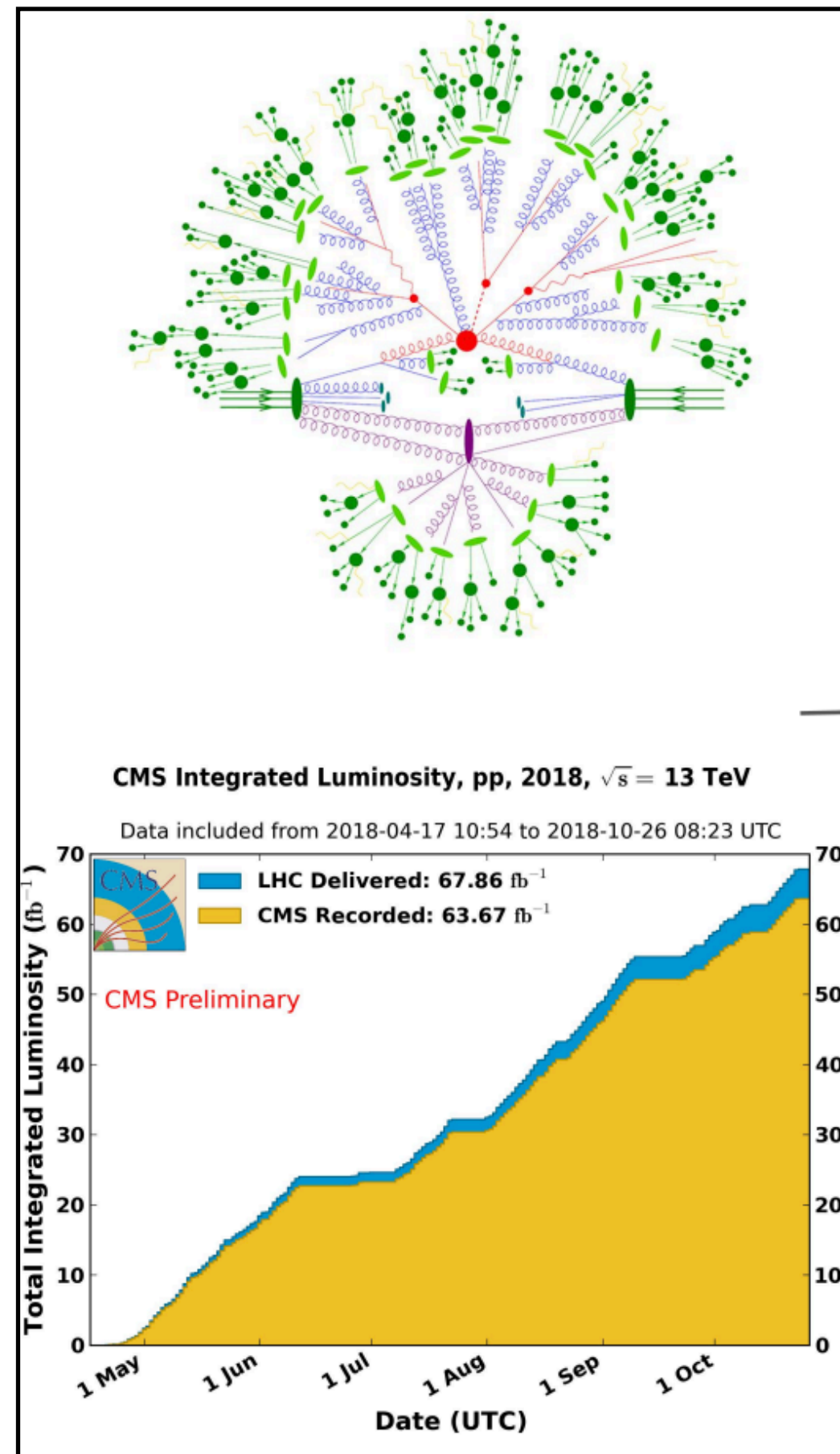
第二届中国CMS冬令营

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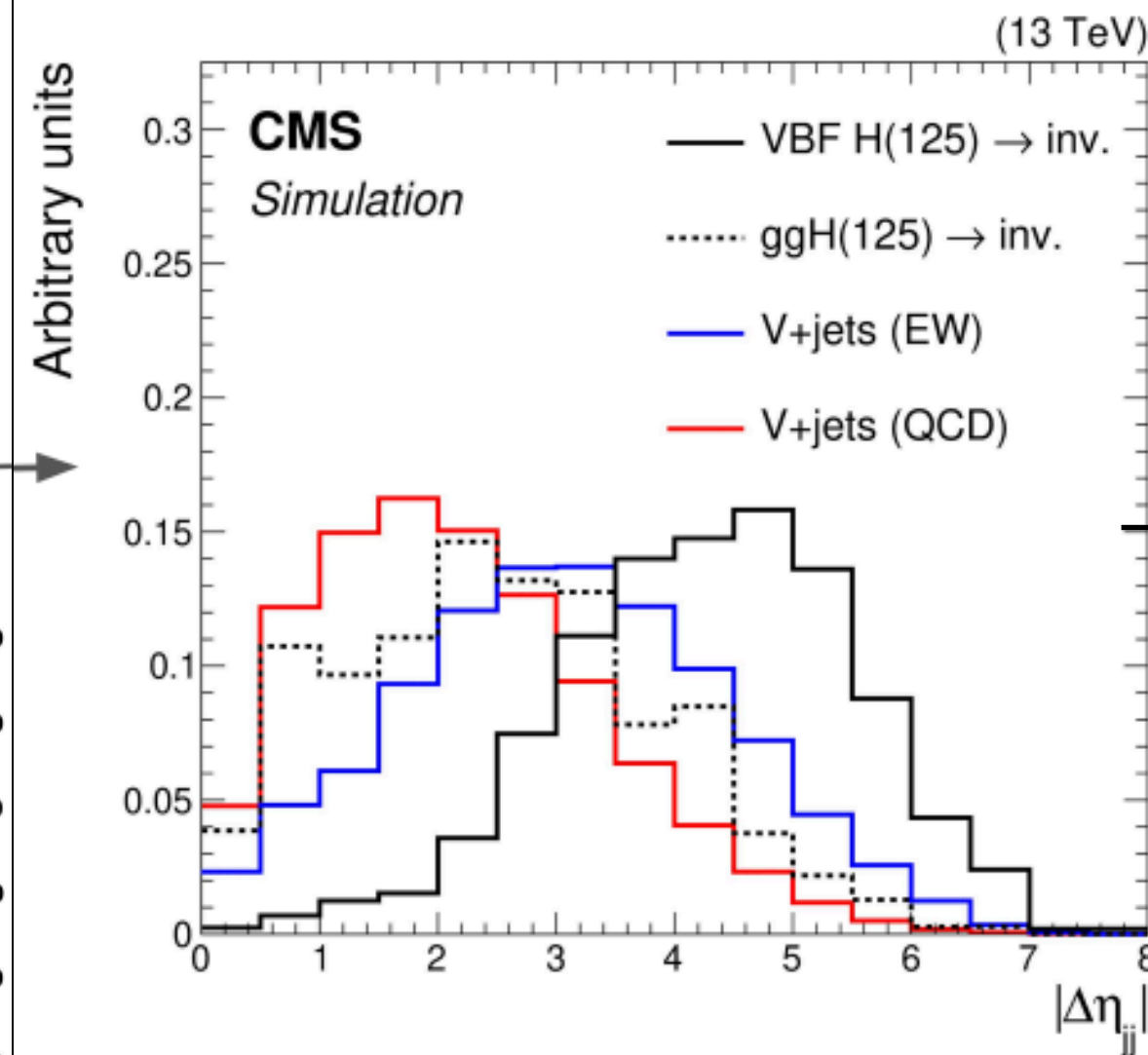
Follow the installation instruction

`/publicfs/cms/user/wangchu/CMSDAS_Stat/README.md`

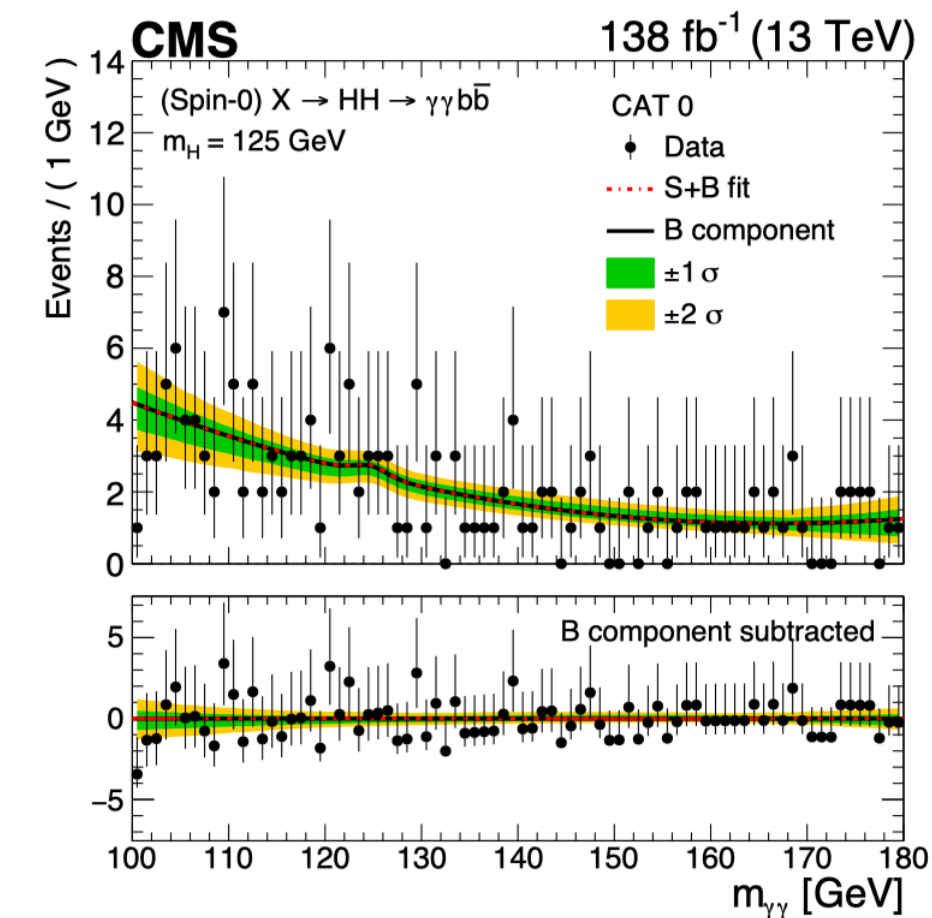
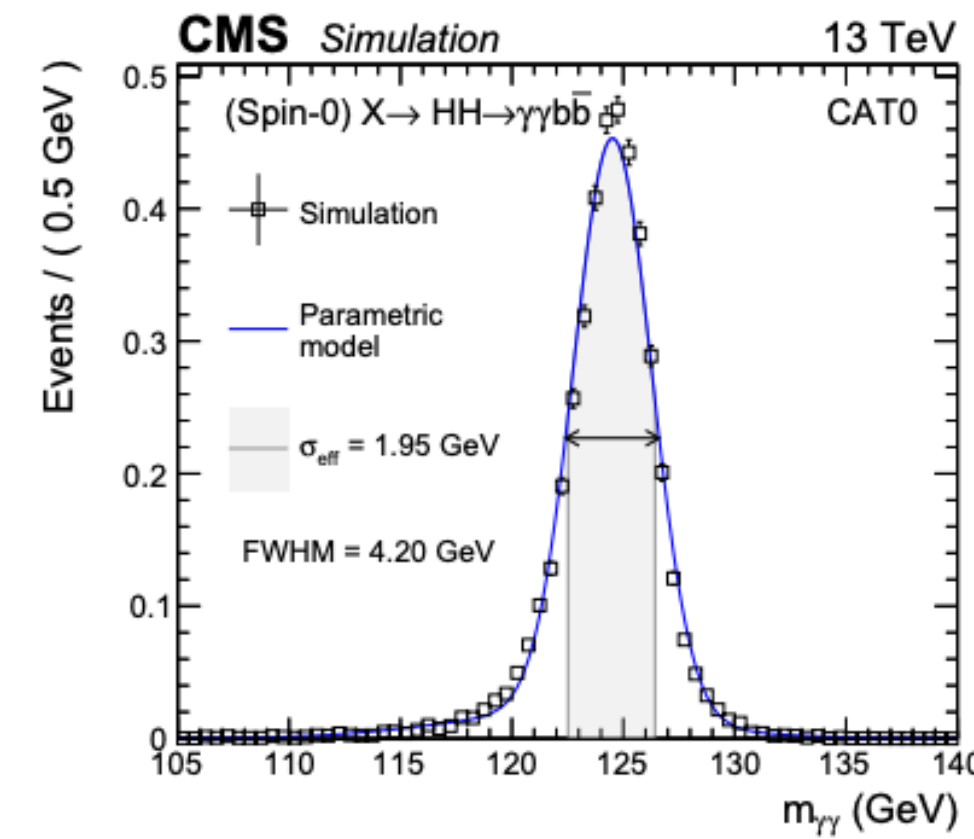
Common procedures of the Data analysis



