

13TeV VBF $H \rightarrow \gamma\gamma$ analysis

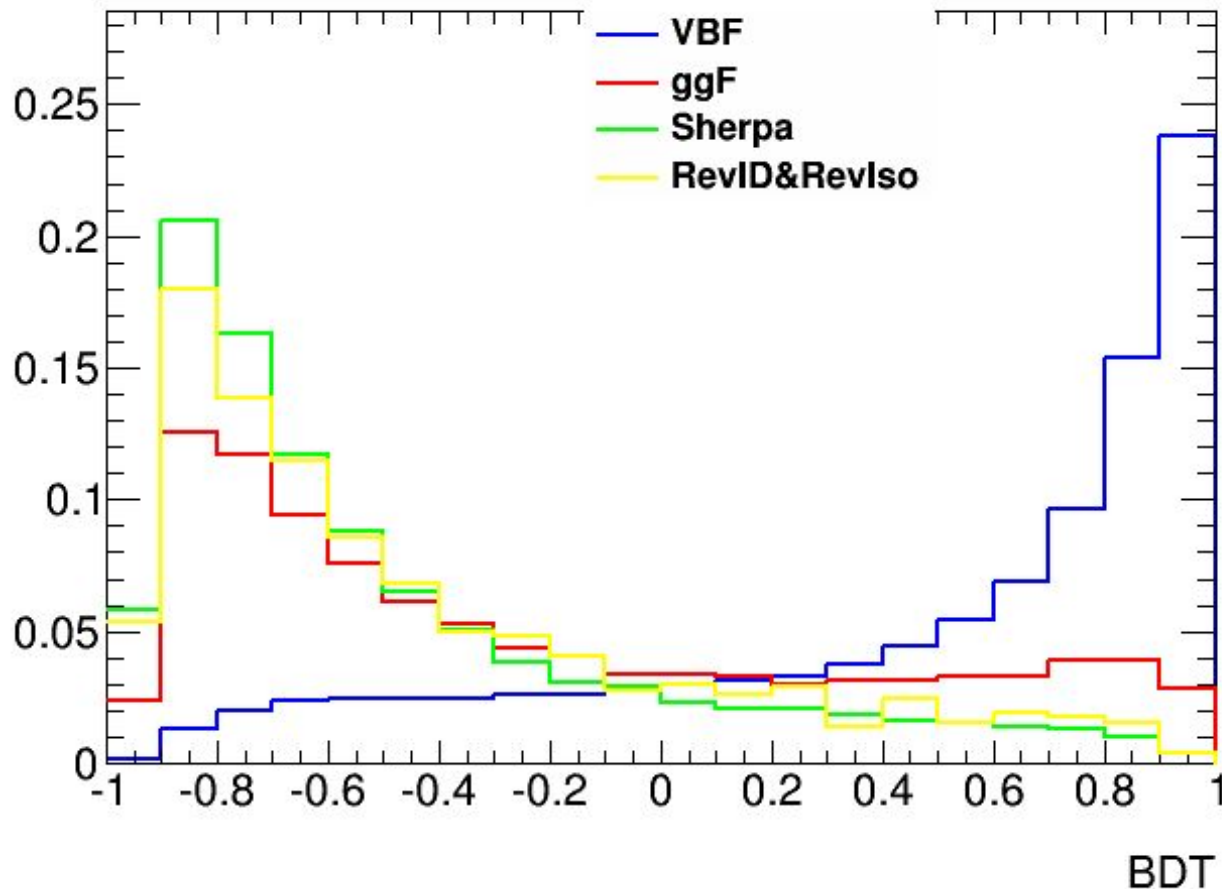
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09-21

MVA

- input variables
 - $m_{jj}, \Delta\eta_{jj}, \Delta\Phi_{\gamma\gamma, jj}, p_{Tt}, \Delta R_{y,j}^{\min}, \eta^*$
- preselection
 - $N_{jet} \geq 2, \Delta\eta_{jj} > 2, \eta^* < 5$
- configuration
 - default one same as run1
- training sample
 - SignalTree:VBF
 - BkgTree:Sherpa $\gamma\gamma$ +jets, RevID, RevIso
 - Ratio: 1:1
- optimization
 - scan BDT cut value
 - get threshold value for VBF tight and loose category
 - plot signal and bkgshape in each category

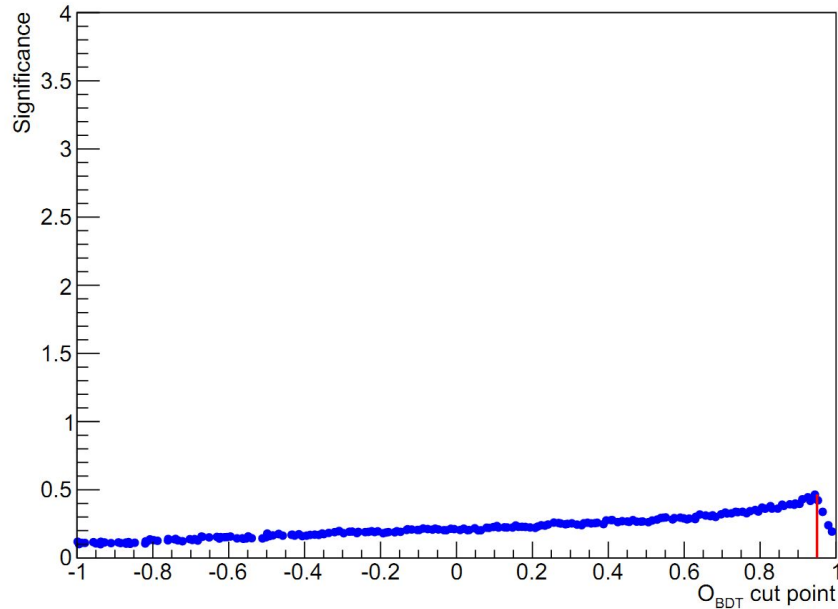
BDT value

- obtain weight xml file after training
- calculate BDT value with xml file
- plot BDT value of signal and bkg samples

