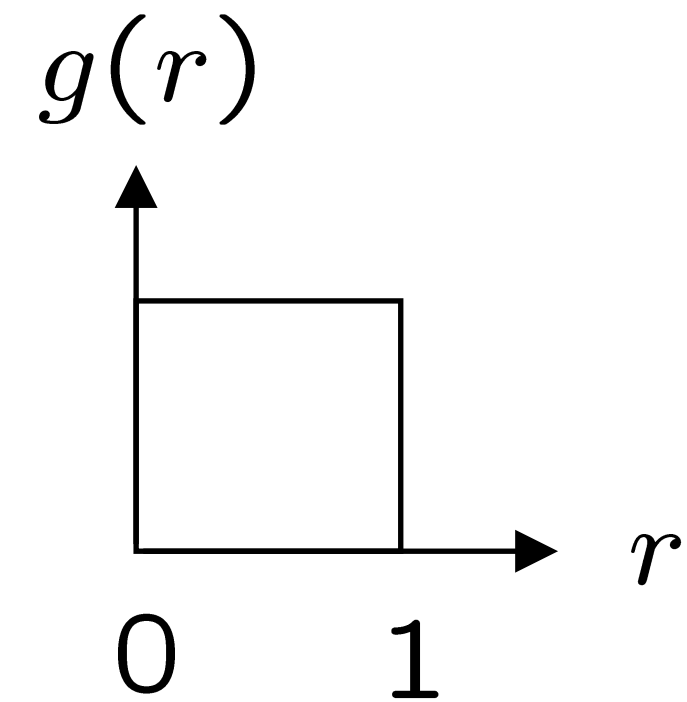


Monte Carlo Method

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The Monte Carlo method

- What it is: a numerical technique for calculating probabilities and related quantities using sequences of random numbers.
- The usual steps:
 - (1) Generate sequence r_1, r_2, \dots, r_m uniform in $[0, 1]$.
 - (2) Use this to produce another sequence x_1, x_2, \dots, x_n distributed according to some pdf $f(x)$ in which we are interested (x can be a vector)
- MC generated values = “simulated data”, “simulation”
=> use for testing statistical procedures



Random number generators

- Goal: generate uniformly distributed values in $[0, 1]$.
- Tools/algorithms to generate random numbers are called “random number generator”
- There are many algorithms to do that, such as multiplicative linear congruential generator (MLCG), Mersenne twister algorithm, etc.,
- They are implemented, and we can use directly:
 - C++ standard library(stdlib): rand() function can generate random numbers.
 - ROOT: TRandom, TRandom2, TRandom3, where TRandom3 gives much better results [F. James, Comp. Phys. Comm. 60 (1990) 111]

The transformation method

- Given r_1, r_2, \dots, r_m uniform in $[0, 1]$, find x_1, x_2, \dots, x_n that follow $f(x)$ by finding a suitable transformation $x(r)$

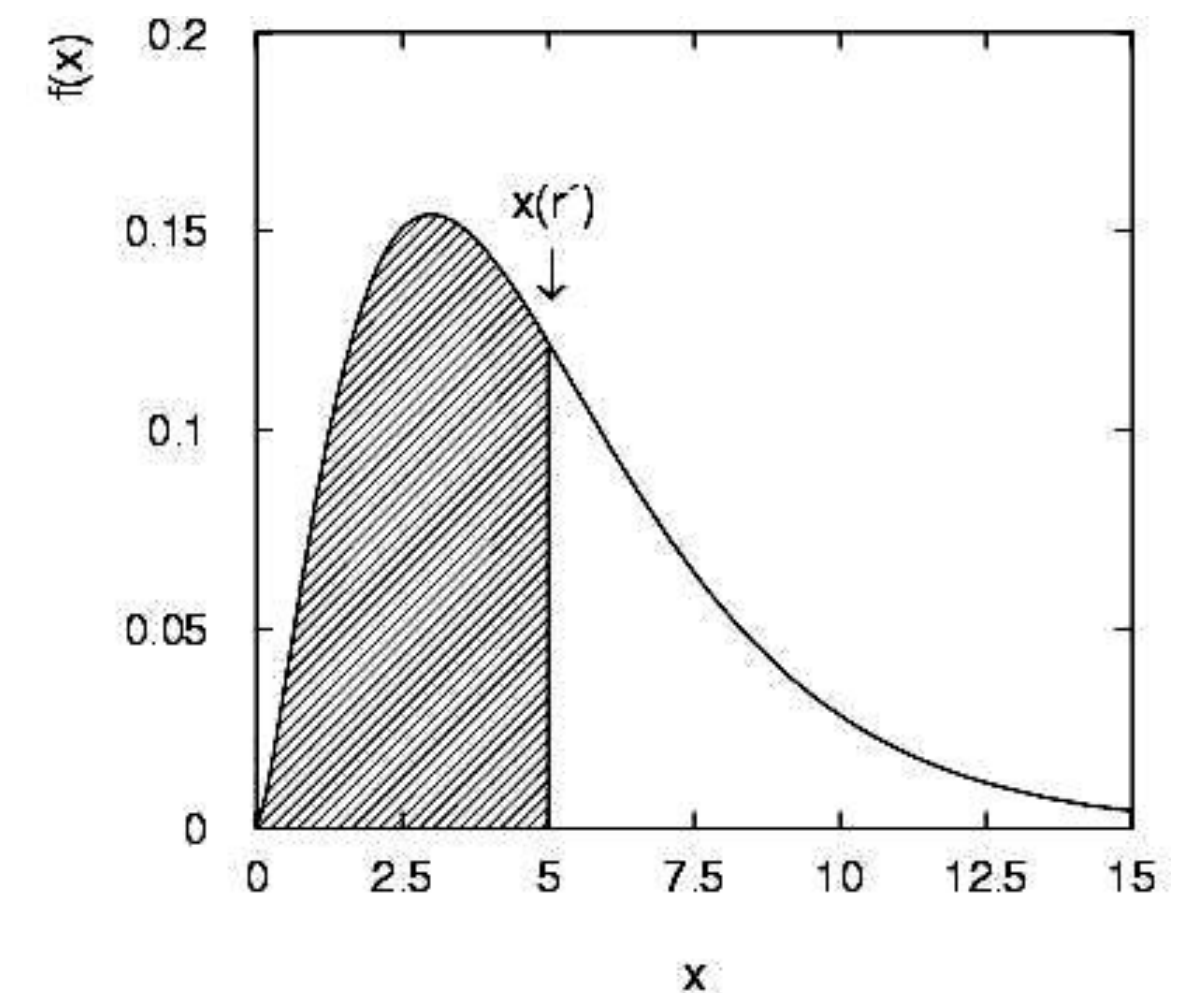
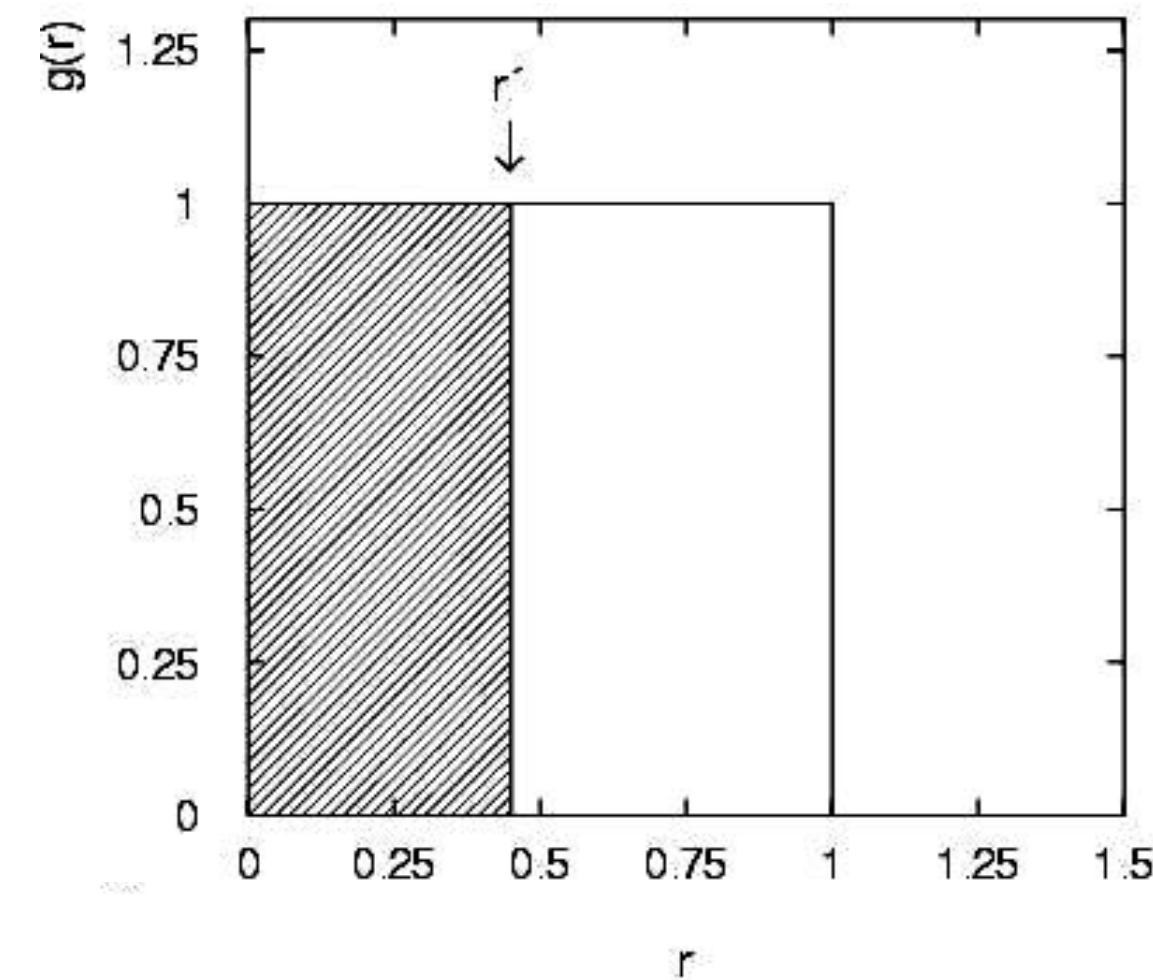
- Remind: previously, we learned given pdf $f(x)$ and function $a(x)$, ask the pdf of a : $g(a)$.

- Here, is given pdf $f(x)$ and $g(r)$, ask the function $x(r)$:

Require: $P(r \leq r') = P(x \leq x(r'))$

$$\text{i.e. } \int_{-\infty}^{r'} g(r) dr = r' = \int_{-\infty}^{x(r')} f(x') dx' = F(x(r'))$$