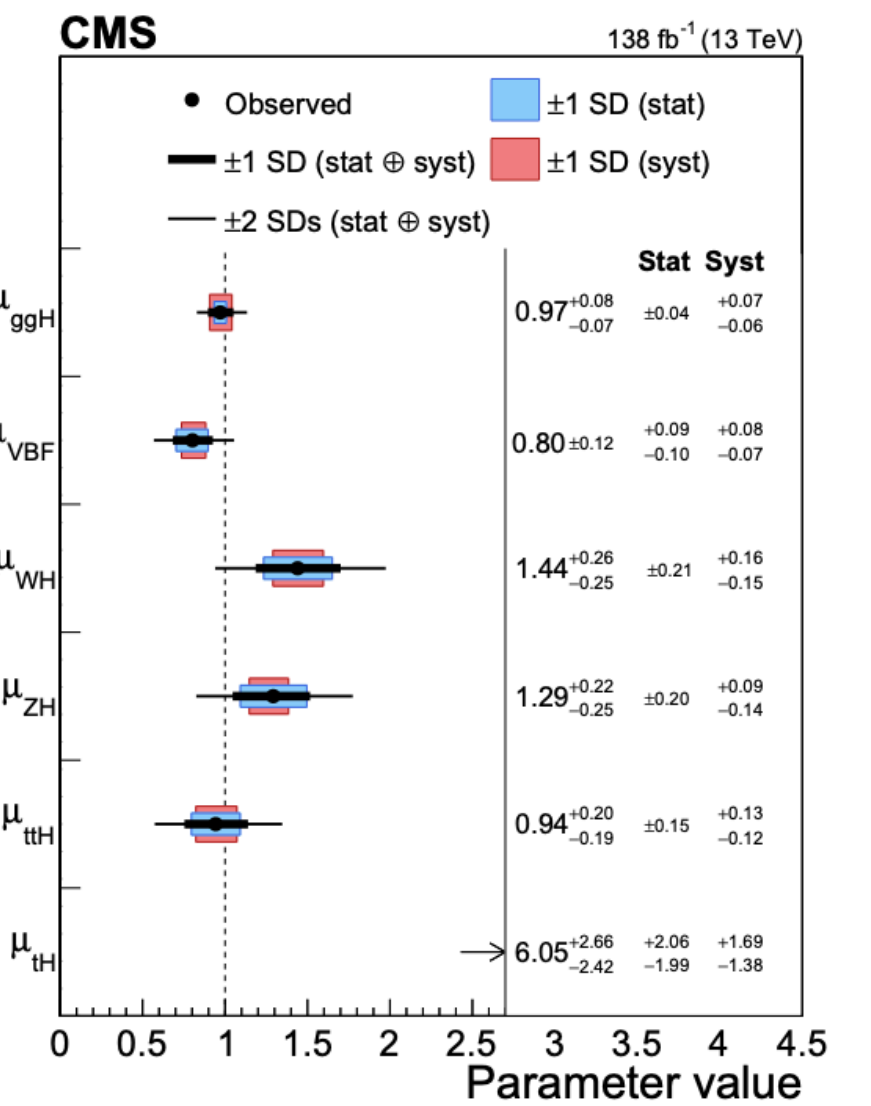
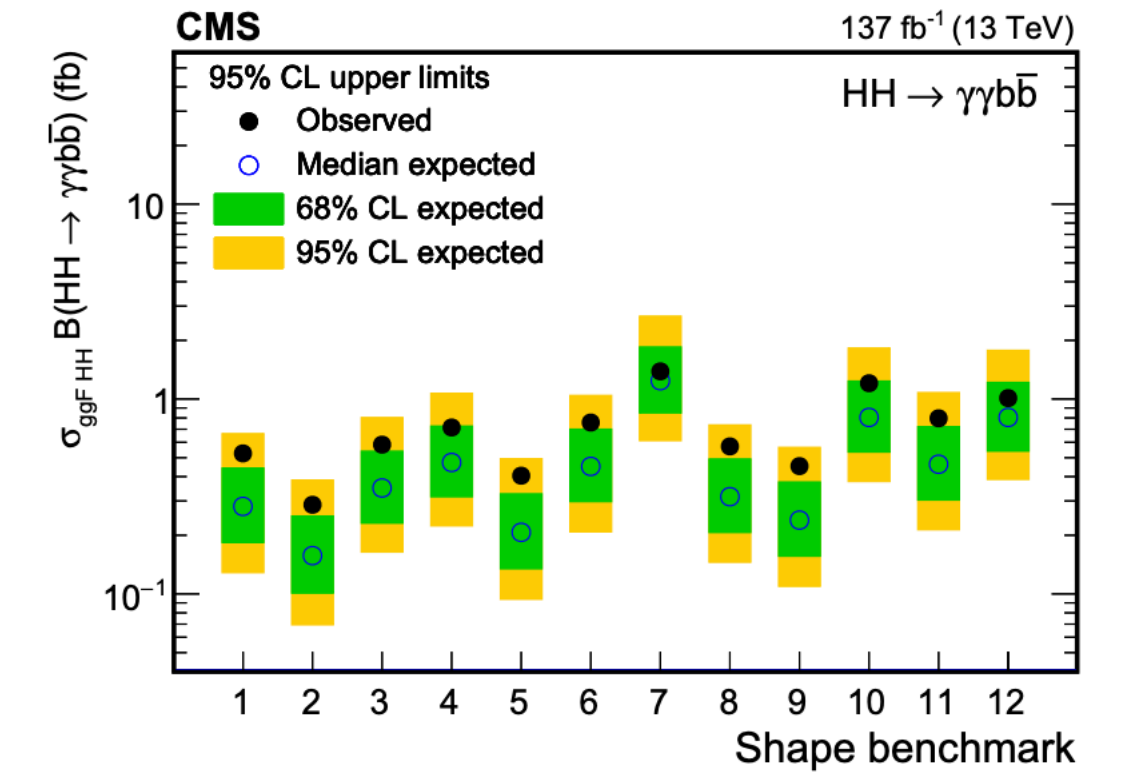
Data and MC samples



Event selection



bkg and sig models



Stat analysis

Common procedures of the stat analysis

Build signal & background models



Evaluate systematic uncertainties



Build **likelihood** with a physics **parameter of interest**, e.g. the signal normalization relative to some reference cross section , and incorporate the systematic uncertainties as **nuisance parameters**



Maximise the likelihood
(minimize negative log of the likelihood)
= **estimate parameters**



Use to define a **test statistic** for hypothesis testing

- **Likelihood** defined as

$$\mathcal{L}(\vec{\alpha}) \propto p(\text{data} | \vec{\alpha})$$

Parameters of the likelihood

Probability to observe the data for a given value of the likelihood parameters

- Note:
 - The likelihood is not a probability (various normalisation terms are ignored)

- Likelihood parameters: $\vec{\alpha} \Rightarrow (\vec{\mu}, \vec{\theta})$

Parameters of Interest (POIs)
= parameters we want to measure

Nuisance parameters
(or NP)

► Expected number of events :

$$n_{\text{exp}} = \mu \sigma_{\text{sig}} \epsilon_{\text{sig}} A_{\text{sig}} L^{\text{int}} + \sigma_{\text{bkg}} \epsilon_{\text{bkg}} A_{\text{bkg}} L^{\text{int}}$$

- μ : signal strength, σ : cross section, ϵ : selection efficiency, A: Detector Acceptance, L: Luminosity

► Construct likelihood by observed events (N) and expected events (n_{exp}):