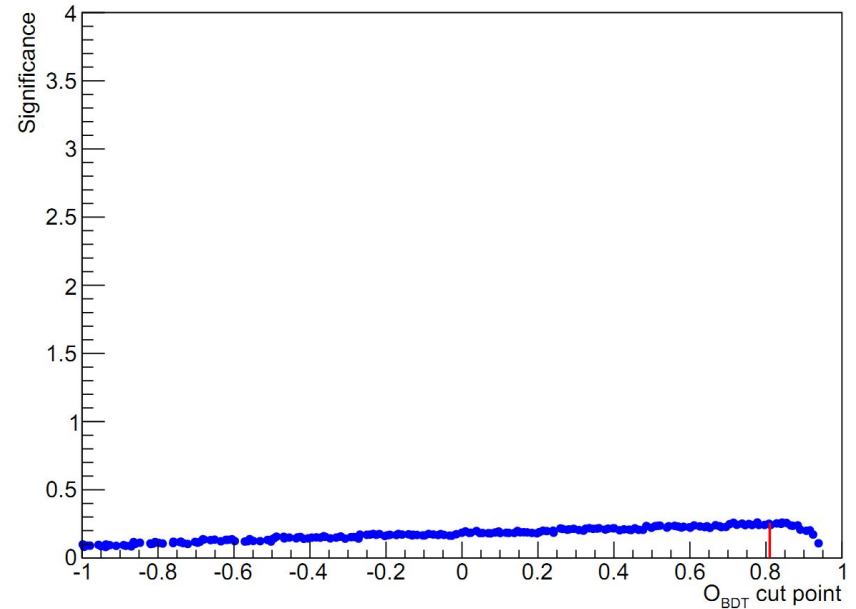


category optimization

Tight VBF Category

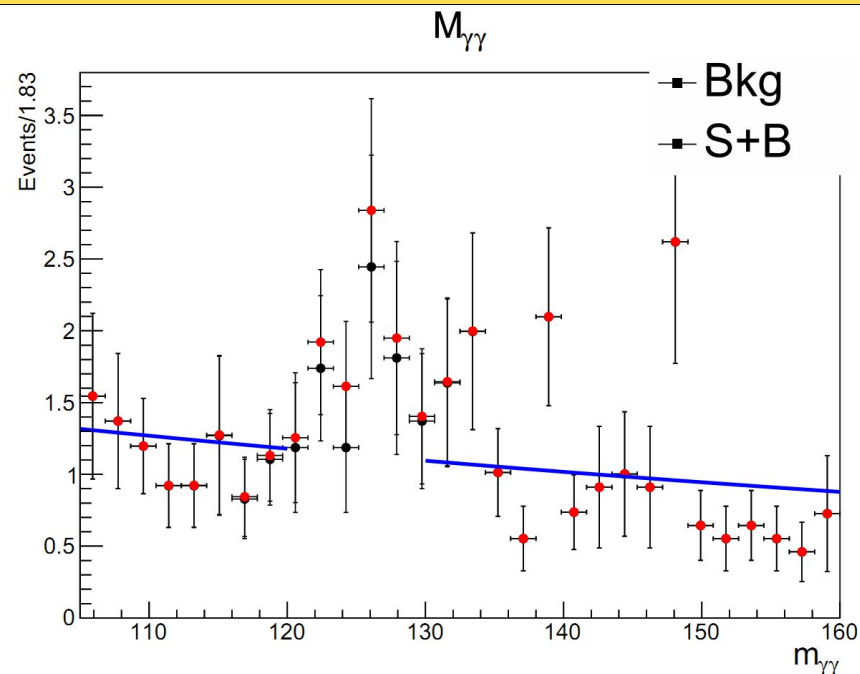
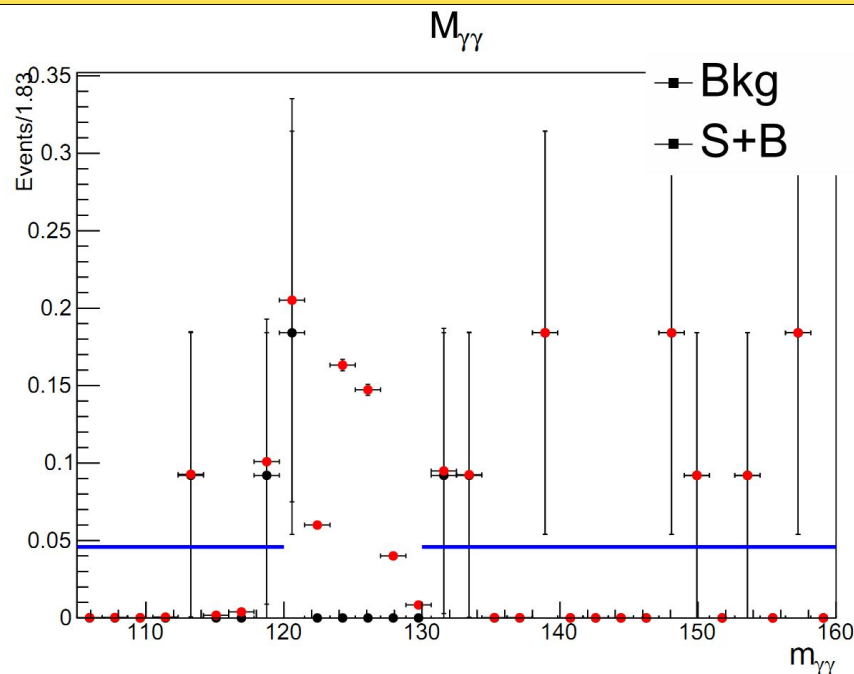