- That is, set $F(x) = r$ and solve for $x(r)$



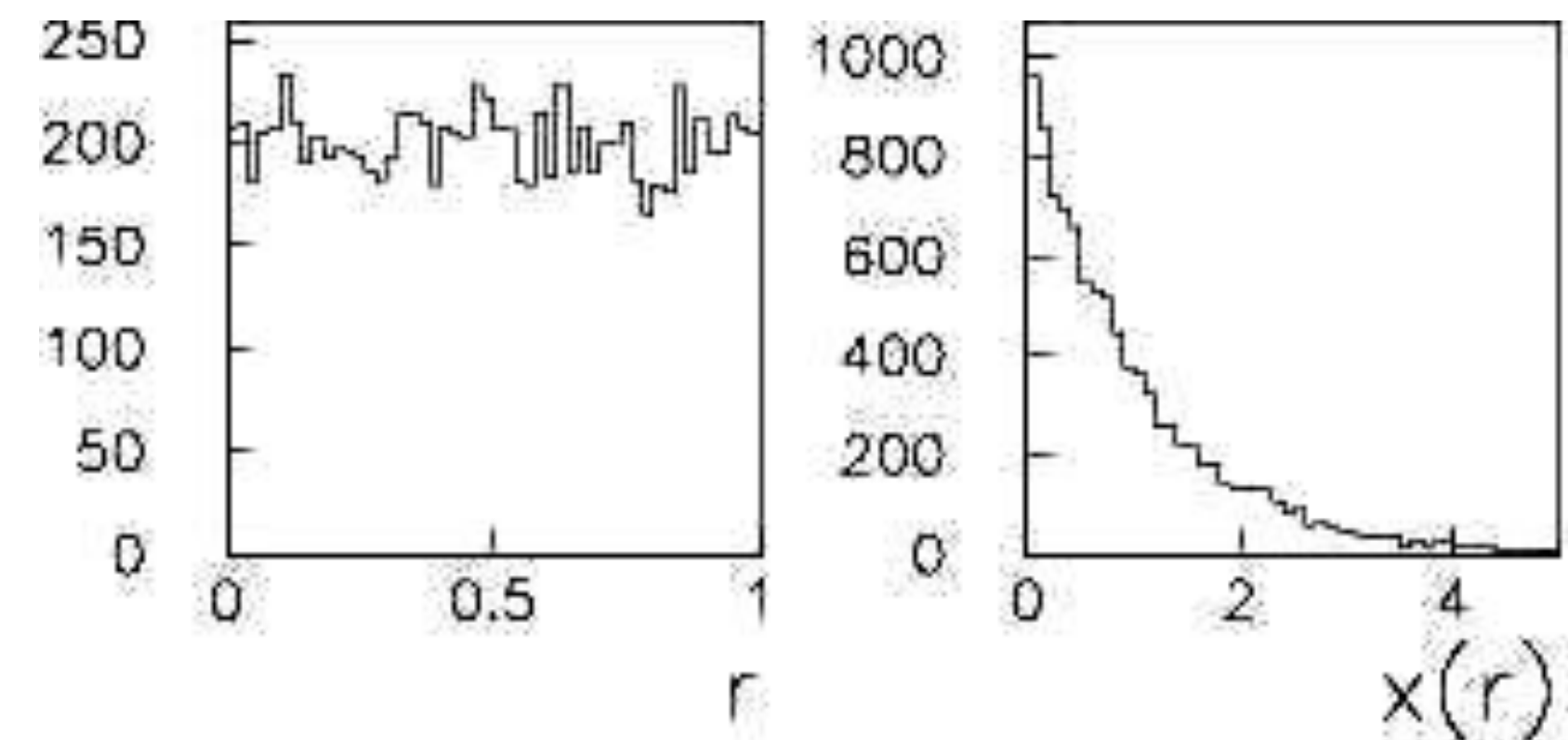
Example of the transformation method

- Exponential pdf: $f(x; \xi) = \frac{1}{\xi} e^{-x/\xi} \quad (x \geq 0)$

Set $\int_0^x \frac{1}{\xi} e^{-x'/\xi} dx' = r$ and solve for $x(r)$.

→ $x(r) = -\xi \ln(1 - r)$ ($x(r) = -\xi \ln r$ works too.)

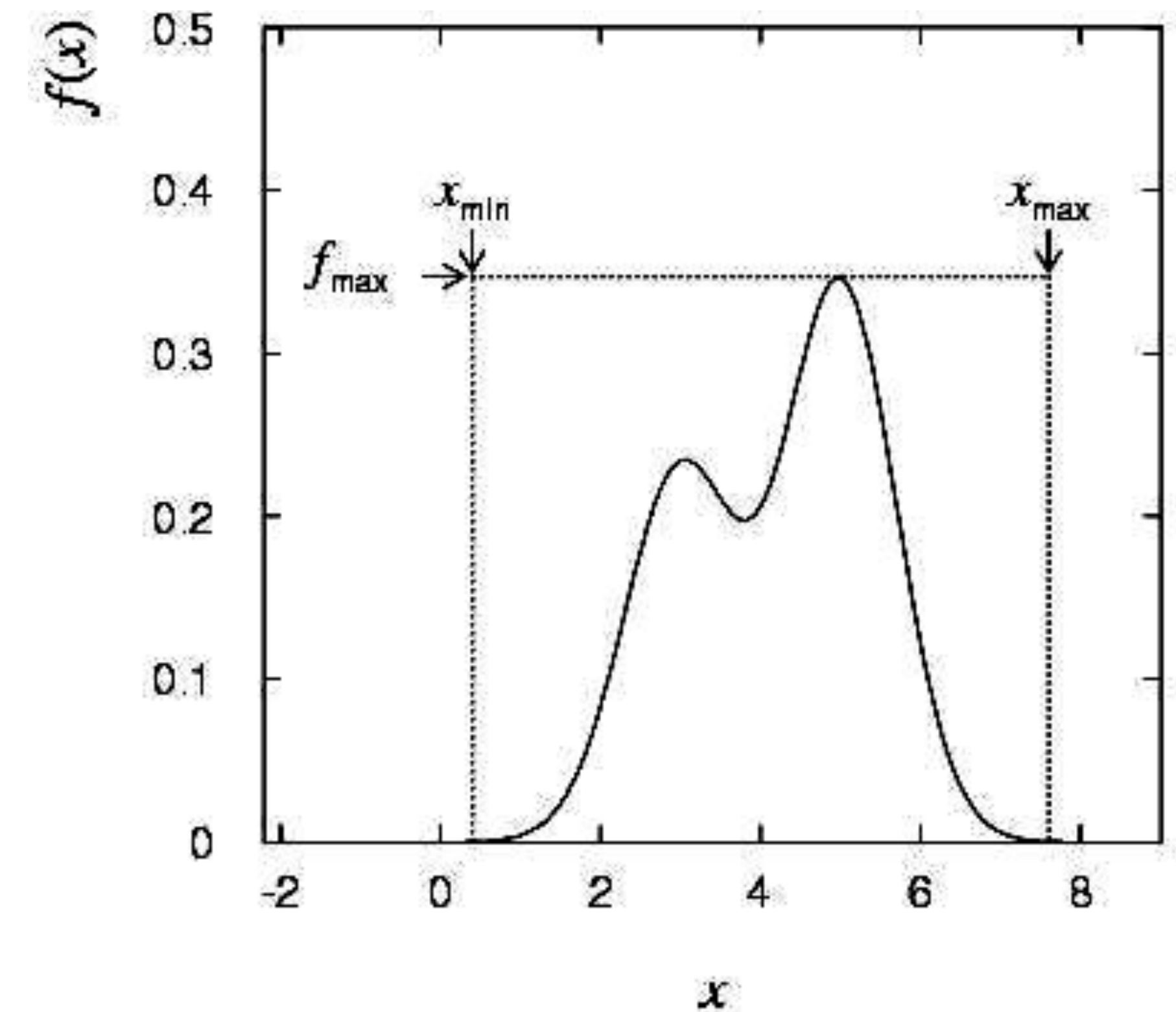
- One has to solve for $x(r)$ analytically to use this method, not practical for arbitrary $f(x)$



The acceptance-rejection method

- (1) Generate a random number x , uniform in $[x_{\min}, x_{\max}]$, i.e. $x = x_{\min} + r_1 (x_{\max} - x_{\min})$, r_1 is uniform in $[0, 1]$.
- (2) Generate a 2nd independent random number u uniformly distributed between 0 and f_{\max} , i.e. $u = r_2 f_{\max}$
- (3) If $u < f(x)$, then accept x , if not, reject x and repeat
- Practical for $f(x)$ with complicated shapes.

Enclose the pdf in a box

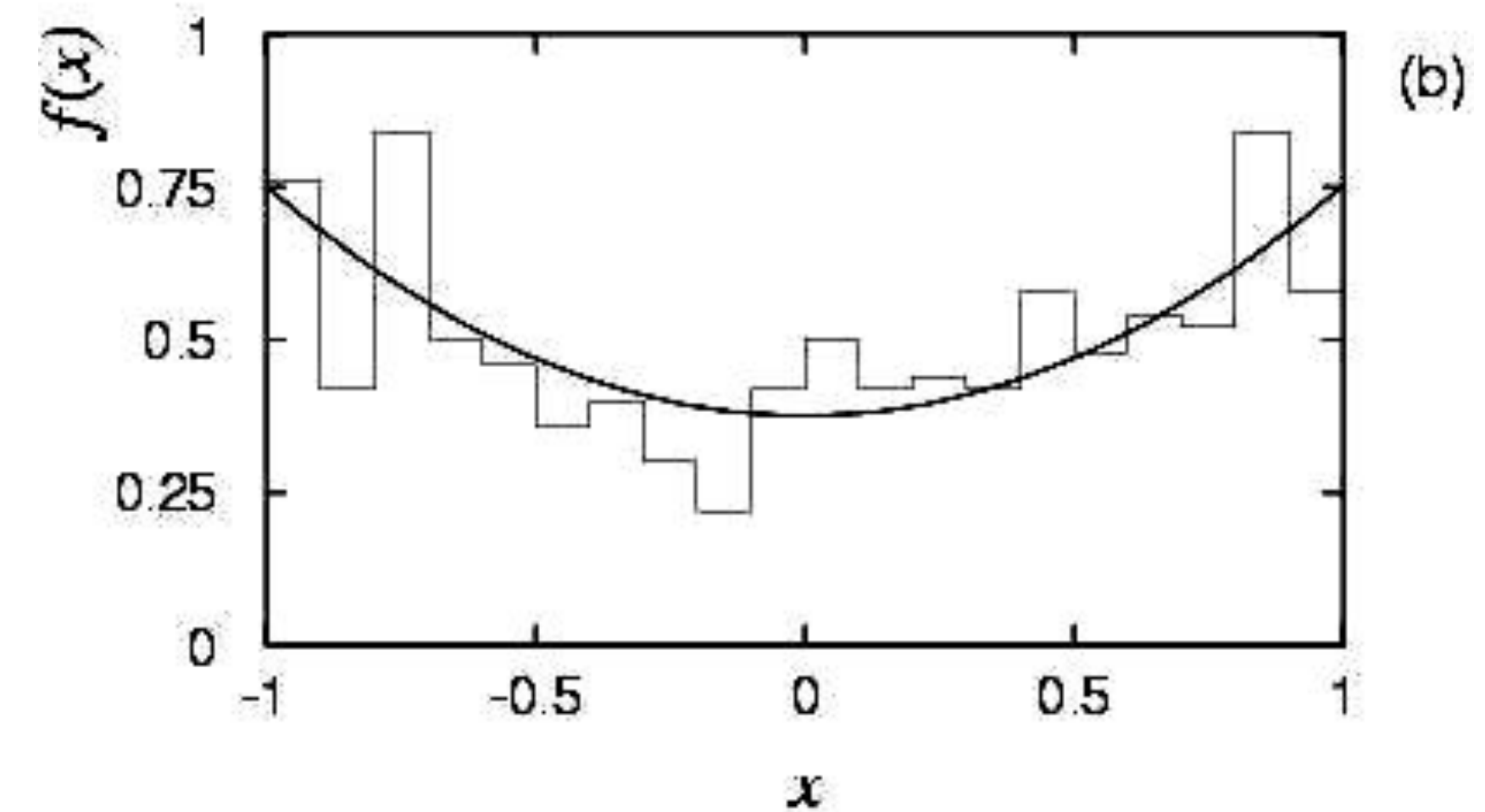
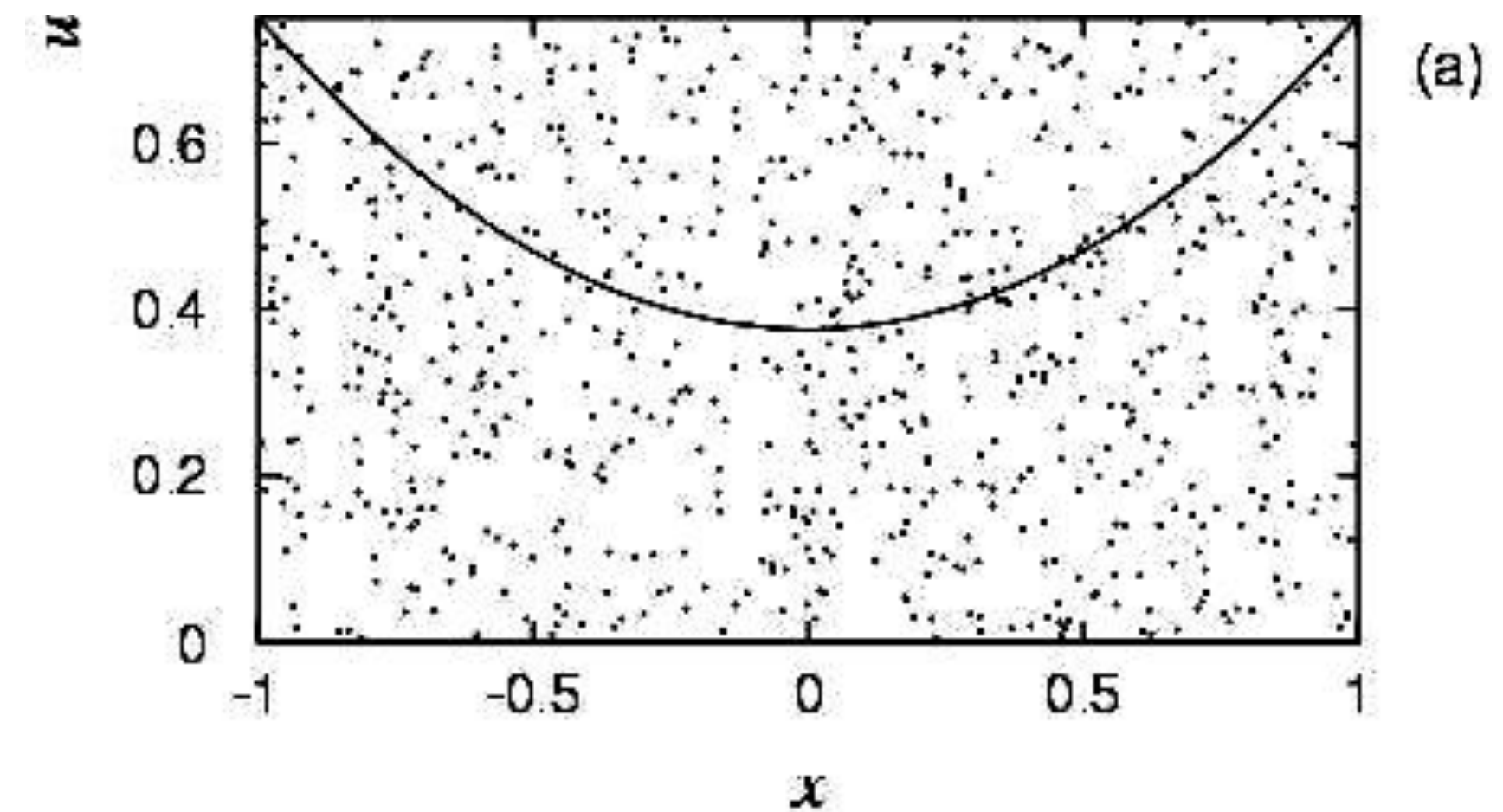