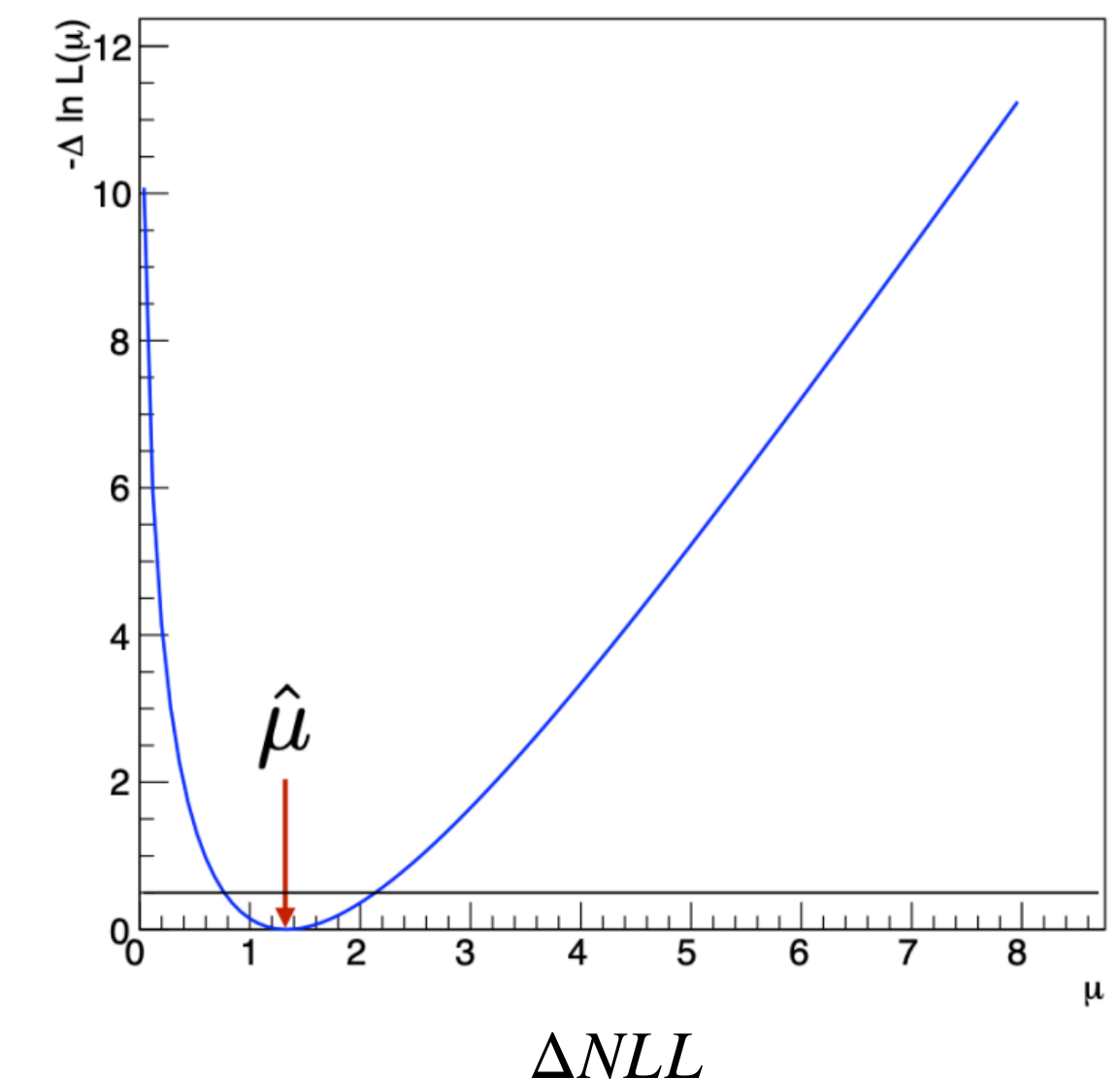
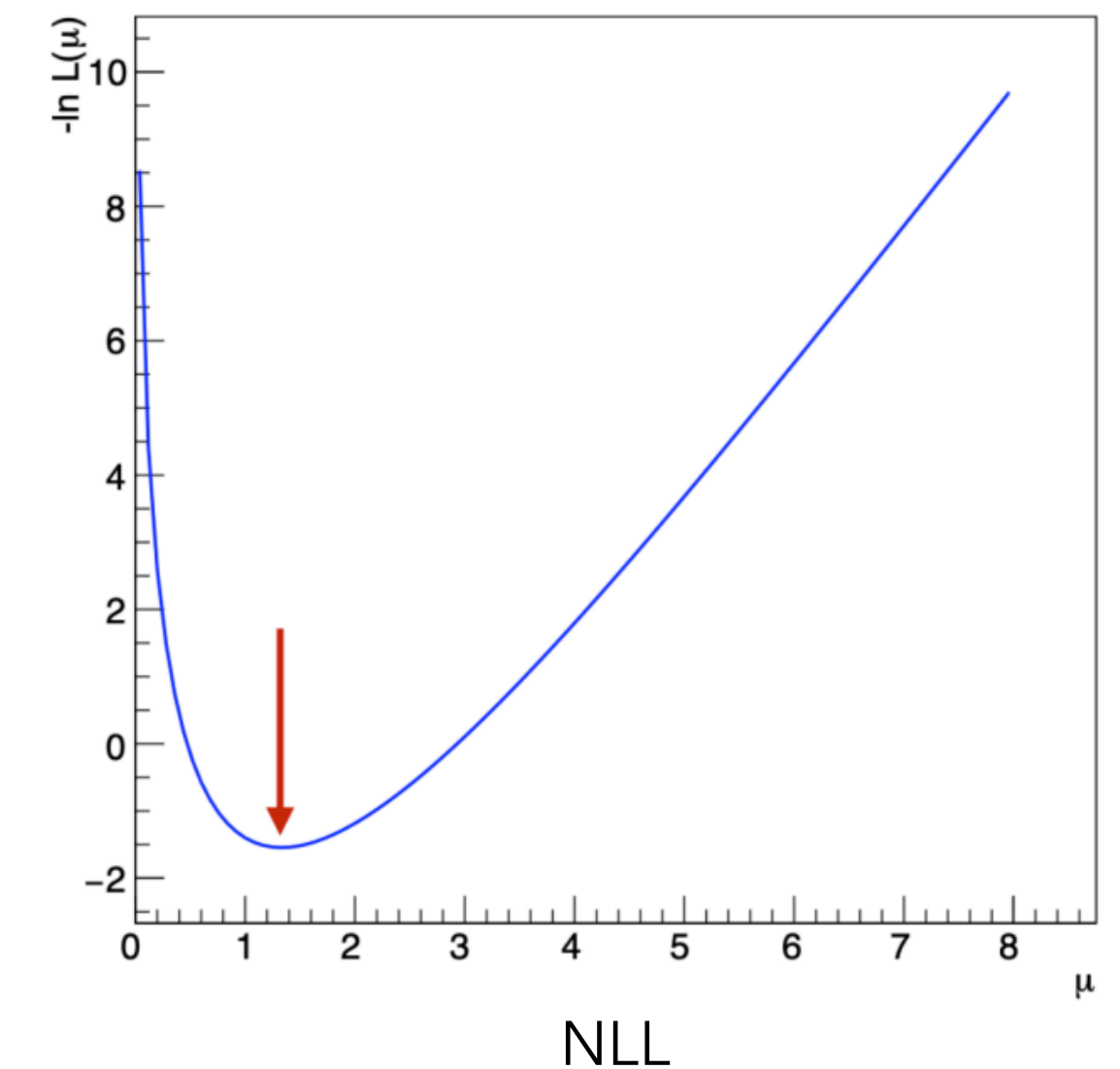
Poisson probability
$$p(N | n_{\text{exp}}) = \frac{n_{\text{exp}}^N e^{-n_{\text{exp}}}}{N!}$$

Description	Observable	Likelihood
Counting	n	Poisson $P(n; S, B) = e^{-(S+B)} \frac{(S+B)^n}{n!}$
Binned shape analysis	$n_i, i = 1 \dots N_{\text{bins}}$	Poisson product $P(\mathbf{n}_i; S, B) = \prod_{i=1}^{N_{\text{bins}}} e^{-(S f_i^{\text{sig}} + B f_i^{\text{bkg}})} \frac{(S f_i^{\text{sig}} + B f_i^{\text{bkg}})^{n_i}}{n_i!}$
Unbinned shape analysis	$m_i, i = 1 \dots n_{\text{evts}}$	Extended Unbinned Likelihood $P(\mathbf{m}_i; S, B) = \frac{e^{-(S+B)}}{n_{\text{evts}}!} \prod_{i=1}^{n_{\text{evts}}} S P_{\text{sig}}(m_i) + B P_{\text{bkg}}(m_i)$

Minimise the likelihood

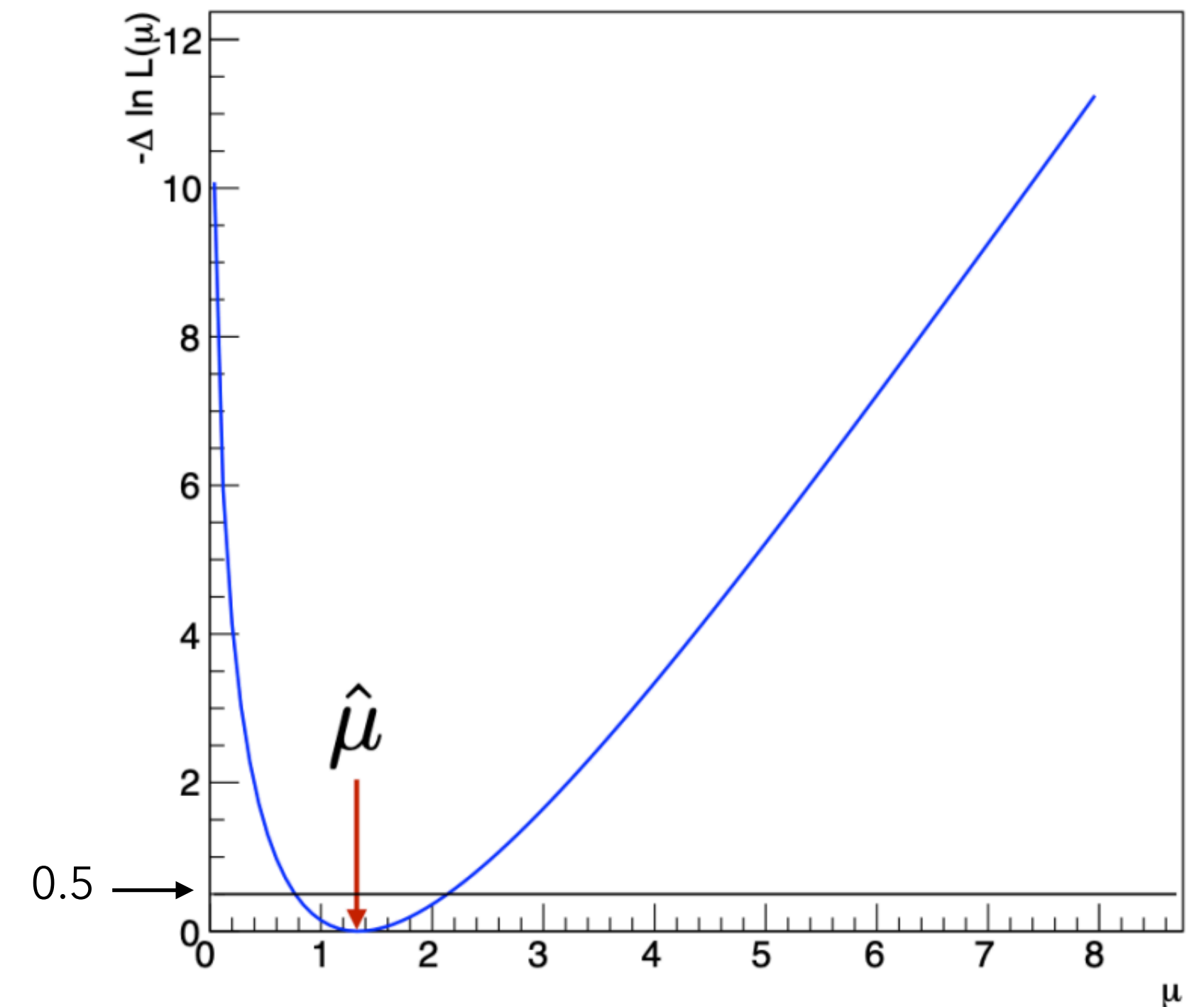
- ▶ Convert likelihood to Negative Log of the Likelihood (NLL), to avoid dealing large or small values
- ▶ When we do the minimisation, only care about the μ at the minimum of the likelihood, denoted by $\hat{\mu}$
 - Because the value of NLL is not important for signal strength scan, we can minus the minimum to get ΔNLL

$$\begin{aligned} -\Delta \ln \mathcal{L} &= -\ln \mathcal{L}(\mu, \hat{\theta}(\mu)) - (-\ln \mathcal{L}(\hat{\mu}, \hat{\theta})) \\ &= -\ln \frac{\mathcal{L}(\mu, \hat{\theta}(\mu))}{\mathcal{L}(\hat{\mu}, \hat{\theta})} \end{aligned}$$



► Set confidence intervals :

- The asymptotic distribution of $-2\Delta NLL$ follows the χ^2 distribution with K degrees of freedom, where the K is Difference in the number of free parameters in the numerator denominator (here $k=1$)
- According to the relationship between the χ^2 distribution and the p-value (p-value), When the degree of freedom $K=1$, if a confidence level of 68% (p-value=0.32) is required, the corresponding χ^2 value should be approximately 1
 - We can calculate $-2\Delta NLL < 1$, to get the 68% confidence interval
 - While $-2\Delta NLL < 3.84$, can get the 95% confidence interval