Loose VBF Category



- significance VS BDT cut value
- tight [0.95,1],0.565 σ
- loose [0.81,0.95],0.326 σ
- combined 0.652 σ

fit sideband



- left is VBF tight ,right is VBF loose
- fit sideband with exponential
- get the bkg events number
- calculate bkg number in [122.5,127.5]GeV
- red point is ggF+VBF+bkg,black point is bkg
- in loose category ,bkg has a peak near 125GeV
- BDT is correlated to $m_{\gamma\gamma}$?

event number after BDT cut

signal region : $m_{\gamma\gamma}$ [122.5,127.5]GeV

	VBF tight Category	VBF loose category	all VBF category
ggF	0.069	0.359	0.428
VBF	0.296	0.629	0.925
bkg	0.122	3.152	3.274
significance	0.565	0.326	0.652

$$\sigma_{VBF} = \sqrt{2 \times ((N_{VBF} + N_{ggF} + N_{Background}) \times \ln(1 + \frac{N_{VBF}}{N_{ggF} + N_{Background}}) - N_{VBF})}$$

- when s+b is not large enough ,this formular may not work
- toy throwing is ongoing

review 8TeV result

8TeV
20.3/fb

Events	Cut-based loose	Cut-based tight	MVA loose	MVA tight
VBF signal	2.6	3.5	3.96	3.22
ggF	1.35	1.05	2.59	0.72
Fitted background	34.8	10.6	40.1	6.7
VBF purity	65.6	76.6	60.0	81.6
VBF significance	0.43	0.98	0.60	1.11
Combined significance	1.07		1.26	

13TeV
4/fb

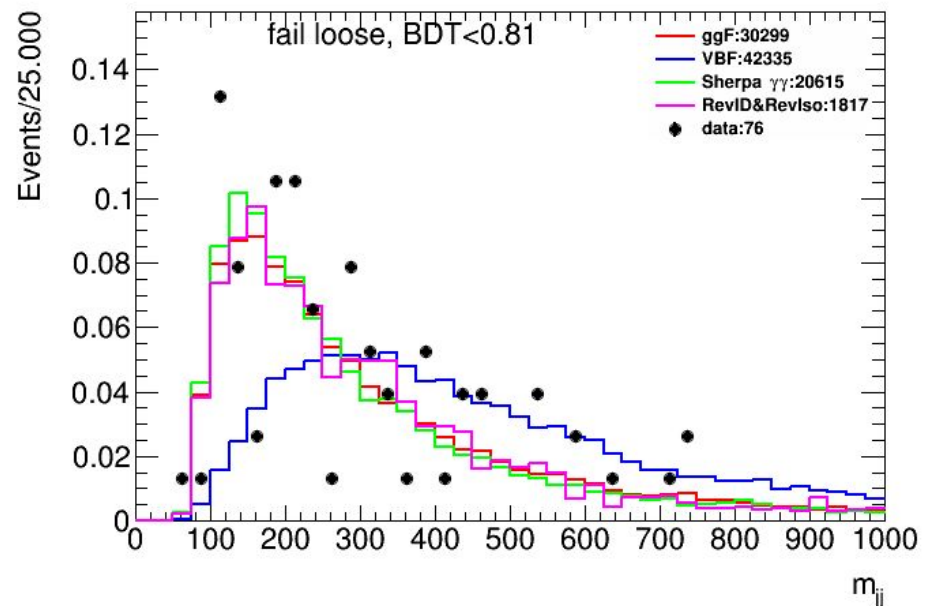
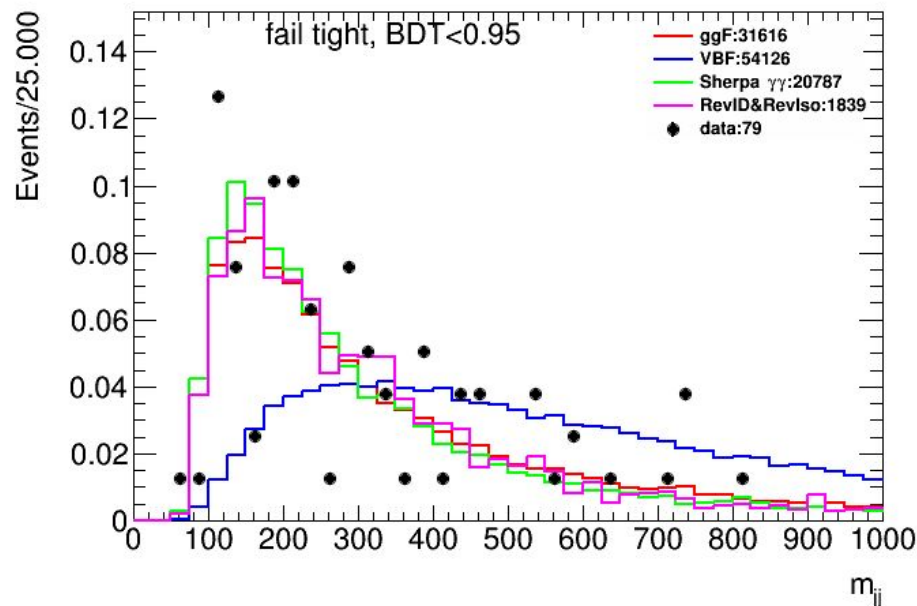
	MVA loose	MVA tight
VBF signal	0.629	0.296
ggF	0.359	0.069
fitted background	33.035	1.343
VBF purity	0.647	0.814
VBF significance	0.326	0.565
Combined significance	0.652	