Example with acceptance-rejection method

$$f(x) = \frac{3}{8}(1 + x^2)$$

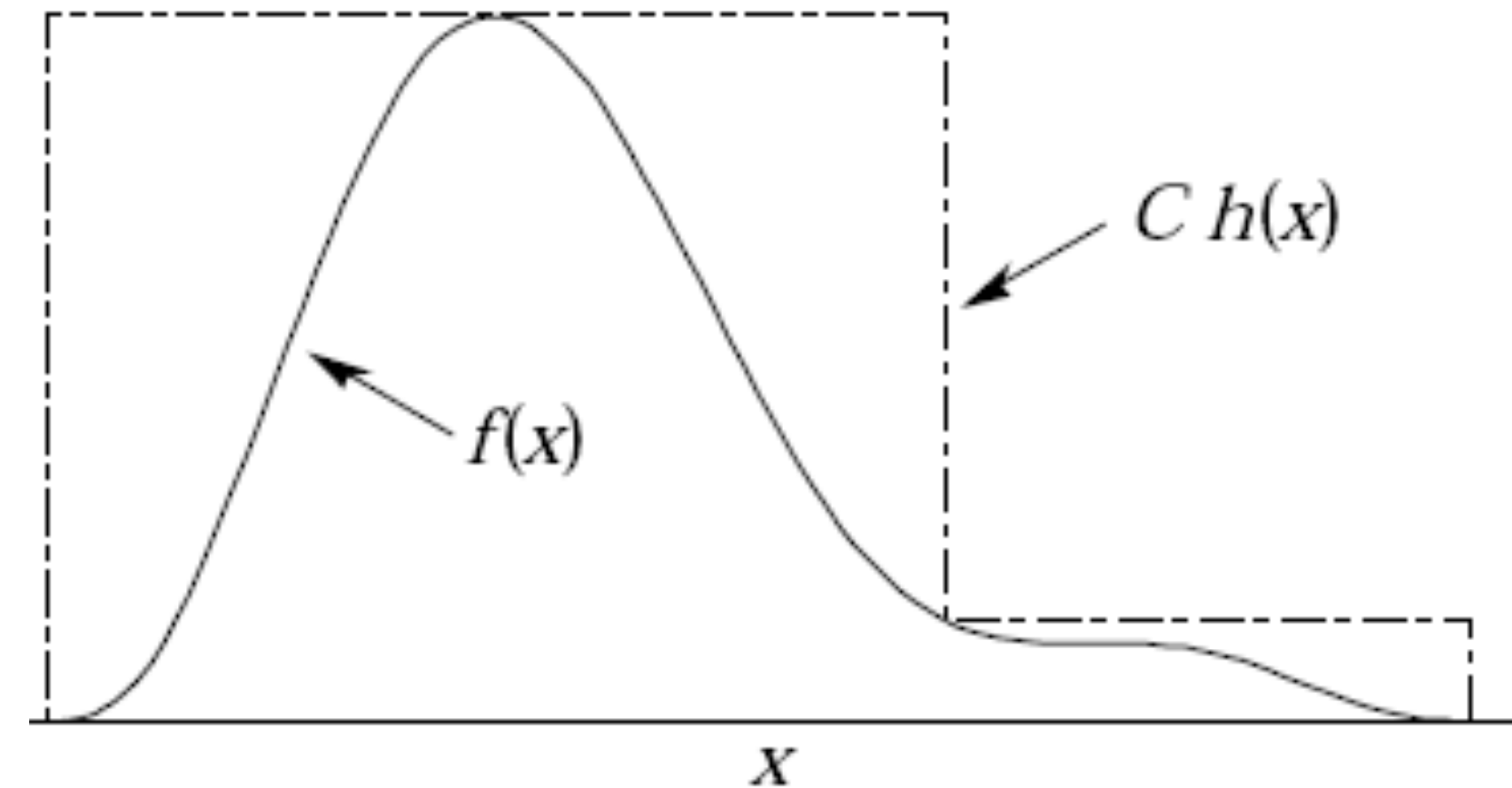
$$(-1 \leq x \leq 1)$$

- Generate random number x in $[-1, 1]$
- Generate another random number u in $[0, 3/4]$
- If $u \leq f(x)$, take the x , and fill it in the histogram. If not, drop the x .



Improving efficiency of the acceptance-rejection method

- The fraction of accepted points is equal to the fraction of the box's area under the curve.
- For very peaked distributions, the fraction maybe very low, thus the algorithm may be slow.
- Improve by enclosing the pdf $f(x)$ in a curve $C h(x)$ that conforms to $f(x)$ more closely, where $h(x)$ is a pdf from which we can generate random values and C is a constant:
 - (1) Generate random variable x according to $C h(x)$ by other method, such as “transformation method”.
 - (2) Generate another random number u uniformly distributed between 0 and $C h(x)$
 - (3) If $u < f(x)$, take accept x .



Moving to the practical programming training.

Backup slides, Probability Density Functions.

Probability functions

- We will now introduce a short list of popular probability functions and pdfs.
- For each functions, we show expectation value, variance, a plot and discuss of some properties and applications.
- See also chapter on probability from pdg.lbl.gov

Binomial distribution

- Consider N independent experiments (Bernoulli trials):
 - outcome of each is “success” or “failure”
 - probability of “success” on any given trial is p
- Define discrete random variable n = number of successes ($0 \leq n \leq N$)
- Probability of a specific outcome (in order), e.g. “s s f s f” is:
$$pp(1 - p)p(1 - p) = p^n(1 - p)^{N-n}$$
- But order is not important; there are $N!/[n!(N - n)!]$ ways (permutations) to get n successes in N trials, total probability for n is sum of probabilities for each permutation.

Binomial distribution

- The binomial distribution is therefore

$$f(n; N, p) = \frac{N!}{n!(N-n)!} p^n (1-p)^{N-n}$$

random variable parameters

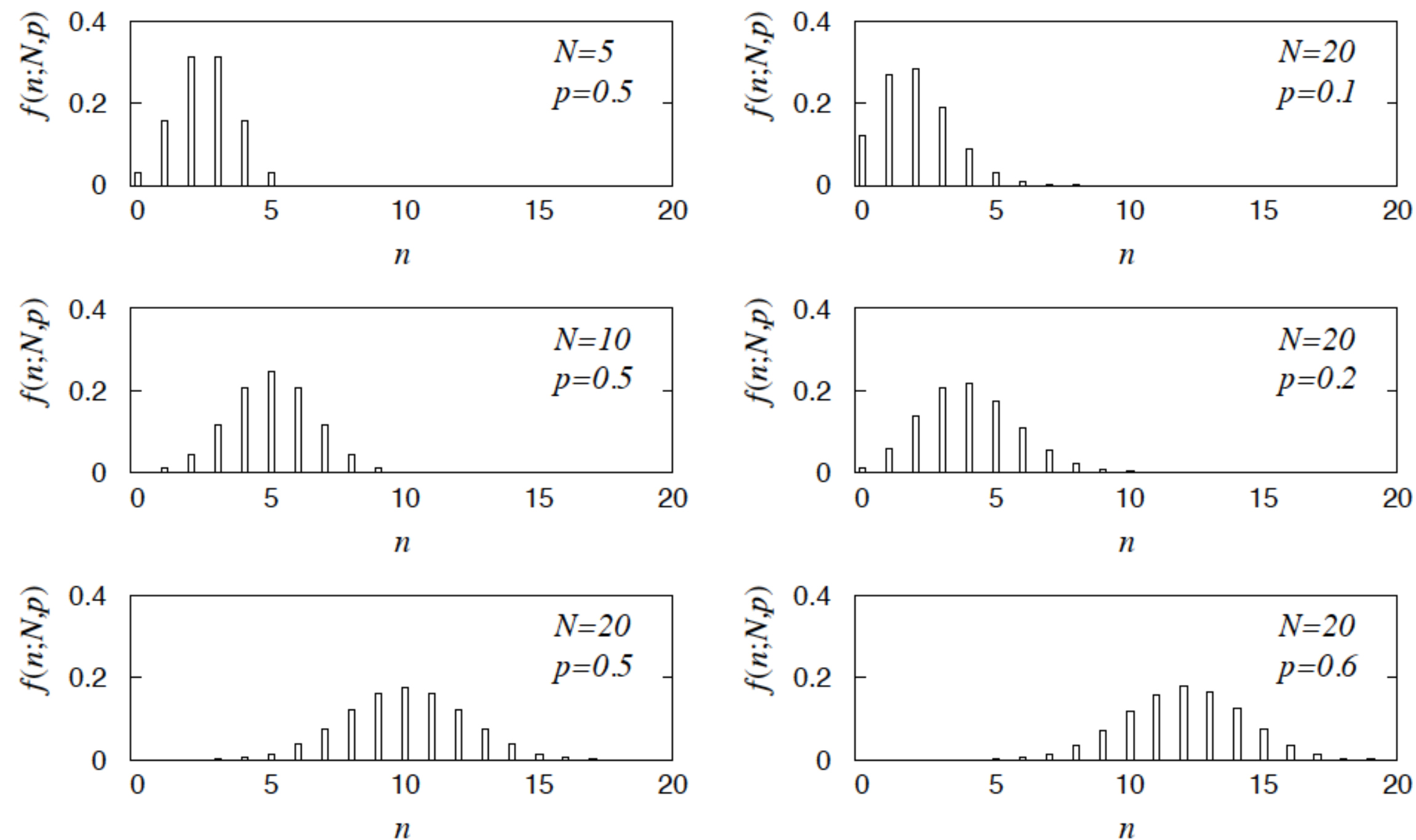
- For the expectation value and variance, we find

$$E[n] = \sum_{n=0}^N n f(n; N, p) = Np$$

$$V[n] = E[n^2] - (E[n])^2 = Np(1-p)$$

Binomial distribution

- Binomial distribution for different parameters values



- Example: observe N decays of W bosons, in which n is the number of $W \rightarrow \mu\nu$ decays (number of “success”), and p is the branching ratio (probability of “success”).