► Nuisance parameters θ :

- Nuisance parameters are parameters that appear in a statistical model that are not our primary parameters of interest, but which still need to be modeled and treated. These are usually parameters that are not directly related to the main goal of the research problem, but have an impact on model fitting and inference

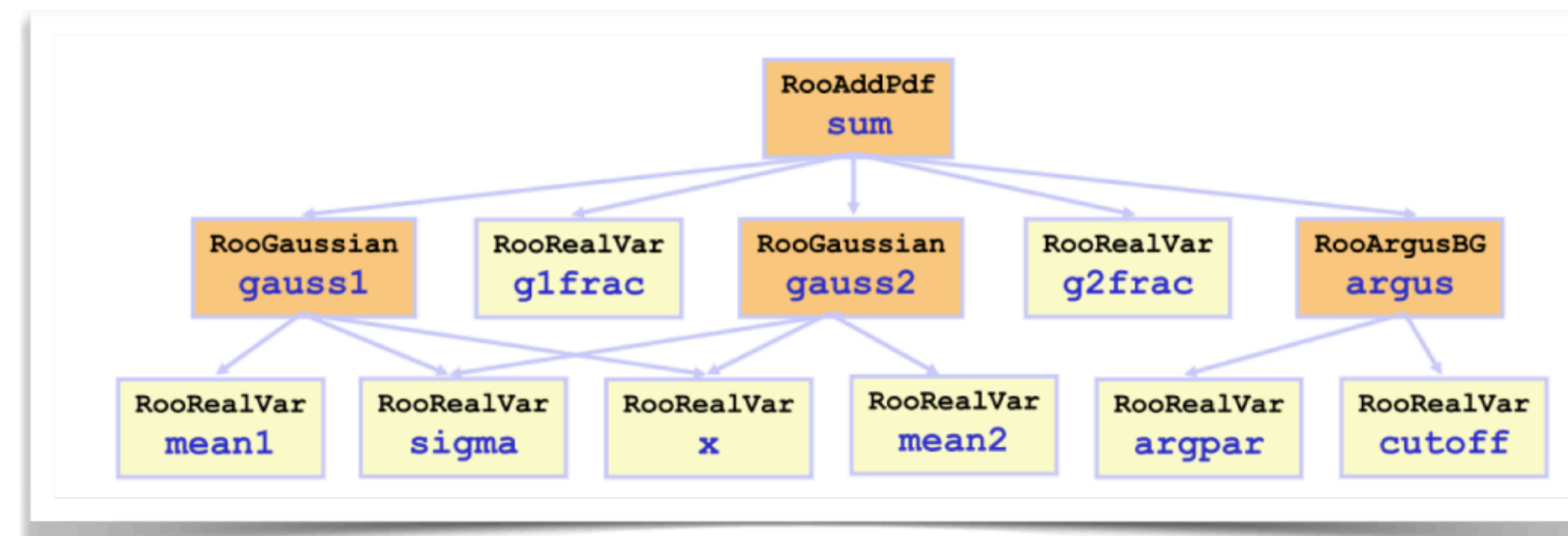
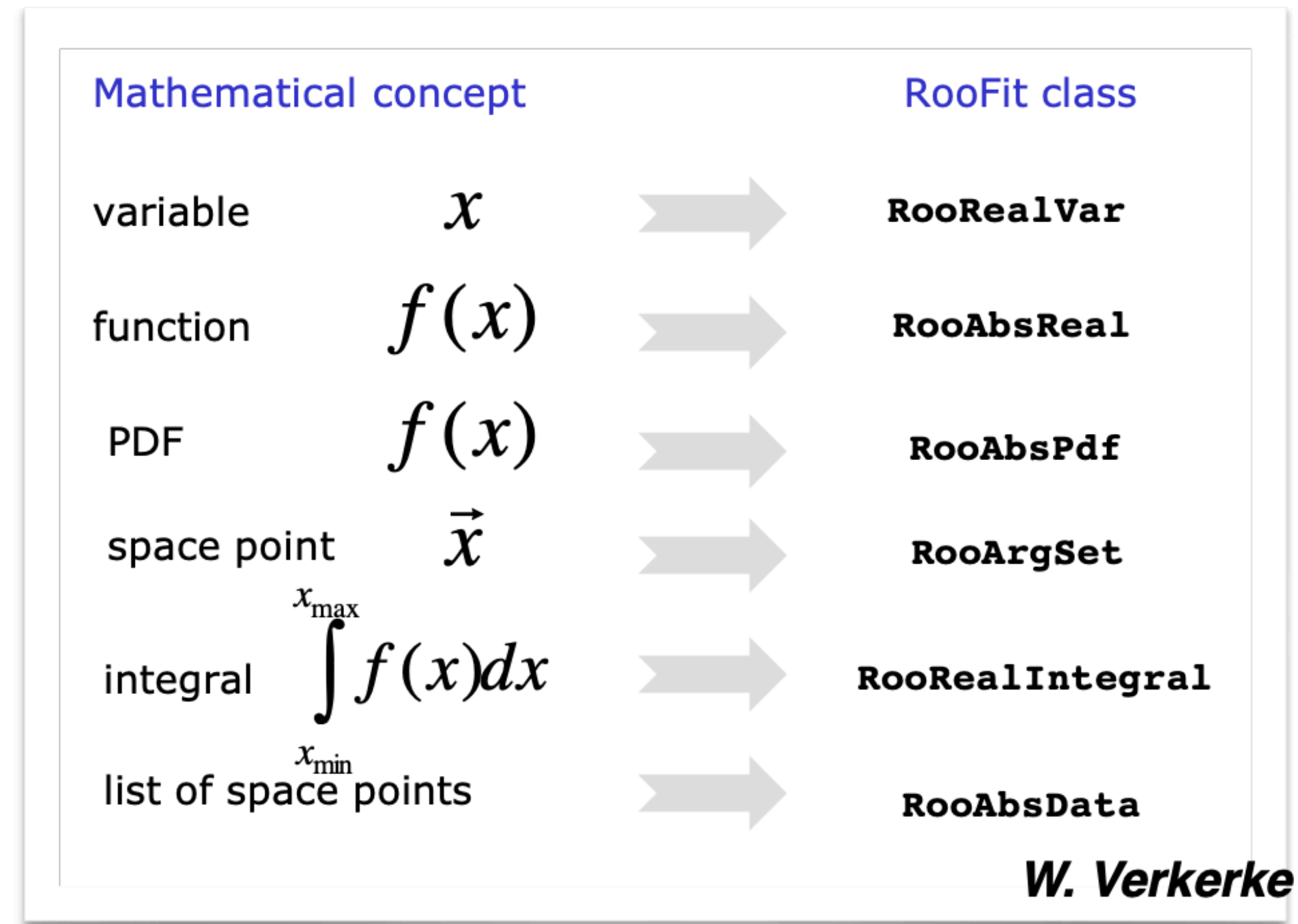
► Eg: luminosity

- If we measured the luminosity has 0.25% uncertainties, then it will either increase the number of instances by a factor of 1.025 or decrease it by a factor of 1/1.025.
- We can add it by a gaussian constraint

$$L^{\text{int}} \rightarrow L^{\text{int}}(1 + 0.025)^{\theta}$$

$$\mathcal{L}(\mu, \theta) = \frac{n_{\text{exp}}^N e^{-n_{\text{exp}}}}{N!} e^{-\frac{1}{2}\theta^2} \quad \text{where}$$
$$n_{\text{exp}} = \mu \sigma_{\text{sig}} \epsilon_{\text{sig}} A_{\text{sig}} L^{\text{int}1.025^{\theta}} + \sigma_{\text{bkg}} \epsilon_{\text{bkg}} A_{\text{bkg}} L^{\text{int}1.025^{\theta}}$$

- Framework built on top of ROOT for statistical analysis
- Objected-oriented approach
 - Specific PDFs deriving from abstract base classes, e.g. **RooGaussian** from **RooAbsPdf**
- Construct mathematical models by connecting objects together
- Provides interfaces for fitting and visualisation



Creating simple variables, pdfs, and likelihood functions with RooFit

Using RooFit to minimize the likelihood function

`/publicfs/cms/user/wangchu/CMSDAS_Stat/README.md`

Signal significance

► **Signal significance: degree of exclusion of background-only (b-only) hypotheses**

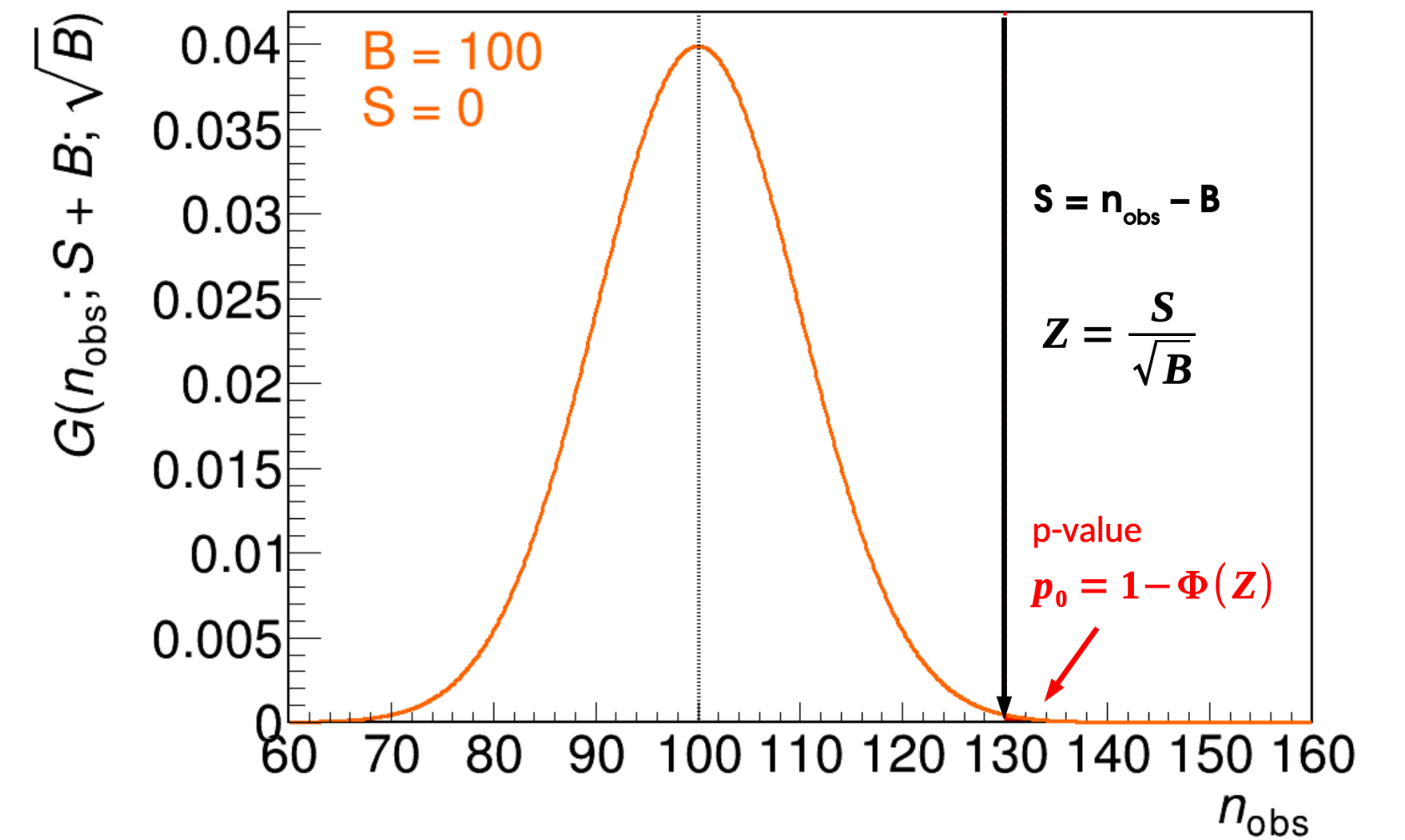
- Can be simply calculated by s/\sqrt{b} or $\sqrt{2n_0 \ln(1 + s/b) - 2s}$
- Generally denoted as Nx sigma, 3x sigma: evidence, 5x sigma: Observation