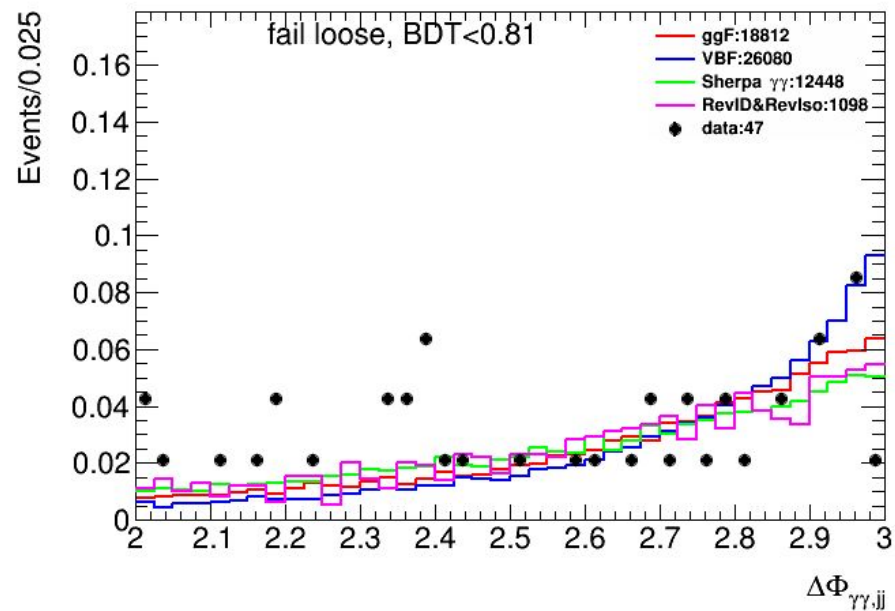
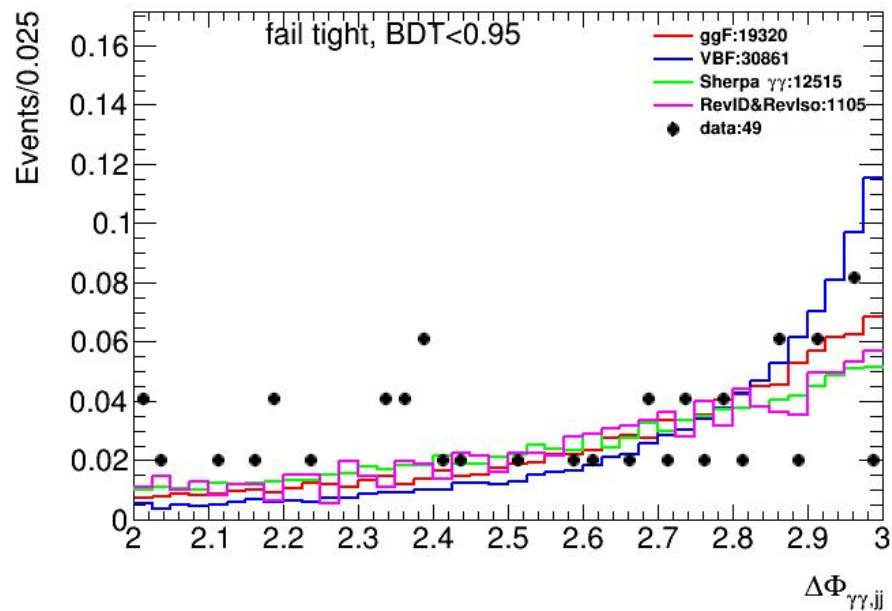
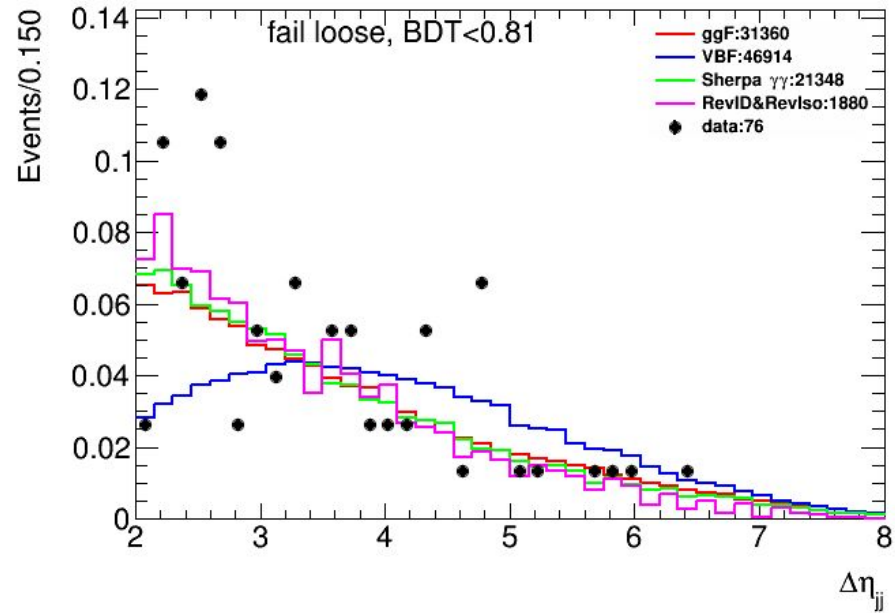
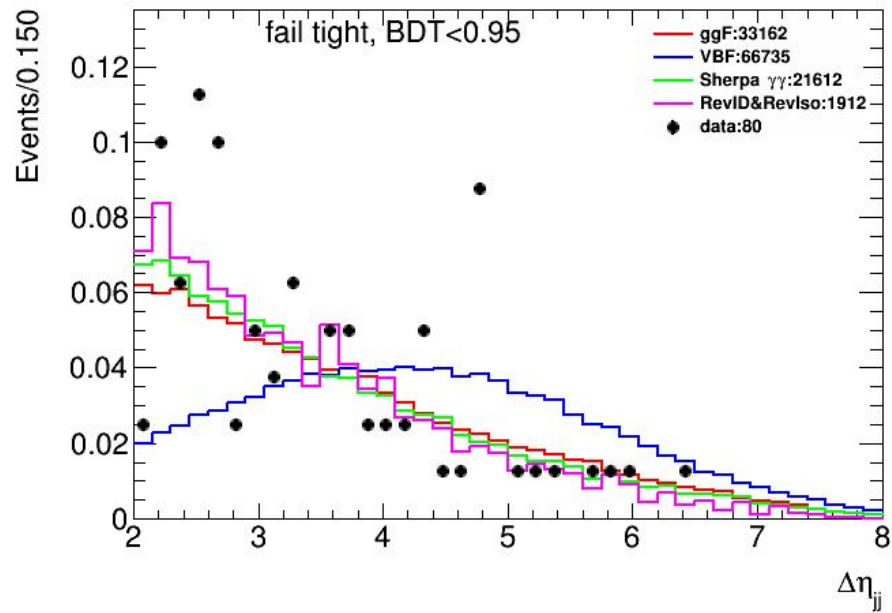
13TeV
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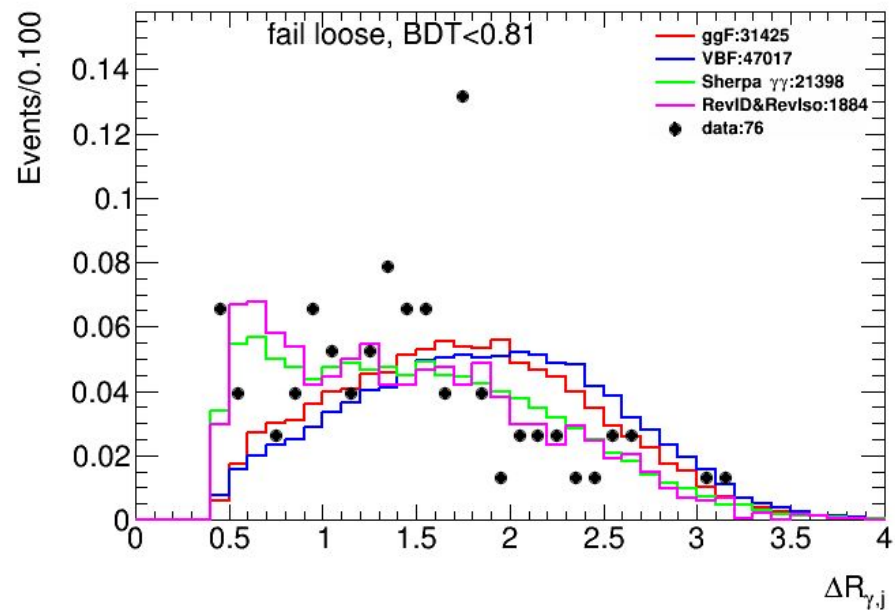