Multinomial distribution

- Like binomial, but now m outcomes instead of two, probabilities are

$$\vec{p} = (p_1, \dots, p_m), \quad \text{with } \sum_{i=1}^m p_i = 1$$

- For N trials, we want the probability to obtain:
 n_1 of outcome 1,
 n_2 of outcome 2,
 \vdots
 n_m of outcome m .

- This is the multinomial distribution for $\vec{n} = (n_1, \dots, n_m)$

$$f(\vec{n}; N, \vec{p}) = \frac{N!}{n_1! n_2! \dots n_m!} p_1^{n_1} p_2^{n_2} \dots p_m^{n_m}$$

Multinomial distribution

- Now consider outcome i as “success”, all others are “failure”.

$$E [n_i] = Np_i, \quad V [n_i] = Np_i (1 - p_i) \quad \text{for all } i$$

- Note, i (success) and NOT- i (failure), actually form a “binomial” structure
- One can also find the covariance to be
$$V_{ij} = Np_i \left(\delta_{ij} - p_j \right)$$
- Example: $\vec{n} = (n_1, \dots, n_m)$ represents a histogram with m bins, N total entries, all entries independent.

Poisson distribution

- Consider binomial n in the limit

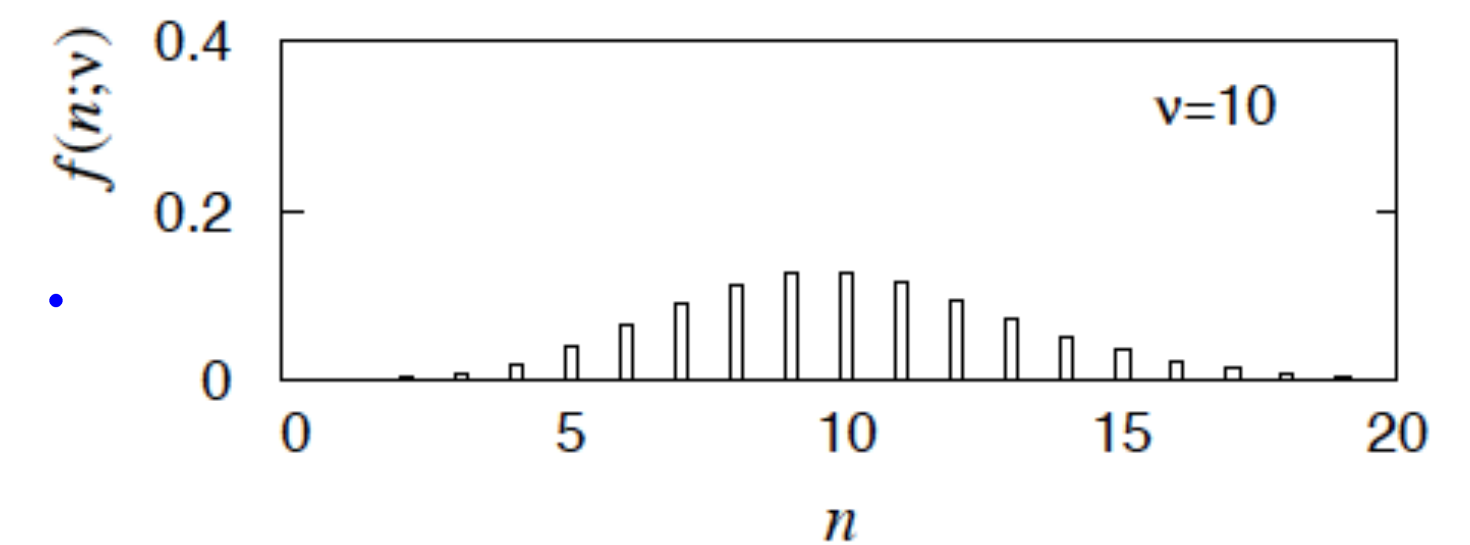
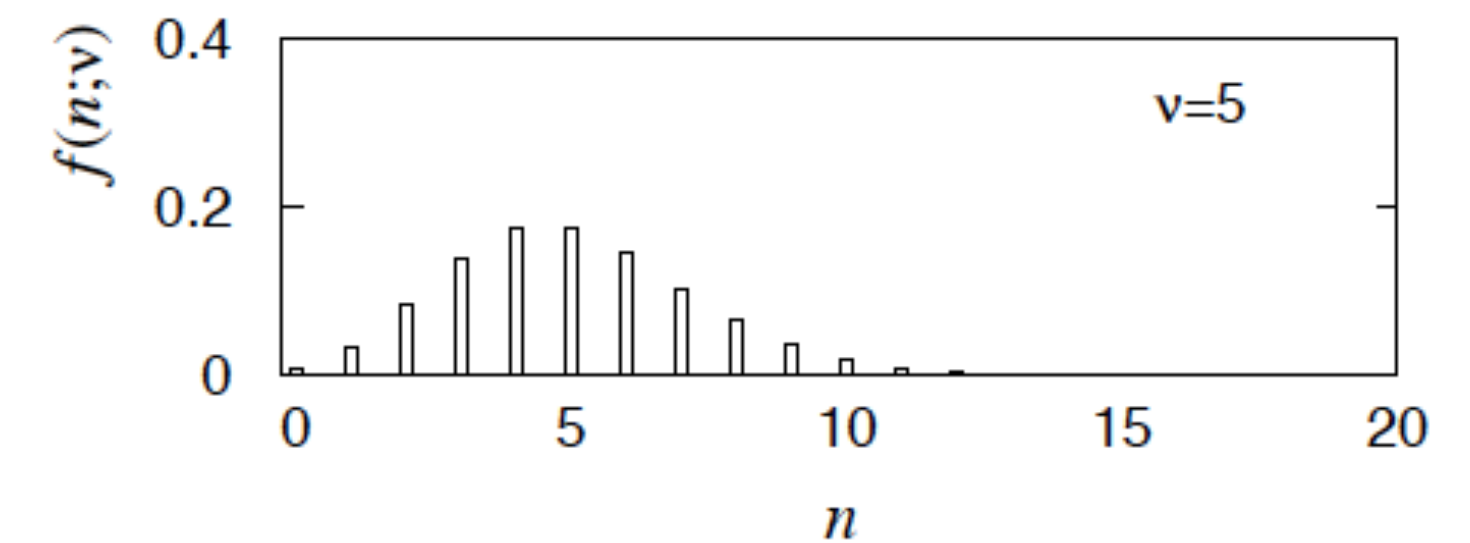
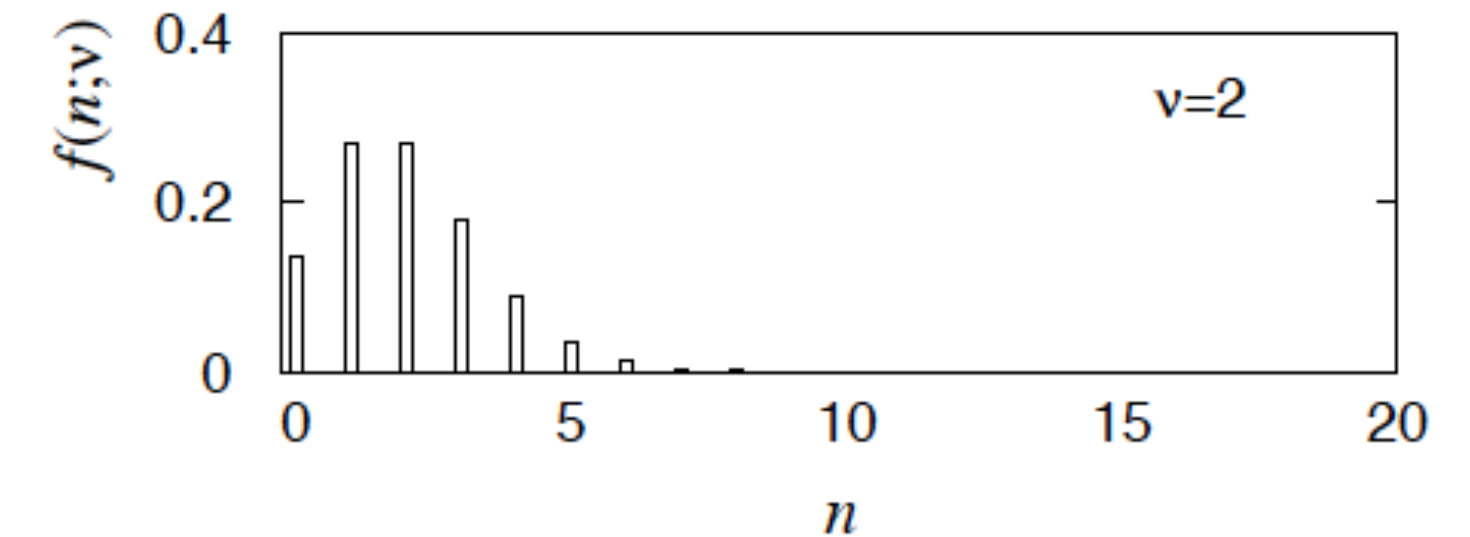
$$N \rightarrow \infty, \quad p \rightarrow 0, \quad E[n] = Np \rightarrow \nu$$

- n follows the Poisson distribution:

$$f(n; \nu) = \frac{\nu^n}{n!} e^{-\nu} \quad (n \geq 0)$$

$$E[n] = \nu, \quad V[n] = \nu$$

- Example: number of scattering events n with cross section σ found for a fixed integrated luminosity, with $\nu = \sigma \int L dt$



Uniform distribution

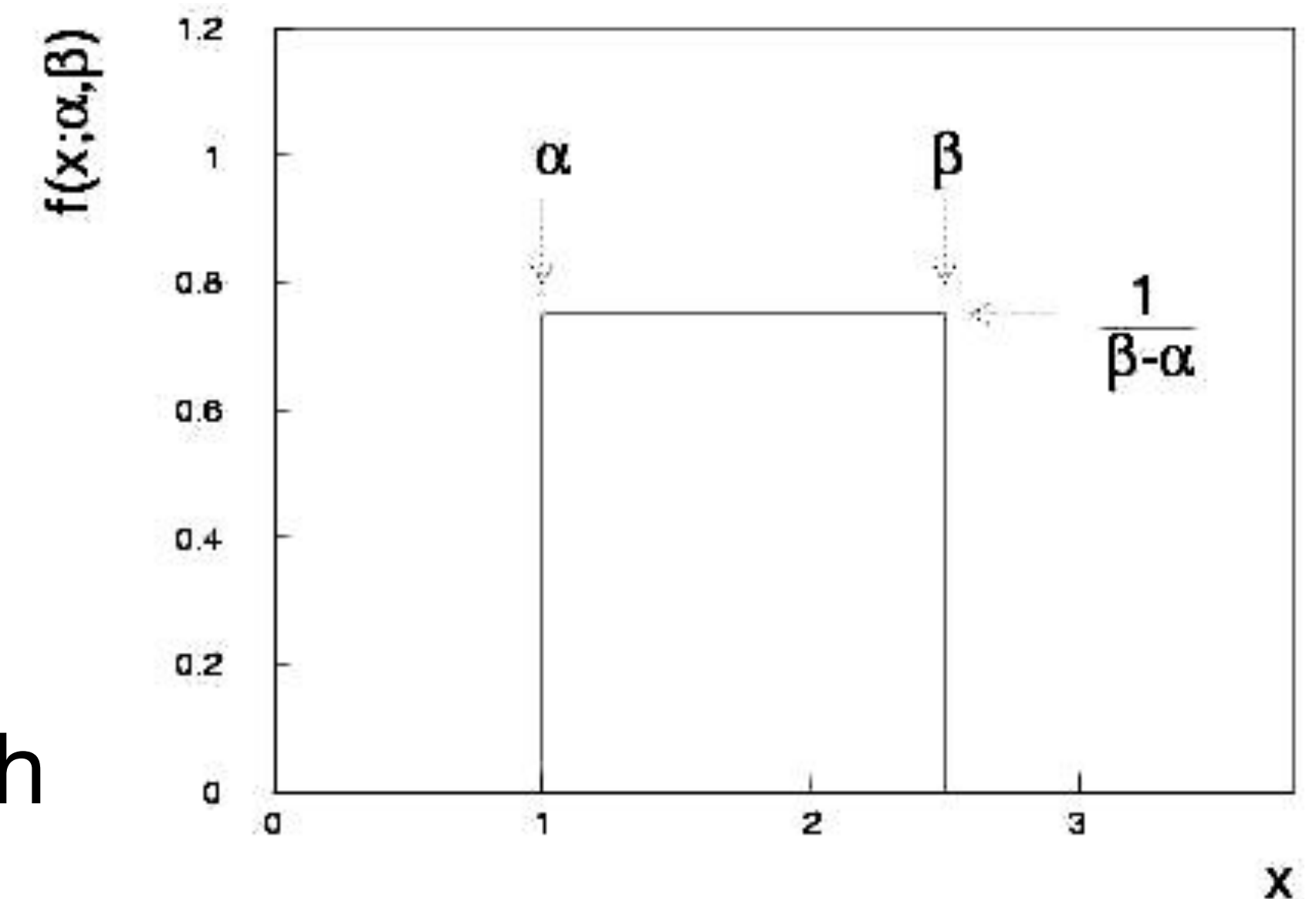
- Consider a continuous random variable with $-\infty < x < \infty$
Uniform pdf is

$$f(x; \alpha, \beta) = \begin{cases} \frac{1}{\beta - \alpha} & \alpha \leq x \leq \beta \\ 0 & \text{otherwise} \end{cases} \quad \begin{aligned} E[x] &= \frac{1}{2}(\alpha + \beta) \\ V[x] &= \frac{1}{12}(\beta - \alpha)^2 \end{aligned}$$