► **Signal significance with hypothesis testing:**

- Null hypothesis H_0 : No signal (b-only)
- Alternative hypothesis H_{alt} : any positive signal
- Discriminant (test statistic) : Likelihood ratio q_0

$$q_0 = -2 \log \frac{L(s = 0)}{L(\hat{s})}$$

- Can calculate p-value, use $p=1 - \Phi(Z)$ to get significance :Z






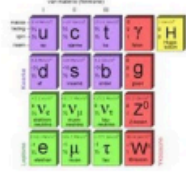
Bkg=100

n_{obs}	S	Z	p_0
105	5	0.5σ	31%
110	10	1σ	16%
120	20	2σ	2.3%
130	30	3σ	0.1%
150	50	5σ	$3 \cdot 10^{-7}$

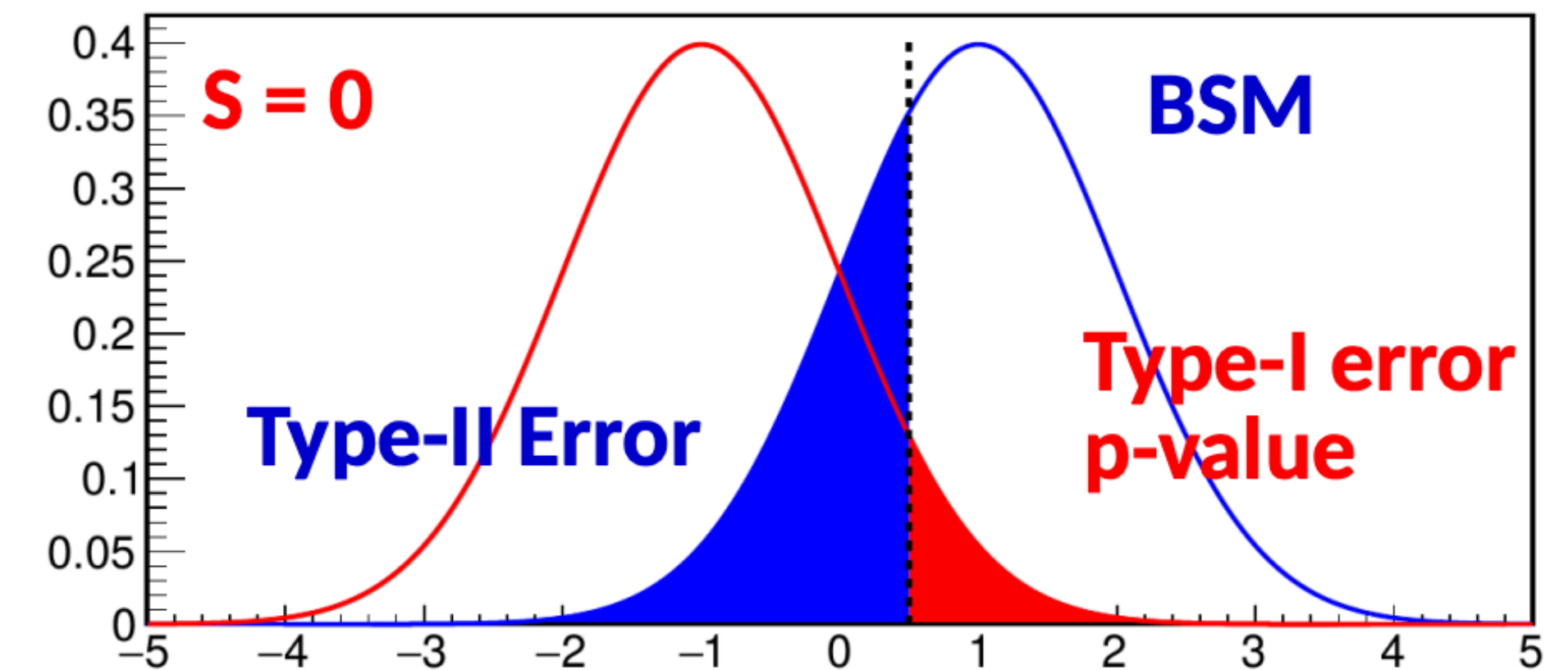
Bkg=100

► Type1 and Type2 errors

- Type 1 error (α error) : Rejecting the null hypothesis when it is true, i.e., drawing conclusions incorrectly, is called a Type 1 error.
- Type 2 error (β error) : Failure to reject the null hypothesis when it is false, i.e., failure to find an effect that actually exists

	Data disfavors H_0 (Discovery claim)	Data favors H_0 (Nothing found)
H_0 is false (New physics!)	Discovery! 	Type-II error (Missed discovery) 
H_0 is true (Nothing new)	Type-I error (False discovery) 	No new physics, none found 

↑ p-value, significance



Compute significance

`/publicfs/cms/user/wangchu/CMSDAS_Stat/README.md`

Set Upper Limit for the parameter

► When searching for undiscovered processes, because their signal significance is too small, they are often measured by setting an upper limit to the range of the parameter

- In practice, the test statistic first needs to be designed

- Considering the upper limit as a one-sided confidence interval $[0, \mu_{up}]$, the previously mentioned likelihood ratio was modified to obtain the new test statistic

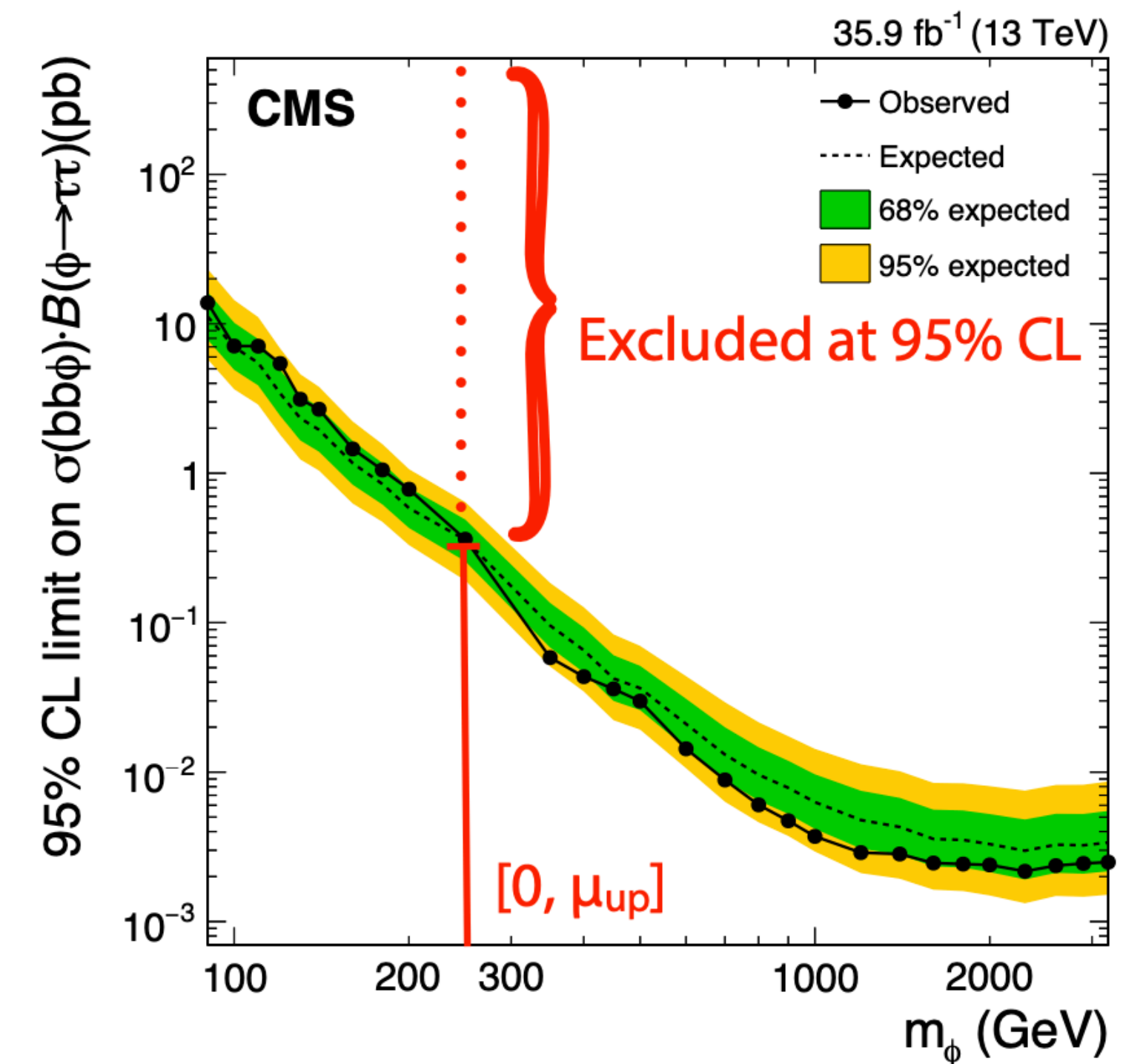
$$q_\mu = -2 \ln \frac{L(\mu, \hat{\theta}_\mu)}{L(\hat{\mu}, \hat{\theta})} \quad \Rightarrow \quad q_\mu = \begin{cases} -2 \ln \frac{L(\mu, \hat{\theta}_\mu)}{L(0, \hat{\theta}_0)} & \hat{\mu} < 0 \\ -2 \ln \frac{L(\mu, \hat{\theta}_\mu)}{L(\hat{\mu}, \hat{\theta})} & 0 \leq \hat{\mu} \leq \mu \\ 0 & \hat{\mu} > \mu \end{cases}$$

2-sided confidence intervals

Modified for upper limits

- when $\hat{\mu} < 0$, $\hat{\mu}$ has been set to 0, avoid negative values

- While $\mu < \hat{\mu}$, set test statistic to 0, ensure we can get one-sided intervals