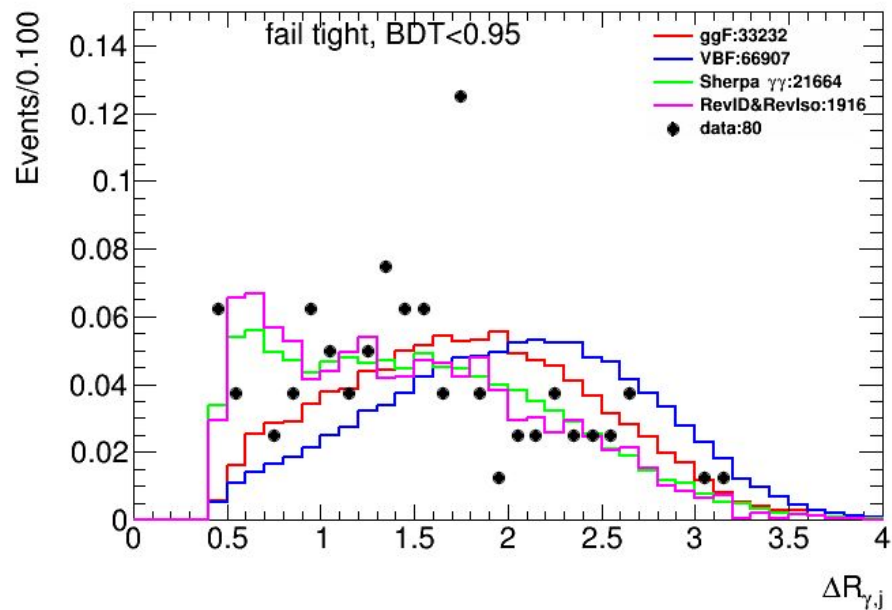
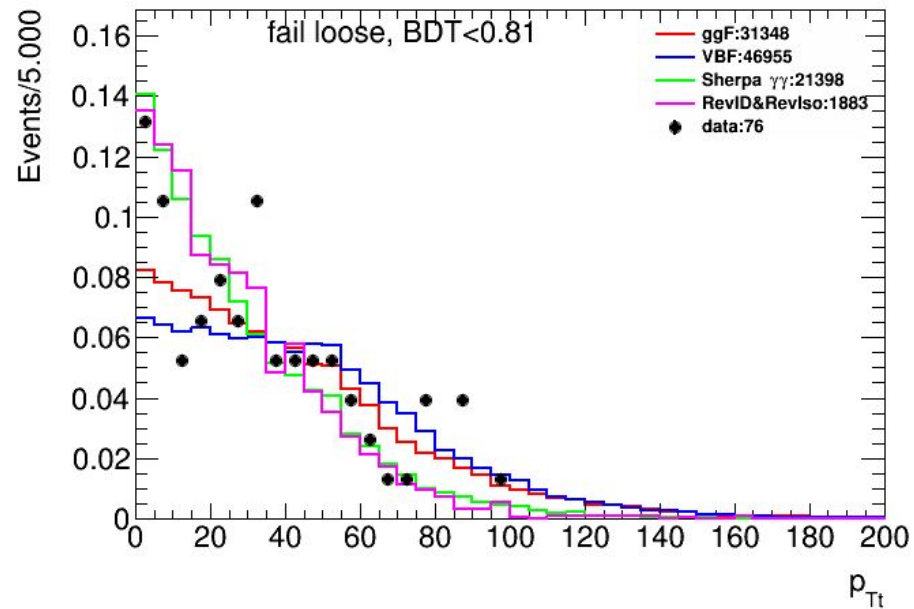
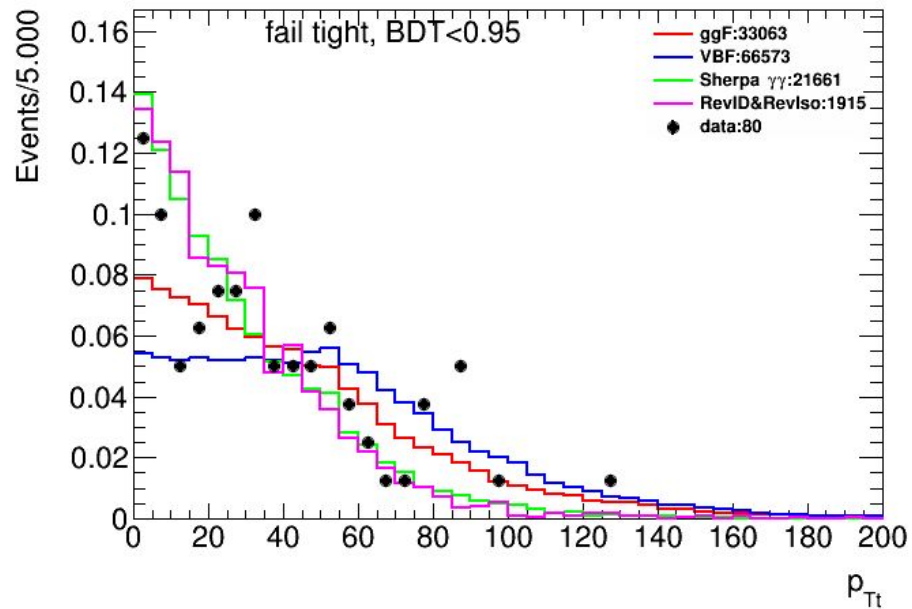
VBF signal	1.572	0.740
ggF	0.898	0.173
fitted background	82.588	3.380
VBF purity	0.647	0.814
VBF significance	0.516	0.900
Combined significance	1.030	