- Note, for any random variable x with cumulative distribution $F(x)$, $y = F(x)$ is uniform in $[0, 1]$
- Example: for $\pi^0 \rightarrow \gamma\gamma$, E_γ is uniform in $[E_{\min}, E_{\max}]$, with

$$E_{\min} = \frac{1}{2}E_\pi(1 - \beta), \quad E_{\max} = \frac{1}{2}E_\pi(1 + \beta)$$

$$\beta = v/c$$



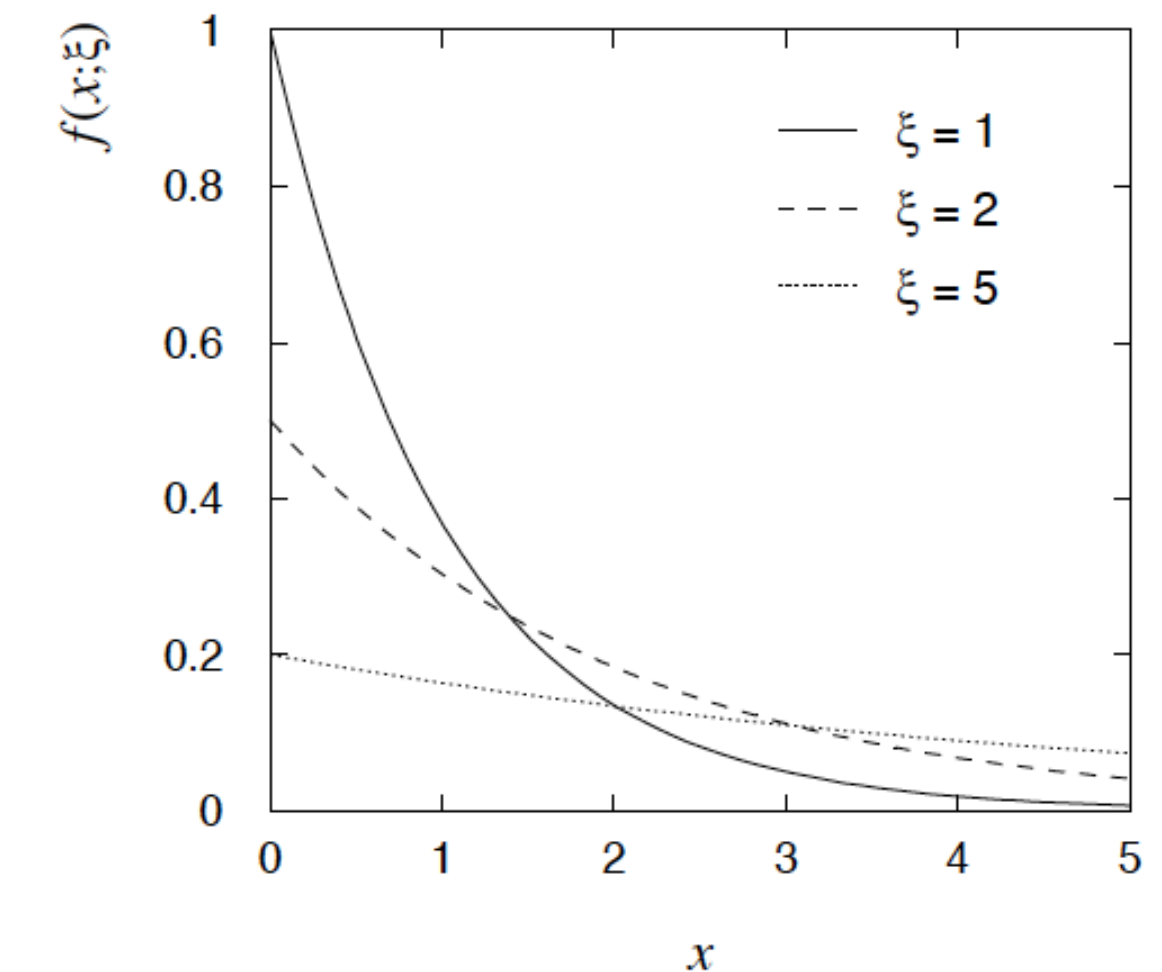
Exponential distribution

- The exponential pdf for the continuous random variable x is defined by:

$$f(x; \xi) = \begin{cases} \frac{1}{\xi} e^{-x/\xi} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

$$E[x] = \xi$$

$$V[x] = \xi^2$$



- Example: proper decay time t of an unstable particle

$$f(t; \tau) = \frac{1}{\tau} e^{-t/\tau} \quad (\tau = \text{mean lifetime})$$

Gaussian distribution

- The Gaussian (normal) pdf for a continuous random variable x is defined by

$$f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-(x-\mu)^2/2\sigma^2}$$

$$E[x] = \mu \quad (\text{N.B. often } \mu, \sigma^2 \text{ denote mean, variance of any}$$

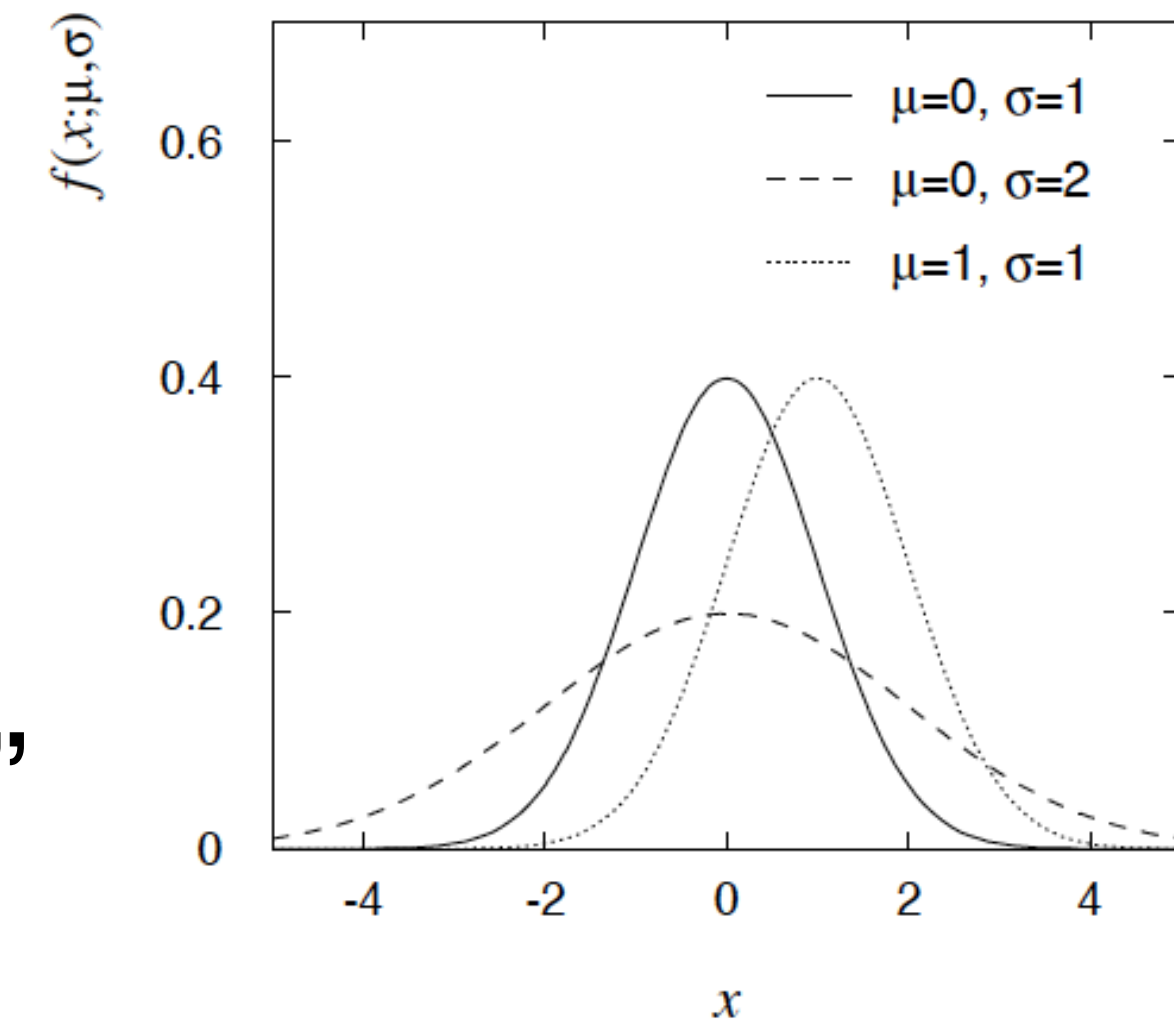
$$V[x] = \sigma^2 \quad \text{r.v., not only Gaussian.)}$$

- Special case: "standard Gaussian"

$$\mu = 0, \sigma^2 = 1$$

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad \Phi(x) = \int_{-\infty}^x \varphi(x') dx'$$

- If $y \sim \text{Gaussian with } \mu, \sigma^2$, then $x = (y - \mu)/\sigma$ follows $\varphi(x)$



Gaussian pdf and the Central Limit Theorem

- The Gaussian pdf is very useful because almost any random variable that is a sum of a large number of small contributions follows it. This follows from the Central Limit Theorem:

- For n independent random variables x_i with finite variances σ_i^2 , otherwise arbitrary pdfs, consider the sum

$$y = \sum_{i=1}^n x_i$$

- In the limit $n \rightarrow \infty$, y is a Gaussian random variable with

$$E[y] = \sum_{i=1}^n \mu_i \quad V[y] = \sum_{i=1}^n \sigma_i^2$$

- Measurement errors are often the sum of many contributions, so frequently measured values can be treated as Gaussian random variables.

Central Limit Theorem (CLT)

- The CLT can be proved using characteristic functions (Fourier transforms).
- For finite n , the theorem is approximately valid to the extent that the fluctuation of the sum is not dominated by one (or few) terms.
=> beware of measurement errors with non-Gaussian tails.
- Good example: velocity components v_x of air molecules.
- OK example: total deflection due to multiple Coulomb scattering.
(Rare large angle deflections give non-Gaussian tail.)
- Bad example: energy loss of charged particle traversing thin gas layer.