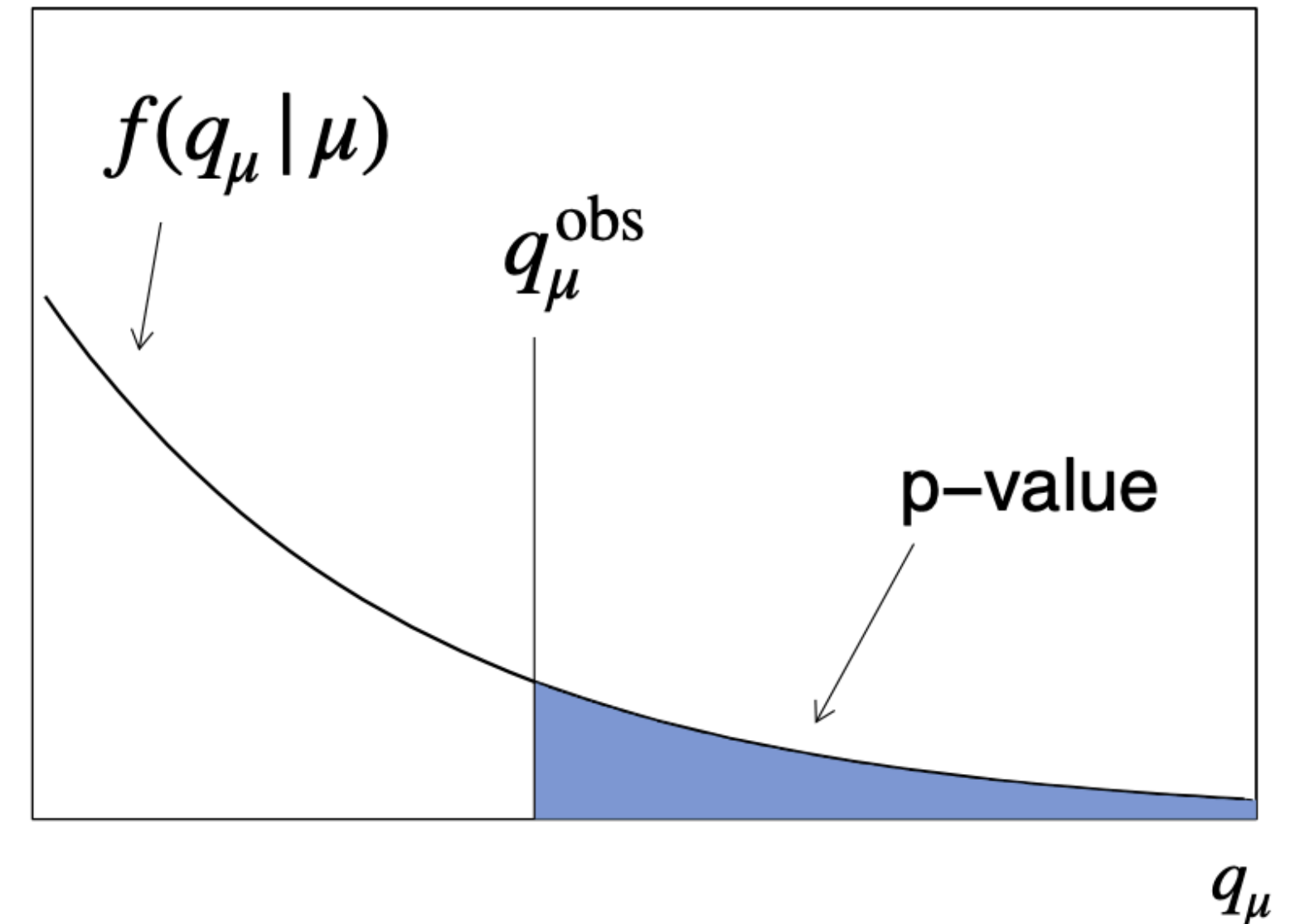


Set Upper Limit for the parameter

- ▶ With the distribution of the test statistic, the p-value can be calculated

$$p_{\mu} = P(q_{\mu} > q_{\mu}^{\text{obs}} | \mu) = \int_{q_{\mu}^{\text{obs}}}^{+\infty} f(q_{\mu} | \mu, \hat{\theta}_{\mu}) dq_{\mu}$$

- ▶ In the high-energy physics community, it is common to use the CLs criterion to set different confidence levels (commonly 95% CLs)



$$CL_s = \frac{CL_{s+b}}{CL_b}$$
$$CL_{s+b} = P(q_{\mu} > q_{\mu}^{\text{obs}} | \text{sig} + \text{bkg}) = \int_{q_{\mu}^{\text{obs}}}^{+\infty} f(q_{\mu} | \mu, \hat{\theta}_{\mu})$$
$$CL_b = P(q_{\mu} > q_{\mu}^{\text{obs}} | \text{bkg only}) = \int_{q_{\mu}^{\text{obs}}}^{+\infty} f(q_{\mu} | 0, \hat{\theta}_0)$$