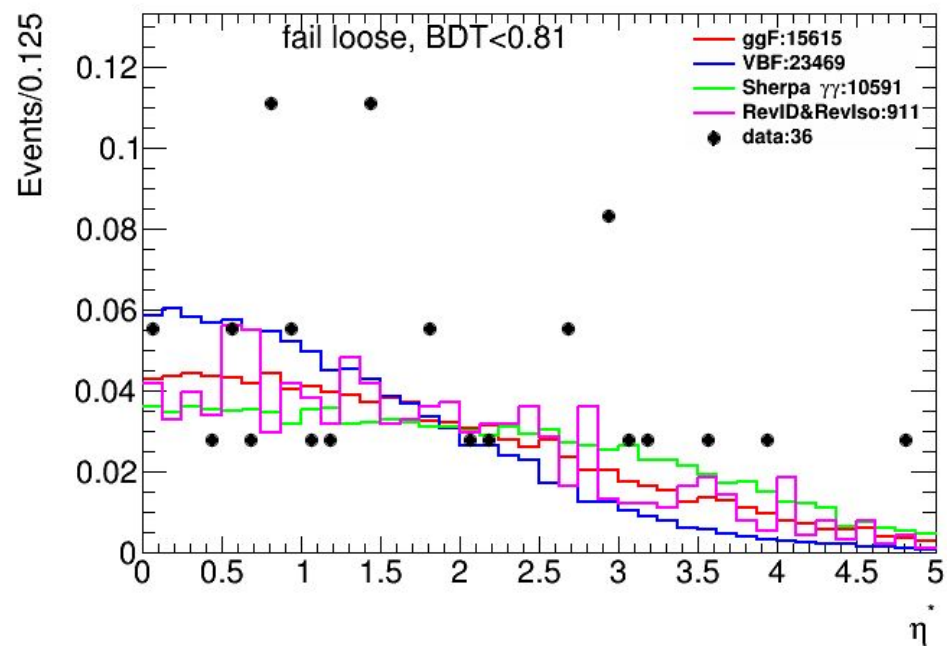
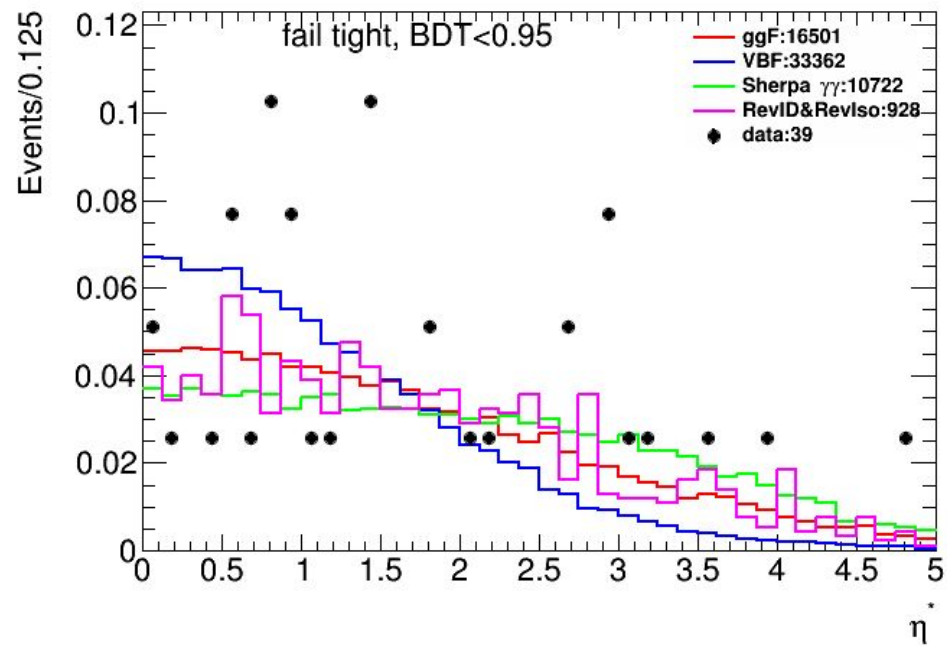
variable modeling validation