Multivariate Gaussian distribution

- Multivariate Gaussian pdf for the vector $\vec{x} = (x_1, \dots, x_n)$

$$f(\vec{x}; \vec{\mu}, V) = \frac{1}{(2\pi)^{n/2} |V|^{1/2}} \exp \left[-\frac{1}{2} (\vec{x} - \vec{\mu})^T V^{-1} (\vec{x} - \vec{\mu}) \right]$$

- $\vec{x}, \vec{\mu}$ are column vectors, $\vec{x}^T, \vec{\mu}^T$ are transpose (row) vectors

$$E[x_i] = \mu_i, \quad \text{COV}[x_i, x_j] = V_{ij}.$$

- For n=2 this is

$$f(x_1, x_2; \mu_1, \mu_2, \sigma_1, \sigma_2, \rho) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \\ \times \exp \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 - 2\rho \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) \right] \right\}$$

- where $\rho = \text{COV}[x_1, x_2] / (\sigma_1\sigma_2)$ is the correlation coefficient.

Chi-square (χ^2) distribution

- The chi-square pdf for the continuous random variable z ($z \geq 0$) is defined by

$$f(z; n) = \frac{1}{2^{n/2} \Gamma(n/2)} z^{n/2-1} e^{-z/2}$$

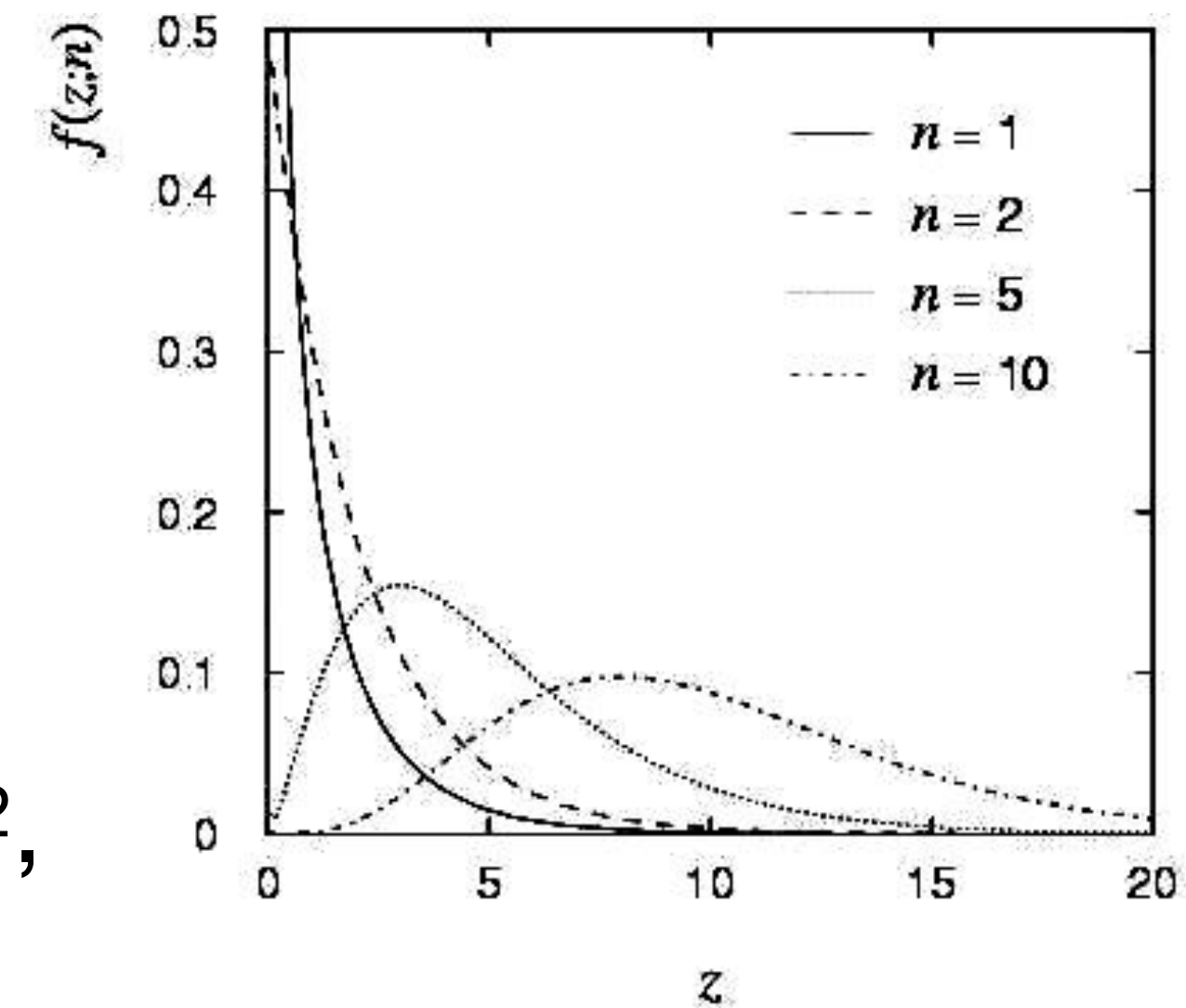
$n = 1, 2, \dots$ = number of 'degrees of freedom' (dof)

$$E[z] = n, \quad V[z] = 2n.$$

- For independent Gaussian x_i , $i = 1, \dots, n$, means μ_i , variances σ_i^2 ,

$$z = \sum_{i=1}^n \frac{(x_i - \mu_i)^2}{\sigma_i^2} \quad \text{follows } \chi^2 \text{ pdf with } n \text{ dof.}$$

- Example: goodness-of-fit test variable especially in conjunction with method of least squares.



Cauchy (Breit-Wigner) distribution

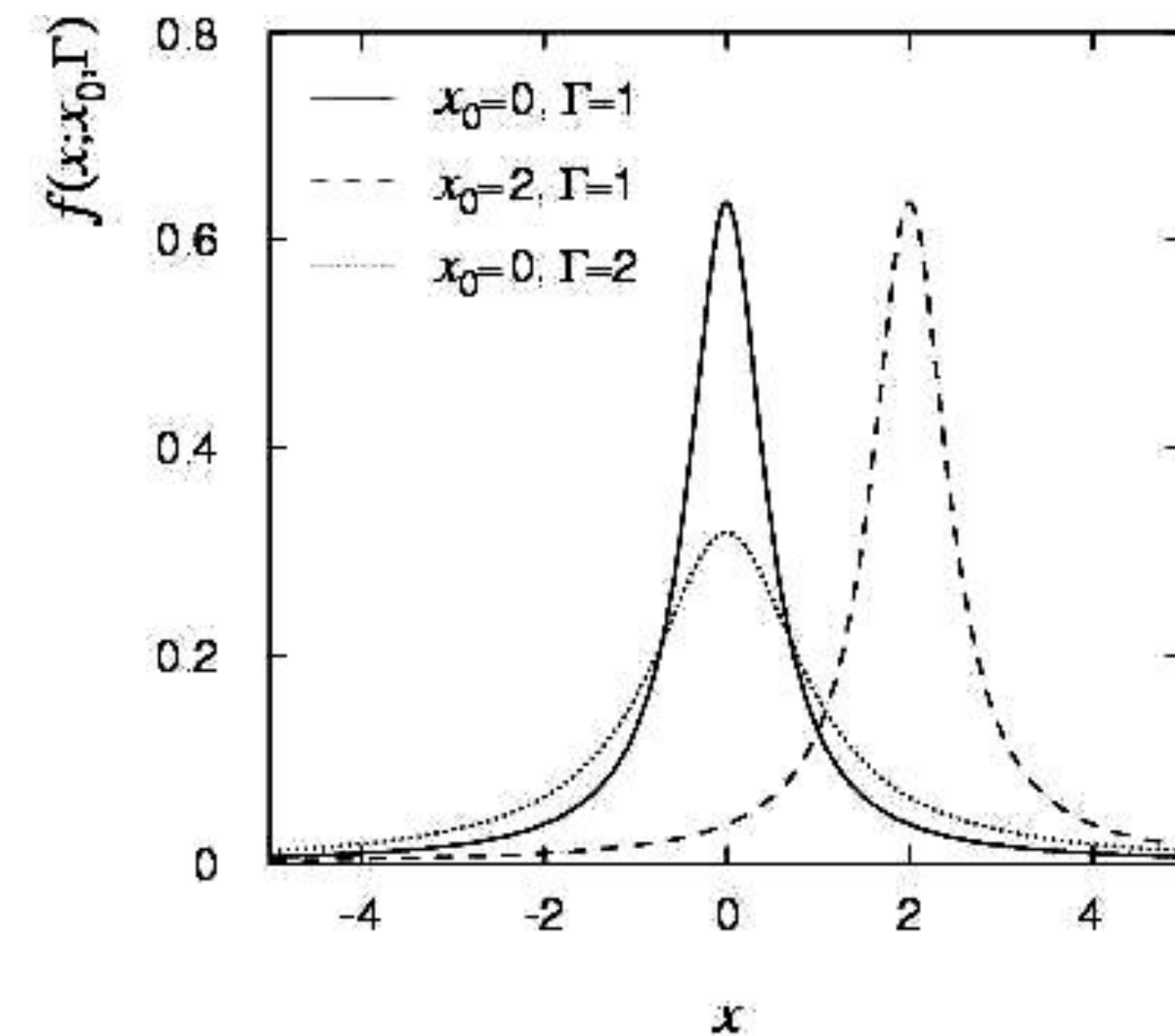
- The Breit-Wigner pdf for the continuous random variable x is defined by

$$f(x; \Gamma, x_0) = \frac{1}{\pi} \frac{\Gamma/2}{\Gamma^2/4 + (x - x_0)^2}$$

($\Gamma = 2, x_0 = 0$ is the Cauchy pdf.)

- $E[x]$ is not well defined, $V[x] \rightarrow \infty$
 $x_0 = \text{mode}$ (most probable value)
 $\Gamma = \text{full width at half maximum}$

- Example: mass of resonance particles, e.g. $Z^0, \rho, K^*, \phi^0, \dots$
- $\Gamma = \text{decay rate}$ (inverse of mean lifetime)



Landau distribution

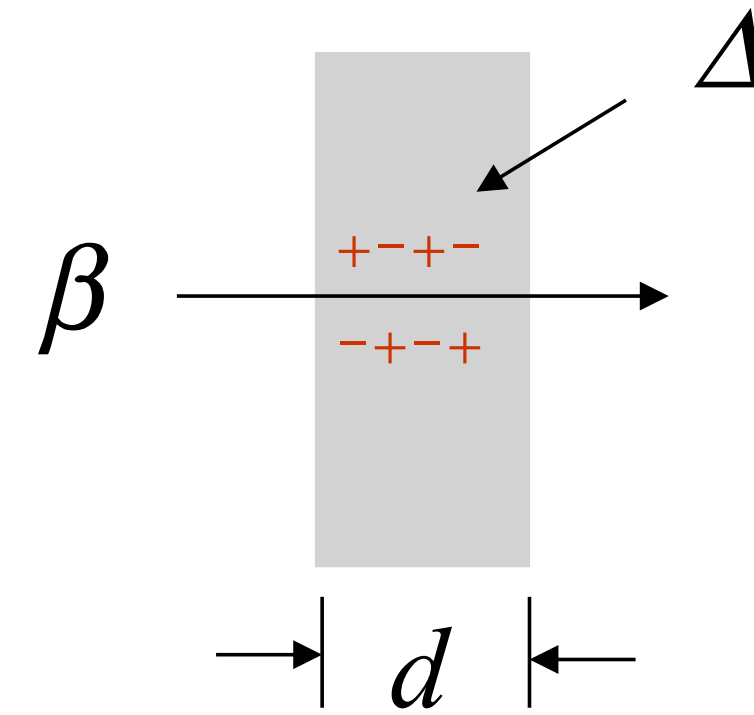
- For a charged particle with $\beta = v/c$ traversing a layer of matter of thickness d , the energy loss Δ follows the Landau pdf:

$$f(\Delta; \beta) = \frac{1}{\xi} \phi(\lambda) ,$$

$$\phi(\lambda) = \frac{1}{\pi} \int_0^\infty \exp(-u \ln u - \lambda u) \sin \pi u \, du ,$$

$$\lambda = \frac{1}{\xi} \left[\Delta - \xi \left(\ln \frac{\xi}{\epsilon'} + 1 - \gamma_E \right) \right] ,$$

$$\xi = \frac{2\pi N_A e^4 z^2 \rho \sum Z}{m_e c^2 \sum A} \frac{d}{\beta^2} , \quad \epsilon' = \frac{I^2 \exp \beta^2}{2m_e c^2 \beta^2 \gamma^2} .$$