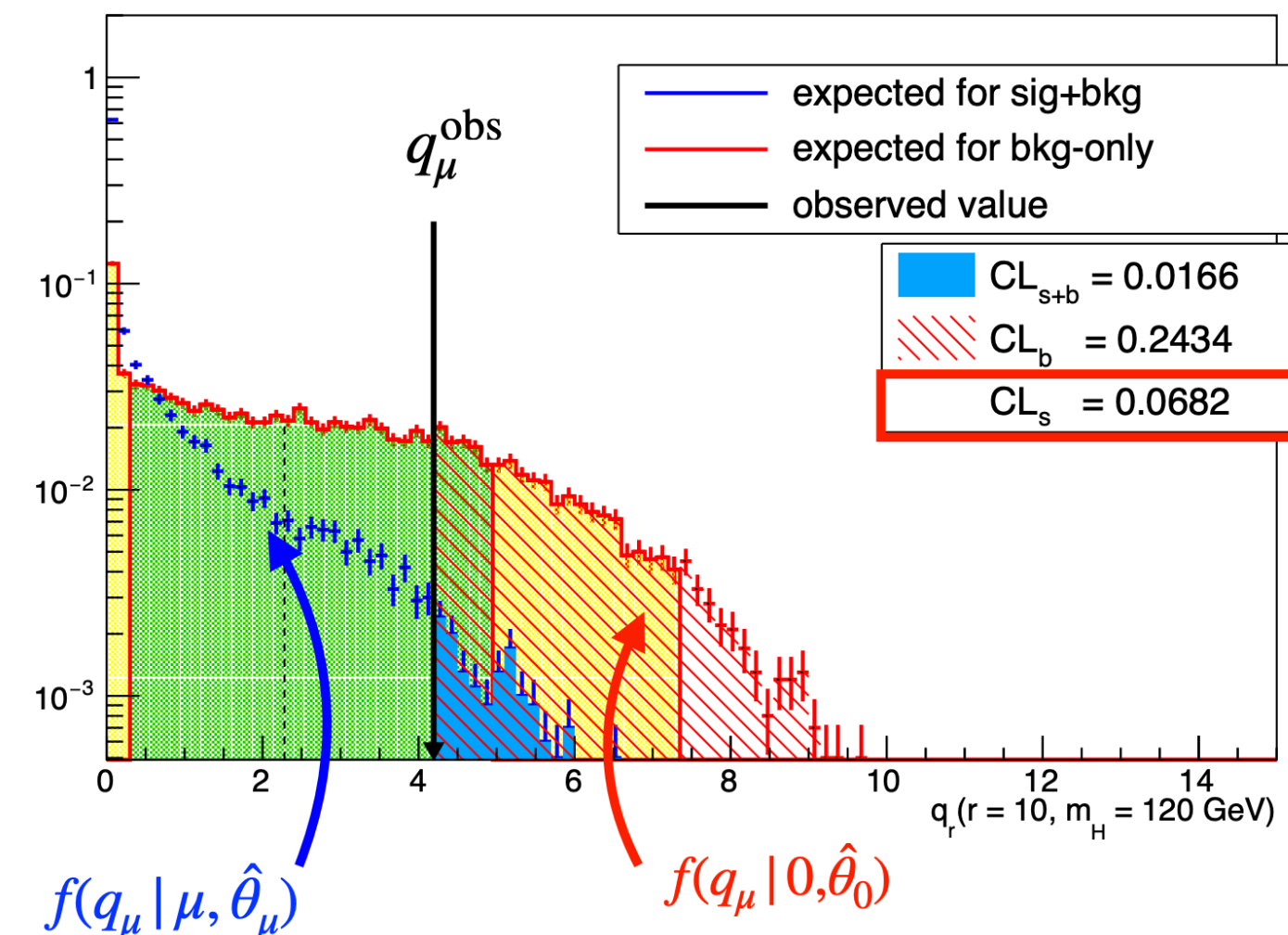
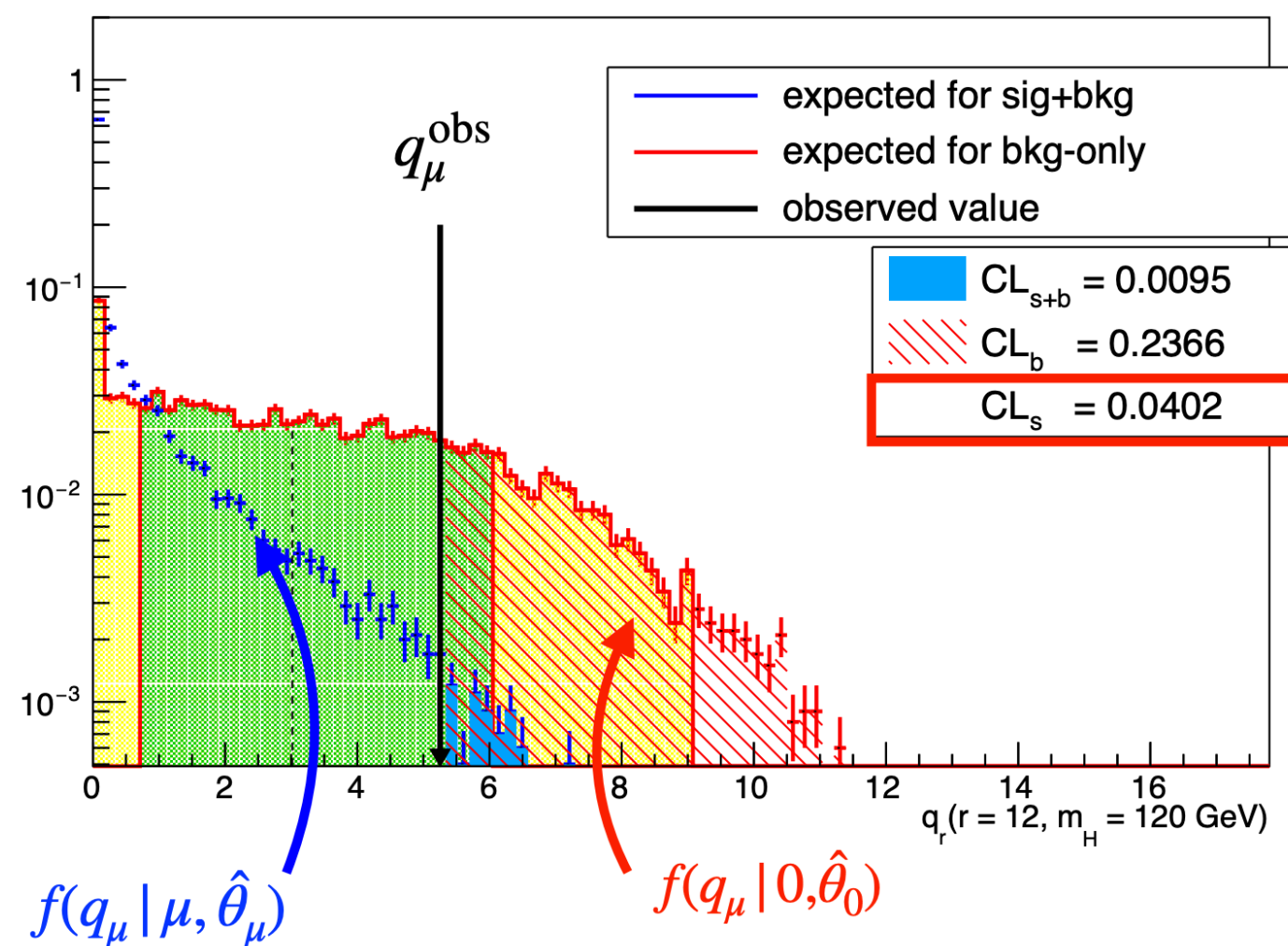
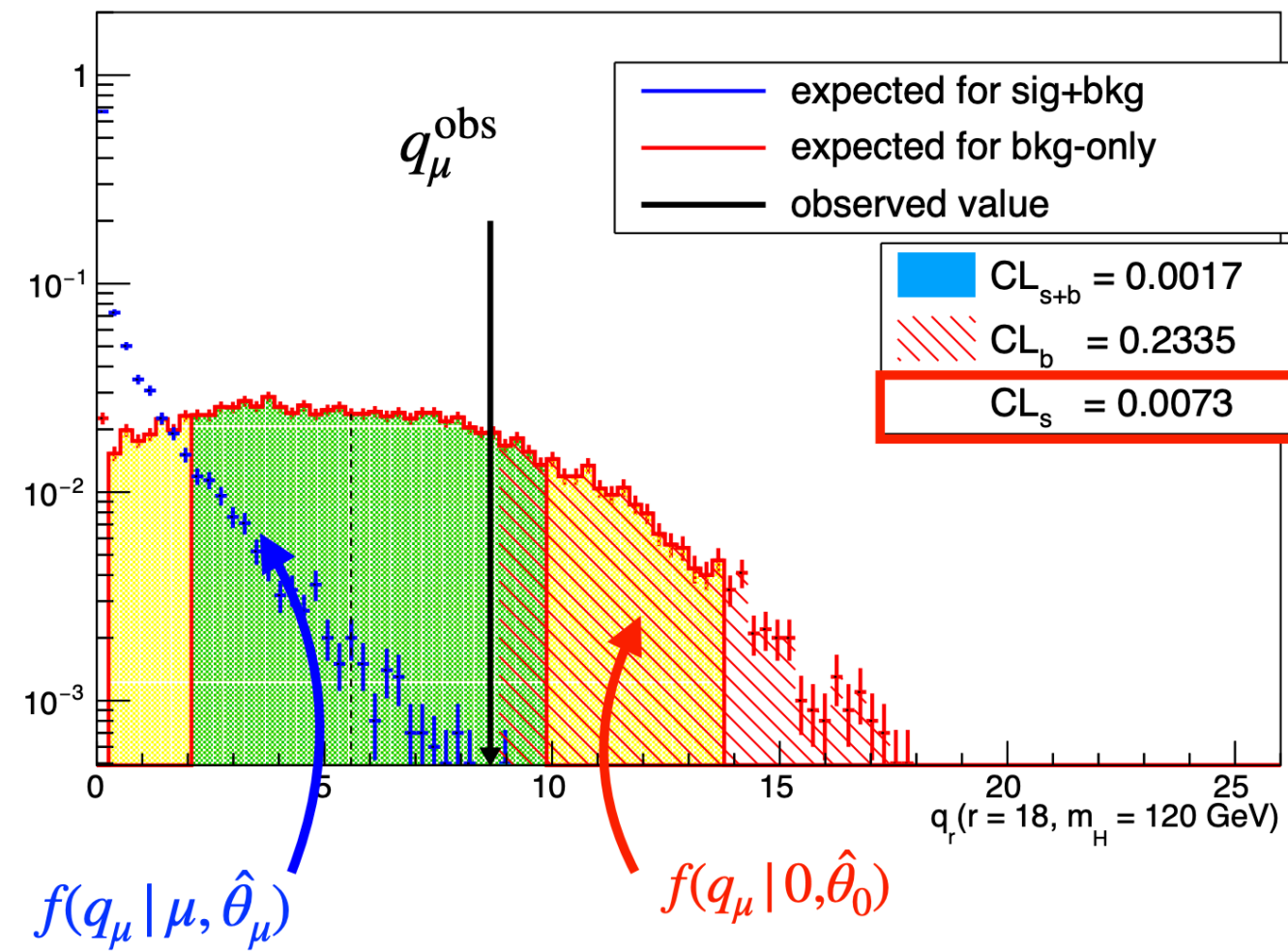
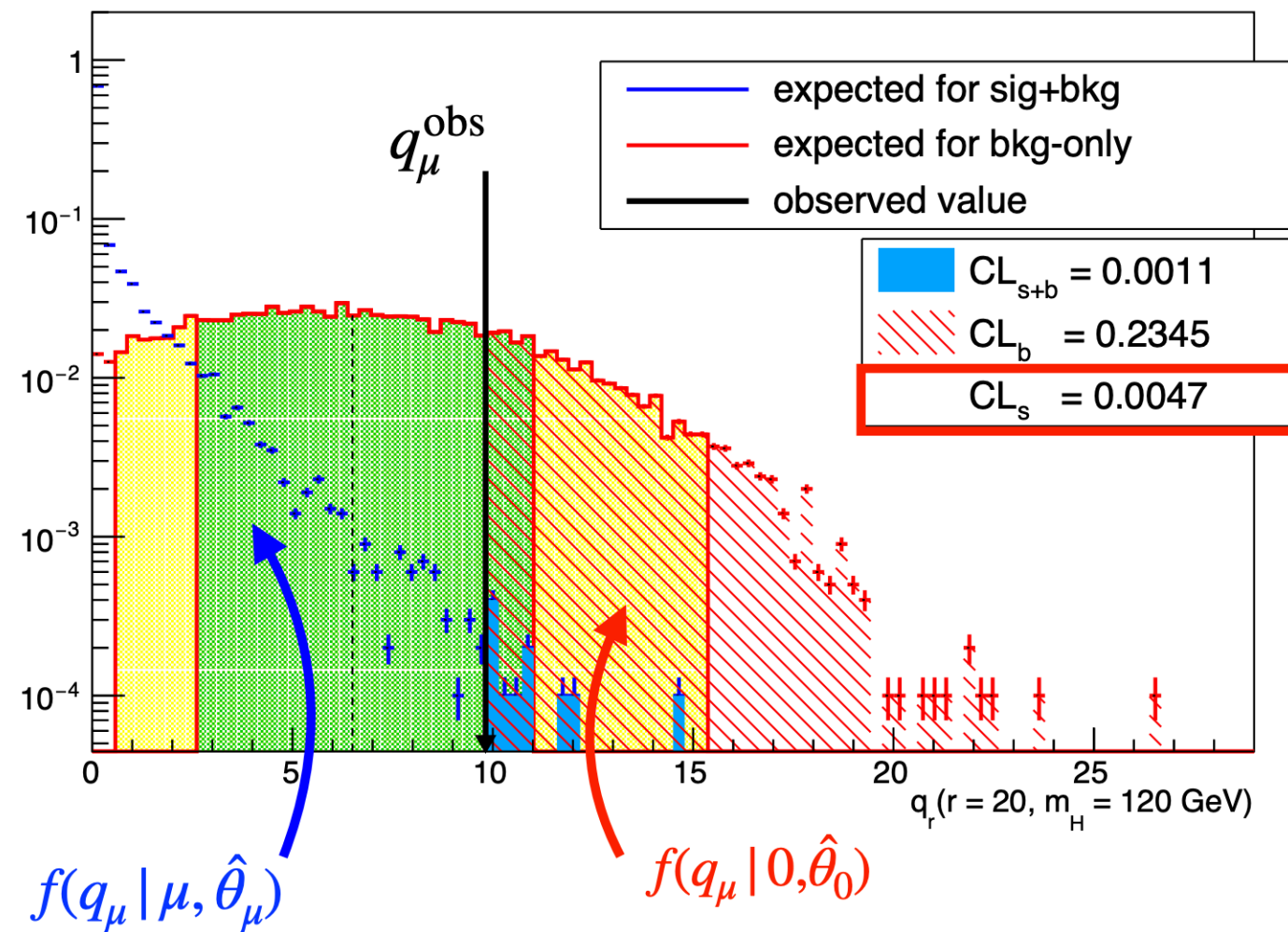
► Target:

- Determine the minimum value of signal strength μ_{up} when the CLs value is less than some determined p-value α ($\alpha=0.05$ at 95% CLs)

► Workflows:

- For each μ , generate some toy datasets based on s+b and b-only hypothesis
- Calculate the test statistics q_{μ} , build the distribution of q_{μ} in s+b and b-only hypothesis
- Calculate the p-values, namely CL_{s+b} and CL_b
- Calculate CLs, we can get the upper limit at 95% CLs while the CLs crossed 0.05

Toy method



- By scanning the (r) value, one can see:
 - Between $r = 12$ and $r = 10$, the value of CLs crosses 0.05, then the observed upper limit is between 12 and 10
- If you want to claim the Expected upper bound, you need to replace q_μ^{obs} with a different quantile value of CLb
- Commonly used quantiles are:
 - $[0.025, 0.16, 0.5, 0.84, 0.975]$
 - Corresponding to the median and $\pm 1/2\sigma$

- ▶ When the model is too complex, if the Toy method is utilized to take the upper limit, as it needs to generate a large number of Toy samples and calculate the p-values, it will consume a lot of time and computational resources
- ▶ Therefore, Asymptotic approximation can be utilized to save resources and time

- Do not need to generate the toys for p-values

$$q_{\mu,A} = -2\ln \frac{L(\text{Asimov}|\mu, \theta(\mu))}{L(\text{Asimov}|\hat{\mu}, \hat{\theta})} \quad q_{\mu} = -2\ln \frac{L(\text{data}|\mu, \theta(\mu))}{L(\text{data}|\hat{\mu}, \theta(\hat{\mu}))} \quad \begin{aligned} CL_{sb} &= 1 - \Phi(q_{\mu} + q_{\mu,A}/2 * \sqrt{q_{\mu,A}}) \\ CL_b &= 1 - \Phi(q_{\mu} - q_{\mu,A}/2 * \sqrt{q_{\mu,A}}) \end{aligned} \quad q_{\mu,A} = [\Phi^{-1}(CL_b) - \Phi^{-1}(CL_{s+b})]^2/2$$

- For the derivation see. [[Cowan, Cranmer, Gross, Vitells 2013](#)]
- A in the formula represents Asimov dataset. Asimov dataset is an idealized dataset, which suppressed stat uncertainties
- Eg, when we want to calculate the expected median at 95% CLs:
 - $CL_b=0.5$, $CL_s=0.05$, $CL_{s+b}=0.5*0.05=0.025$, then $q_{\mu,A}=3.84/2$
- Scan μ by using formula, once the value crossed $3.84/2$, then we found the median.
- CL_b/CL_{s+b} could be calculated by q_{μ} and $q_{\mu,A}$
 - For observed limit, replace the Asimov dataset to data, calculate CL_{s+b} and CL_b , then CLs

- ▶ **Combine: RooStats / RooFit - based software tools used for statistical analysis in CMS**
- ▶ **It provides a command line interface to many different statistical techniques available inside**
 - Parameter estimation and setting of confidence intervals
 - Signal significance
 - Upper limit
 - Provides many statistical checking tools (FitDiagnostic, Impact, GOF, Bias, etc.)
 -
- ▶ **Github: [Link](#)**
- ▶ **Documents: [Combine Tool](#)**