- make sure our background variable modeling (Sherpa + RevIso&RevID) is consistent with real background
- plot variables of samples failing BDT tight/loose selection
- 85 pb⁻¹ data
 - 2jets : 179 events
 - VBF preselection : 80 events
 - BDT : no events in tight ,4 events in loose









other issues

- I find a very stupid mistake.....

$$\sigma_{VBF} = \sqrt{2 \times (N_{VBF} + N_{ggF} + N_{Background}) \times \ln \left(1 + \frac{N_{VBF}}{N_{ggF} + N_{Background}} \right) - N_{VBF}}$$

- this formula is copied from ATL-COM-PHYS-2013-076
- but it is wrong.....the right one should be

$$\sigma_{VBF} = \sqrt{2 \times ((N_{VBF} + N_{ggF} + N_{Background}) \times \ln(1 + \frac{N_{VBF}}{N_{ggF} + N_{Background}}) - N_{VBF})}$$

- which is consistent with Jin's code
- so the cut-base number counting result showed last week should be wrong ,the correct one is in next page

significance ---number counting

- signal region: $m_{\gamma\gamma}$ [121,131]GeV

	tight	weighted	loose	weighted	all	weighted
ggF	3990	1.022	10804	2.767	14794	3.789
VBF	45411	1.937	66780	2.849	112194	4.786
Sherpa	141	12.98	667	61.40	808	74.37
Rev	14	1.208	55	4.746	69	5.954

$$\sigma_{VBF} = \sqrt{2 \times ((N_{VBF} + N_{ggF} + N_{Background}) \times \ln(1 + \frac{N_{VBF}}{N_{ggF} + N_{Background}}) - N_{VBF})}$$

	tight	loose	all
sigma	0.486648	0.340871	0.5945

compatible with throwing toys!

to do list

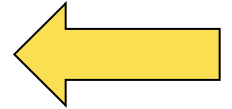
- MVA configuration optimization
- 25ns data
 - another $\sim 85\text{pb}^{-1}$
 - 276262-276954

formula derivation--copy from Cowan

Cowan, Cranmer, Gross, Vitells, arXiv:1007.1727, EPJC 71 (2011) 1554

Distribution of q_0 in large-sample limit

Assuming approximations valid in the large sample (asymptotic) limit, we can write down the full distribution of q_0 as



$$f(q_0|\mu') = \left(1 - \Phi\left(\frac{\mu'}{\sigma}\right)\right) \delta(q_0) + \frac{1}{2} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{q_0}} \exp\left[-\frac{1}{2} \left(\sqrt{q_0} - \frac{\mu'}{\sigma}\right)^2\right]$$

The special case $\mu' = 0$ is a “half chi-square” distribution:

$$f(q_0|0) = \frac{1}{2} \delta(q_0) + \frac{1}{2} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{q_0}} e^{-q_0/2}$$

In large sample limit, $f(q_0|0)$ independent of nuisance parameters; $f(q_0|\mu')$ depends on nuisance parameters through σ .

formula derivation

Cowan, Cranmer, Gross, Vitells, arXiv:1007.1727, EPJC 71 (2011) 1554

Cumulative distribution of q_0 , significance

From the pdf, the cumulative distribution of q_0 is found to be

$$F(q_0|\mu') = \Phi\left(\sqrt{q_0} - \frac{\mu'}{\sigma}\right)$$

The special case $\mu' = 0$ is

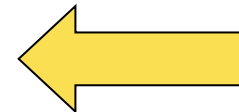
$$F(q_0|0) = \Phi(\sqrt{q_0})$$

The p -value of the $\mu = 0$ hypothesis is

$$p_0 = 1 - F(q_0|0)$$

Therefore the discovery significance Z is simply

$$Z = \Phi^{-1}(1 - p_0) = \sqrt{q_0}$$



formula derivation

$s/\sqrt{b}$ for expected discovery significance

For large $s + b$, $n \rightarrow x \sim \text{Gaussian}(\mu, \sigma)$, $\mu = s + b$, $\sigma = \sqrt{s + b}$.