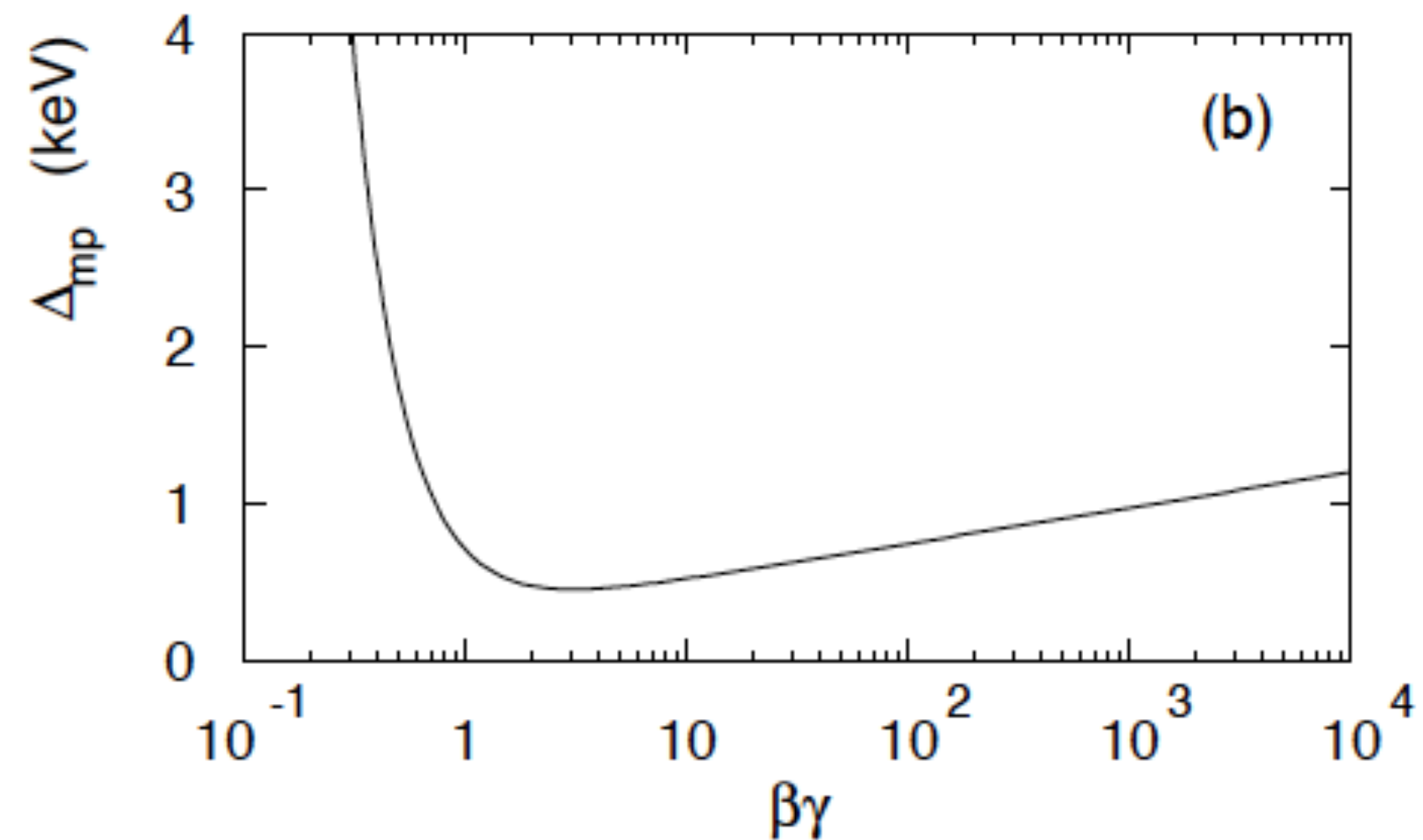
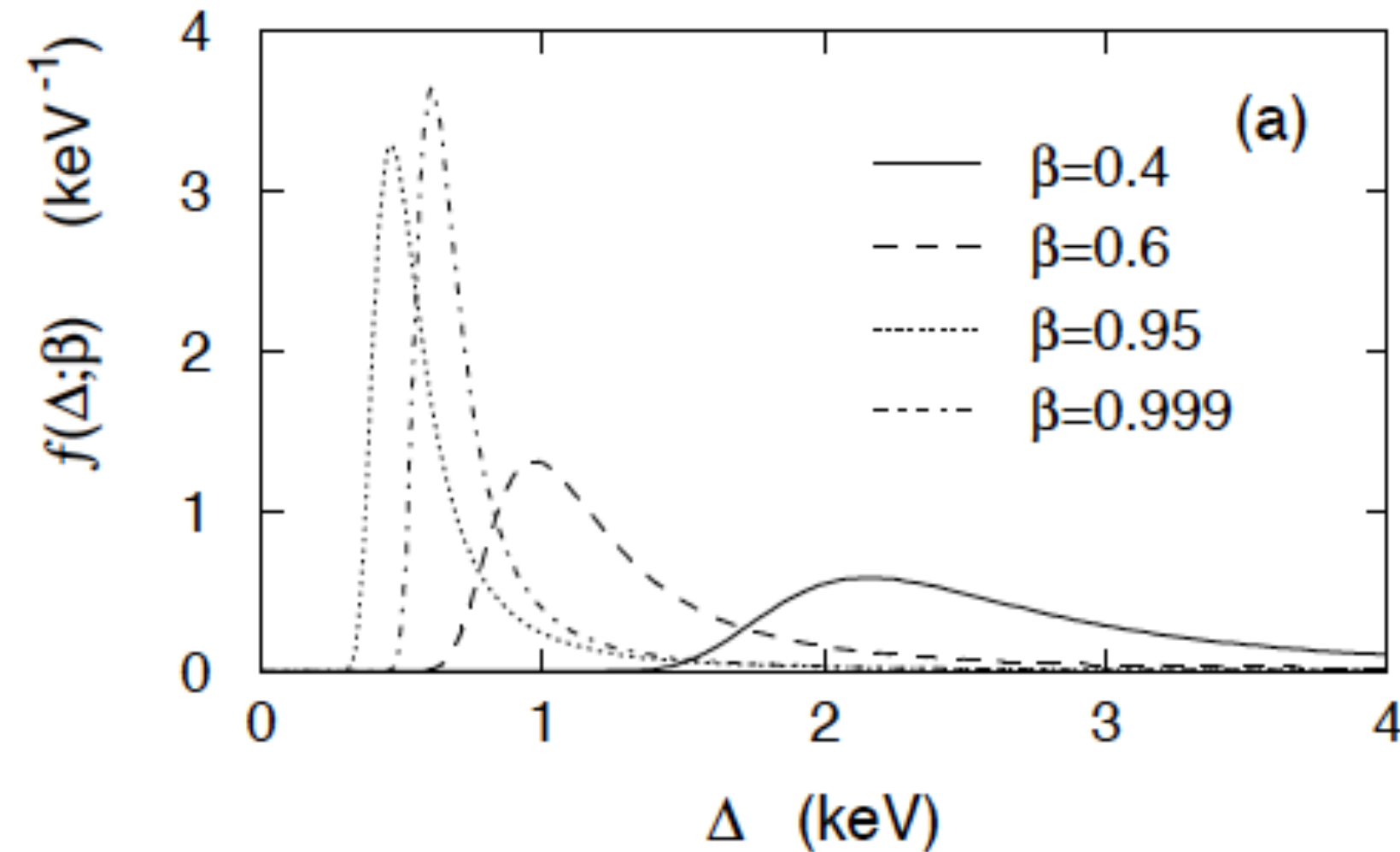
Landau distribution

Long ‘Landau tail’

→ all moments ∞

Mode (most probable value) sensitive to β ,

→ particle i.d.



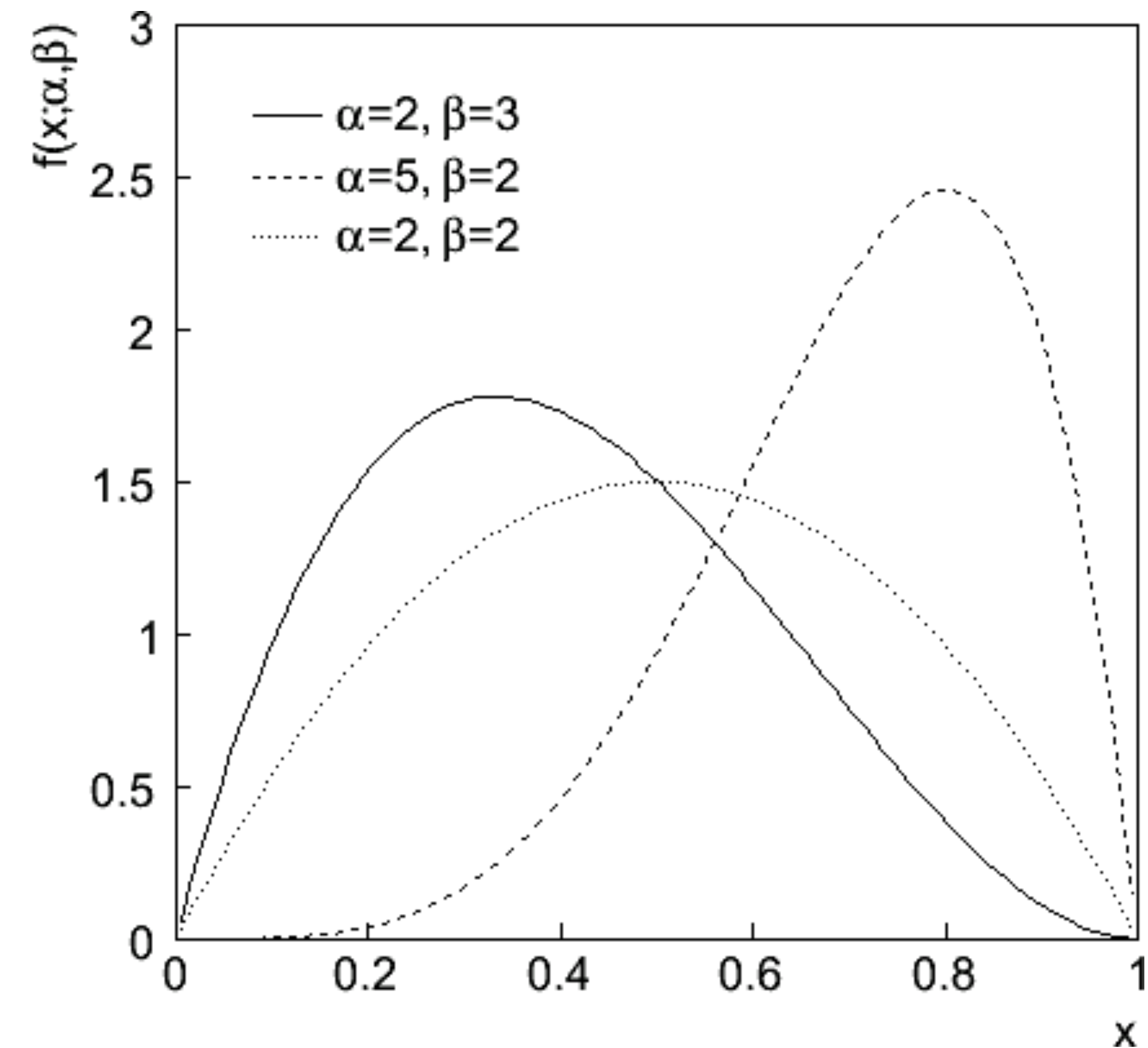
Beta distribution

$$f(x; \alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

$$E[x] = \frac{\alpha}{\alpha + \beta}$$

$$V[x] = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$$

- Often used to represent pdf of continuous non-zero-only random variable between finite limits