► Datacard settings

- For the combine usage, the first step is to generate datacard

Number of bins/channels Number of processes Number of nuisance parameters (*:determined automatically)

```
imax 1 number of bins
jmax 4 number of processes minus 1
kmax * number of nuisance parameters
```

bin	signal_region	Unique channel label				
observation	10.0	Number of observed events in channel				

bin	signal_region	signal_region	signal_region	signal_region	signal_region	Process label
process	ttbar	diboson	Ztautau	jetFakes	bbHtautau	Process ID (<=0 for signal)
process	1	2	3	4	0	Expected number of events
rate	4.43803	3.18309	3.7804	1.63396	0.711064	

Name	Type	Effect on process					Systematic uncertainties
CMS_eff_b	lnN	1.02	1.02	1.02	-	1.02	
CMS_eff_t	lnN	1.12	1.12	1.12	-	1.12	
CMS_eff_t_highpt	lnN	1.1	1.1	1.1	-	1.1	
acceptance_Ztautau	lnN	-	-	1.08	-	-	
acceptance_bbH	lnN	-	-	-	-	1.05	
acceptance_ttbar	lnN	1.005	-	-	-	-	
lumi_13TeV	lnN	1.025	1.025	1.025	-	1.025	
norm_jetFakes	lnN	-	-	-	1.2	-	
xsec_Ztautau	lnN	-	-	1.04	-	-	
xsec_diboson	lnN	-	1.05	-	-	-	
xsec_ttbar	lnN	1.06	-	-	-	-	

► text to workspace

- To saving the running time of the combine, could covert the datacard in text format to RooFit workspace
- `text2workspace.py datacard.txt -m Mass -o workspace.root`

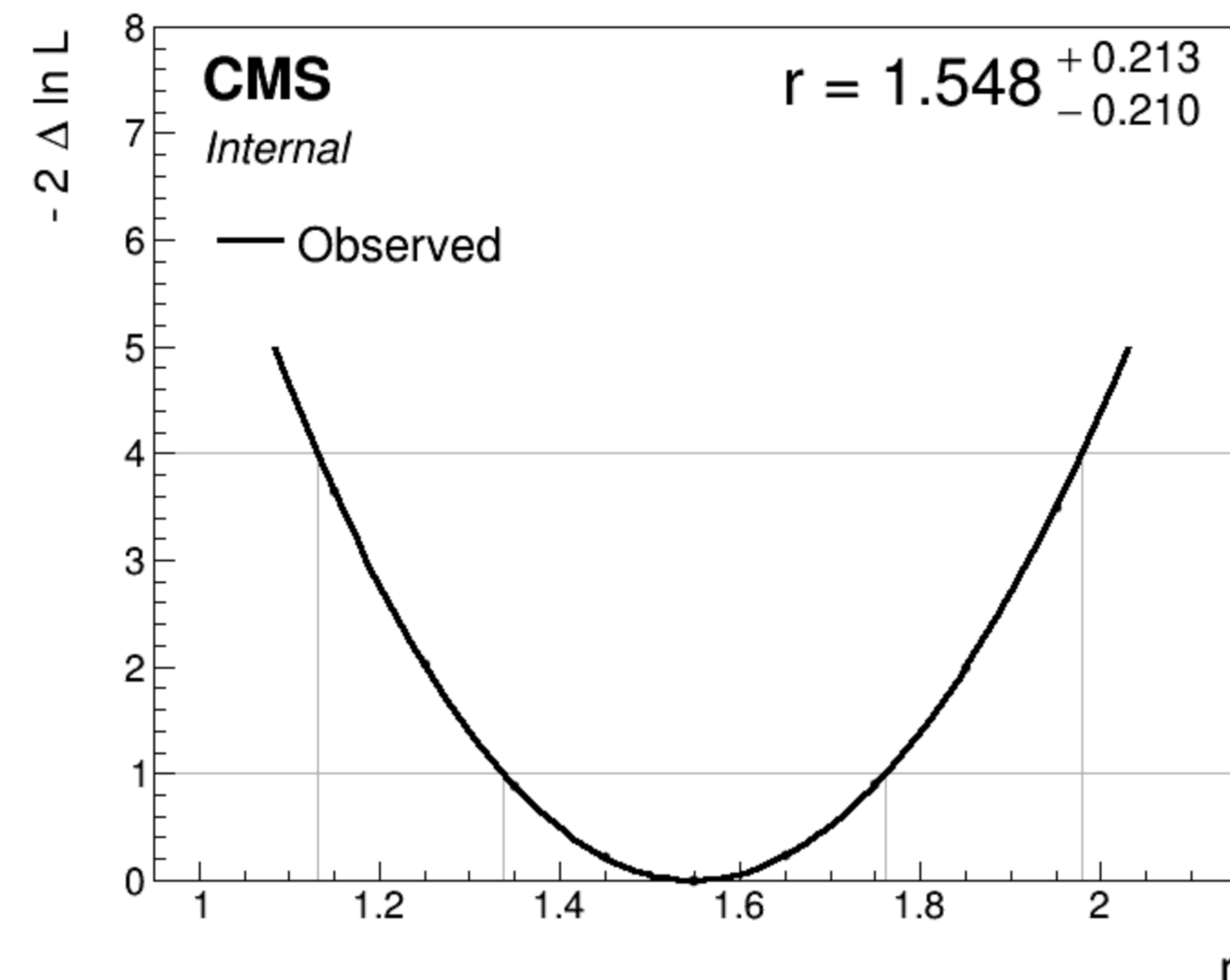
► Best fit of the POI and confidence interval

- bestfit:

- combine -M MultiDimFit
datacard_part1_with_norm.root -m 125 --
freezeParameters MH --saveWorkspace
-n .bestfit

- confidence interval:

- combine -M MultiDimFit
datacard_part1_with_norm.root -m 125 --
freezeParameters MH -n .scan --algo
grid --points 20 --setParameterRanges
r=lo,hi
 - plot1DScan.py
higgsCombine.scan.MultiDimFit.mH125.root
t -o part2_scan



► Set upper limits (Toy method)

- Observed:

- `combine -M HybridNew datacard.txt --LHCmode LHC-limits --saveHybridResult`

```
-- Hybrid New --  
Limit: r < 10.9705 +/- 0.386687 @ 95% CL  
Done in 0.47 min (cpu), 0.47 min (real)
```

- Expected:

- `combine -M HybridNew datacard.txt --LHCmode LHC-limits --saveHybridResult --expectedFromGrid 0.5`

```
-- Hybrid New --  
Limit: r < 14.2678 +/- 0.217055 @ 95% CL  
Done in 0.62 min (cpu), 0.62 min (real)
```

- Plotting:

- `python $CMSSW_BASE/src/HiggsAnalysis/CombinedLimit/test/plotTestStatCLs.py --input higgsCombine.HybridNew.mH120.root --poi r --val all --mass 120`
 - `python printTestStatPlots.py cls_qmu_distributions.root`

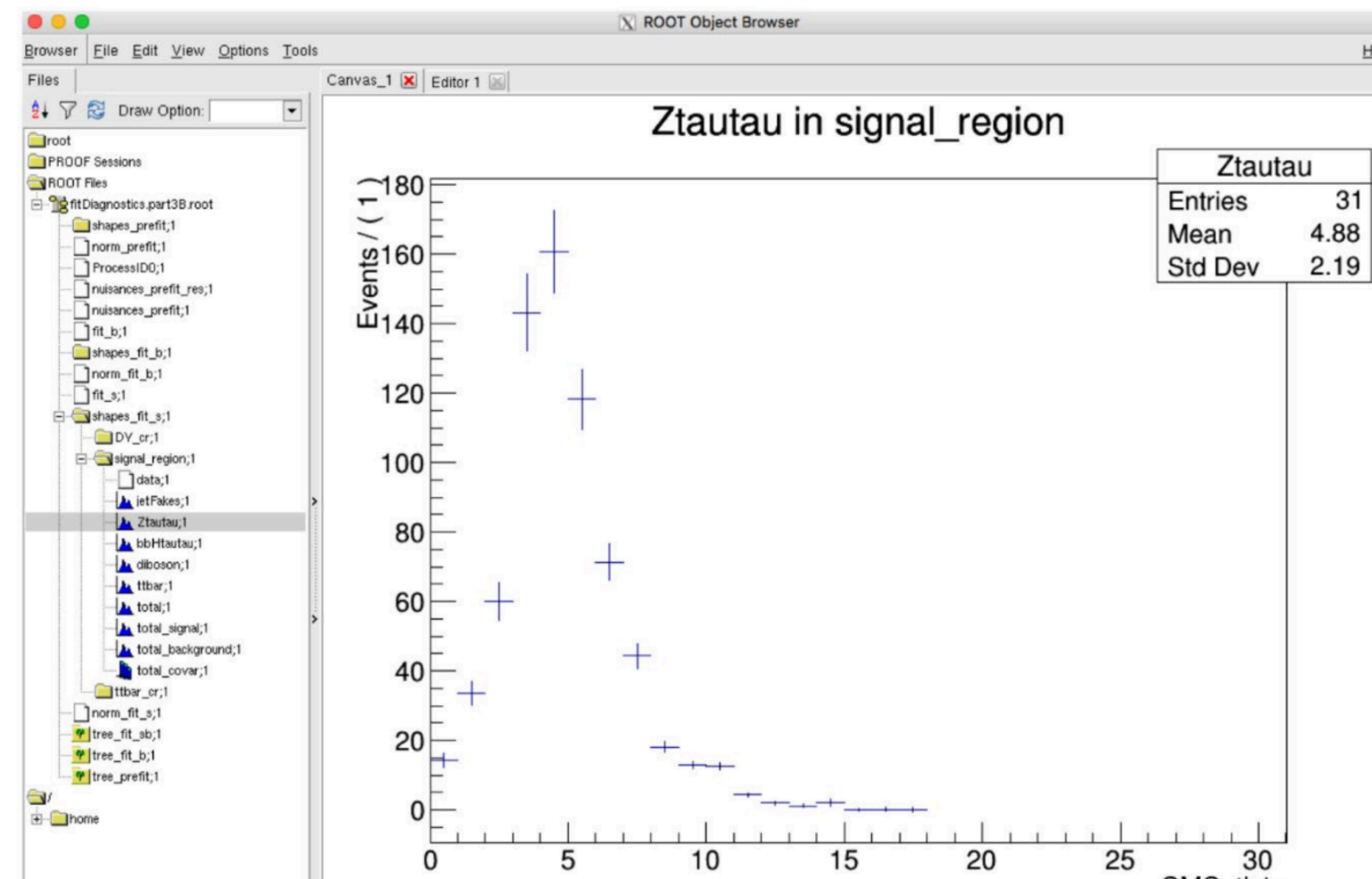
- ▶ **Set upper limits (AsymptoticLimits method)**
 - `combine -M AsymptoticLimits workspace.root`

```
-- AsymptoticLimits ( CLs ) --  
Observed Limit: r < 10.8183  
Expected 2.5%: r < 7.0537  
Expected 16.0%: r < 9.8108  
Expected 50.0%: r < 14.5625  
Expected 84.0%: r < 22.3988  
Expected 97.5%: r < 33.5971
```


► Pre and post fit (FitDiagnostics)

- `combine -M FitDiagnostics workspace.root -m 200 --rMin -1 --rMax 2 --saveShapes --saveWithUncertainties`
- It will do b-only and s+b fits, gotten pre/post fits

Combine will produce pre- and post-fit distributions (for fit_s and fit_b) in the fitdiagnostics.root output file:



Test combine tools

`/publicfs/cms/user/wangchu/CMSDAS_Stat/README.md`