For observed value x_{obs} , p -value of $s = 0$ is $\text{Prob}(x > x_{\text{obs}} | s = 0)$,

$$p_0 = 1 - \Phi\left(\frac{x_{\text{obs}} - b}{\sqrt{b}}\right)$$

Significance for rejecting $s = 0$ is therefore

$$Z_0 = \Phi^{-1}(1 - p_0) = \frac{x_{\text{obs}} - b}{\sqrt{b}}$$

Expected (median) significance assuming signal rate s is

$$\text{median}[Z_0 | s + b] = \frac{s}{\sqrt{b}} \quad \leftarrow$$

formula derivation

Better approximation for significance

Poisson likelihood for parameter s is

$$L(s) = \frac{(s+b)^n}{n!} e^{-(s+b)}$$

For now
no nuisance
params.

To test for discovery use profile likelihood ratio:

$$q_0 = \begin{cases} -2 \ln \lambda(0) & \hat{s} \geq 0, \\ 0 & \hat{s} < 0. \end{cases} \quad \lambda(s) = \frac{L(s, \hat{\hat{\theta}}(s))}{L(\hat{s}, \hat{\theta})}$$

So the likelihood ratio statistic for testing $s = 0$ is

$$q_0 = -2 \ln \frac{L(0)}{L(\hat{s})} = 2 \left(n \ln \frac{n}{b} + b - n \right) \quad \text{for } n > b, \quad 0 \text{ otherwise}$$

formula derivation

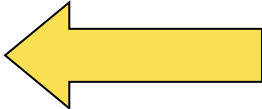
Approximate Poisson significance (continued)

For sufficiently large $s + b$, (use Wilks' theorem), 

$$Z = \sqrt{2 \left(n \ln \frac{n}{b} + b - n \right)} \quad \text{for } n > b \text{ and } Z = 0 \text{ otherwise.}$$

To find $\text{median}[Z|s]$, let $n \rightarrow s + b$ (i.e., the Asimov data set):

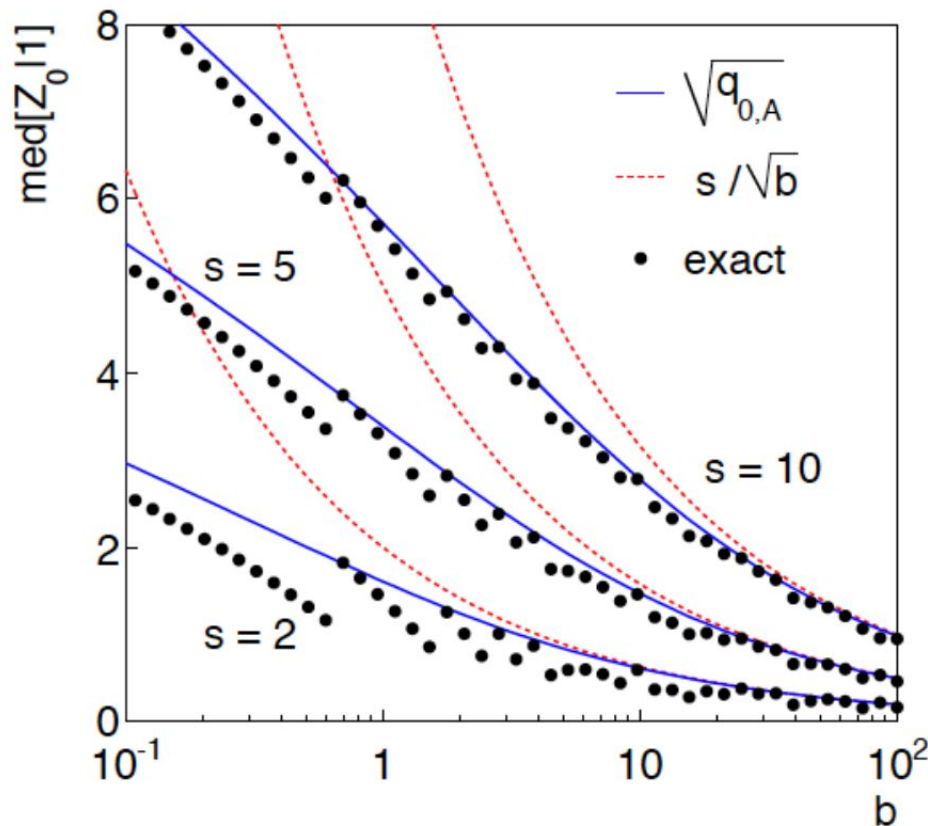
$$Z_A = \sqrt{2 \left((s + b) \ln \left(1 + \frac{s}{b} \right) - s \right)}$$

This reduces to $s/\sqrt{b}$ for $s \ll b$. 

comparison

$n \sim \text{Poisson}(s+b)$, median significance,
assuming s , of the hypothesis $s = 0$

CCGV, EPJC 71 (2011) 1554, arXiv:1007.1727



“Exact” values from MC,
jumps due to discrete data.

Asimov $\sqrt{q_{0,A}}$ good approx.
for broad range of s, b .

$s/\sqrt{b}$ only good for $s \ll b$.