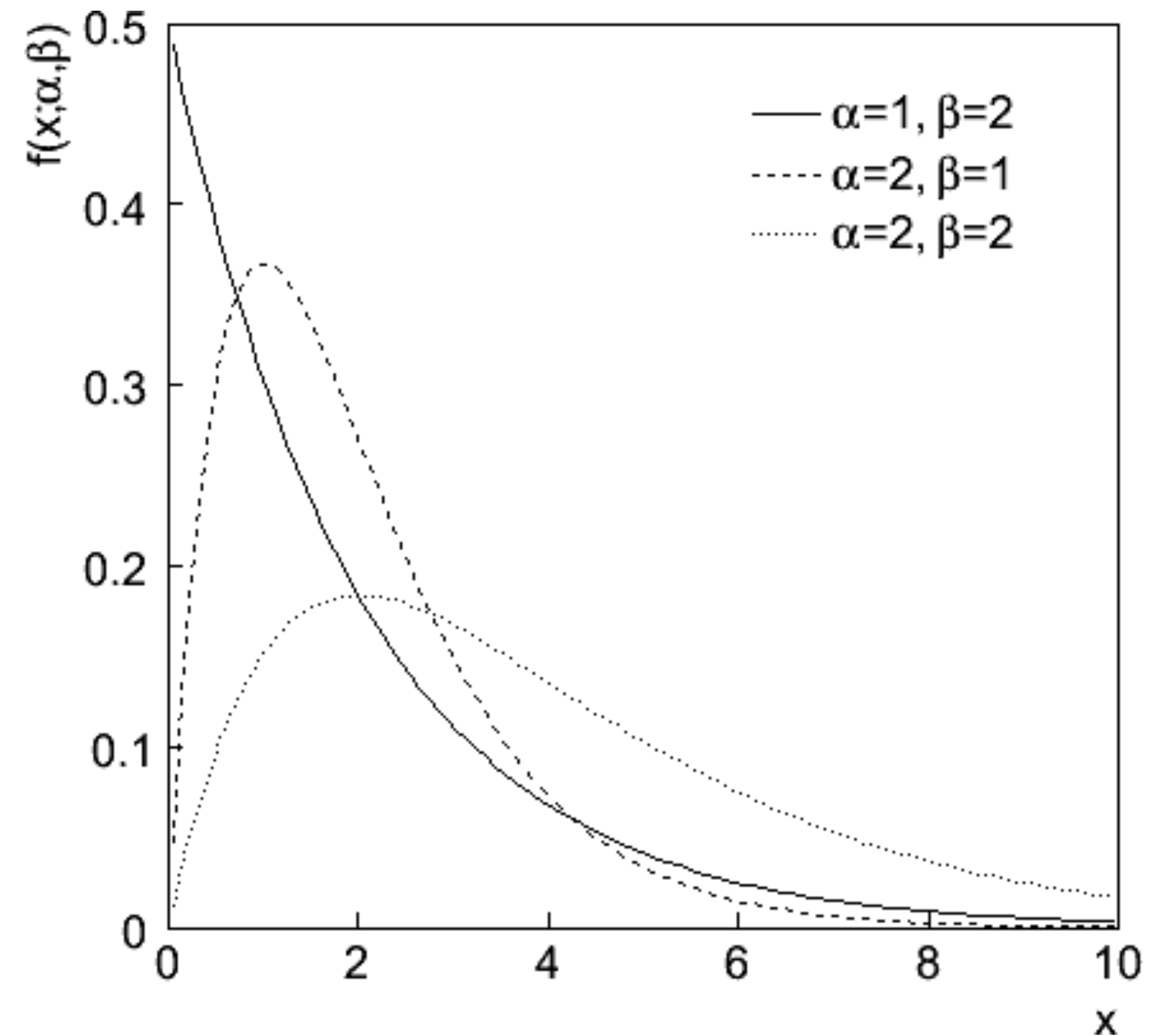
Gamma distribution

$$f(x; \alpha, \beta) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta}$$

$$E[x] = \alpha\beta$$

$$V[x] = \alpha\beta^2$$

- Often used to represent pdf of continuous r.v. nonzero only in $[0, \infty]$.
- Also e.g. sum of n exponential random variables or time until n th event in Poisson process \sim Gamma



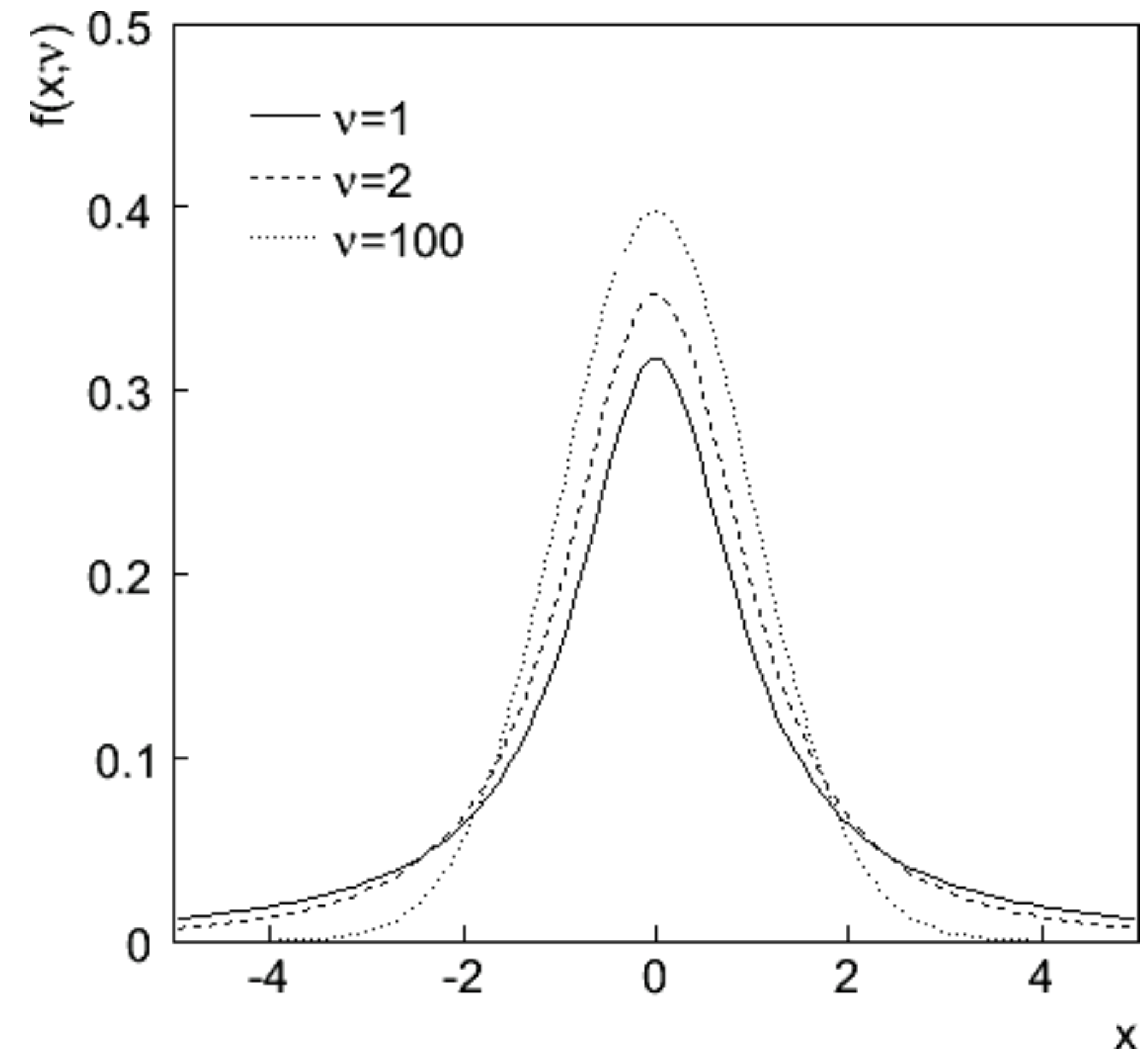
Student's t distribution

$$f(x; \nu) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\nu\pi} \Gamma(\nu/2)} \left(1 + \frac{x^2}{\nu}\right)^{-\left(\frac{\nu+1}{2}\right)}$$

$$E[x] = 0 \quad (\nu > 1)$$

$$V[x] = \frac{\nu}{\nu-2} \quad (\nu > 2)$$

- ν = number of degrees of freedom (not necessarily integer)
- $\nu = 1$ gives Cauchy,
- $\nu \rightarrow \infty$ gives Gaussian.



Student's t distribution

- If $x \sim$ Gaussian with $\mu = 0$, $\sigma^2 = 1$, and $z \sim \chi^2$ with n degrees of freedom, then $t = x / (z/n)^{1/2}$ follows Student's t with $\nu = n$.
- This arises in problems where one forms the ratio of a sample mean to the sample standard deviation of Gaussian random variables.
- The Student's t provides a bell-shaped pdf with adjustable tails, ranging from those of a Gaussian, which fall off very
- quickly, ($\nu \rightarrow \infty$, but in fact already very Gauss-like for $\nu =$ two dozen), to the very long-tailed Cauchy ($\nu = 1$).

Monte Carlo event generators

- Simple example: $e^+e^- \rightarrow \mu^+\mu^-$

- Generate $\cos \theta$ and ϕ

$$f(\cos \theta; A_{\text{FB}}) \propto (1 + \frac{8}{3}A_{\text{FB}} \cos \theta + \cos^2 \theta) ,$$

$$g(\phi) = \frac{1}{2\pi} \quad (0 \leq \phi \leq 2\pi)$$

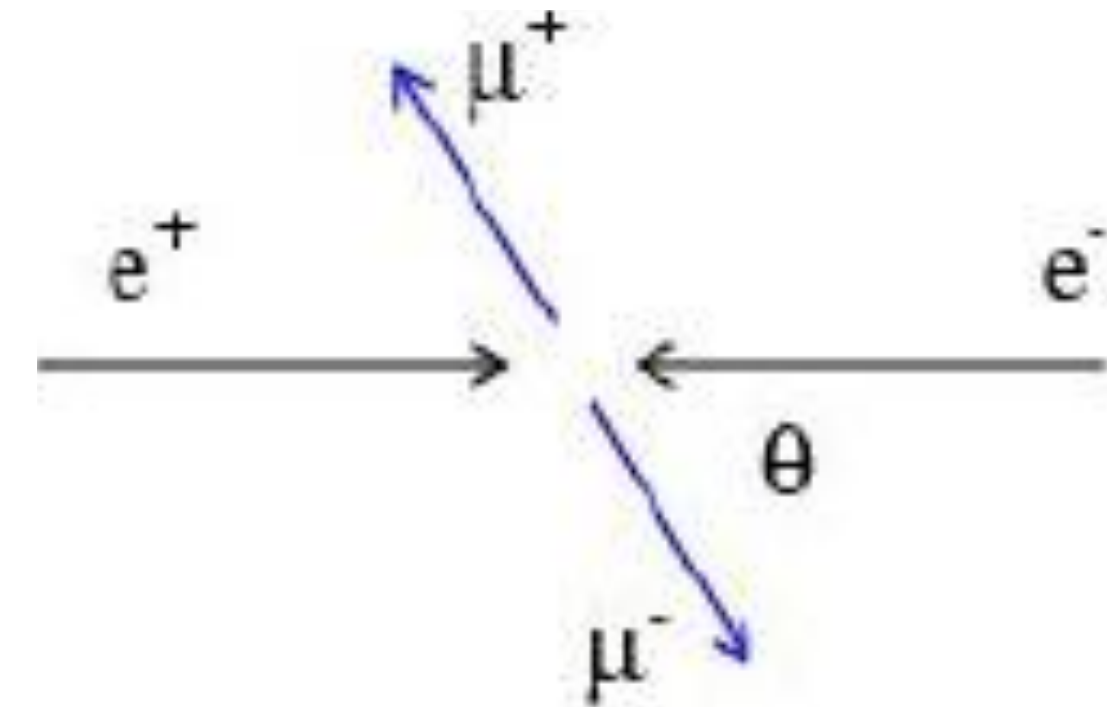
- Less simple: “event generators” for a variety of reactions:

$$e^+e^- \rightarrow \mu^+\mu^-, \text{ hadrons, ...}$$

$$pp \rightarrow \text{hadrons, D-Y, SUSY, ...}$$

- e.g. PYTHIA, POWHEG, HERWIG, ISAJET ...

- Output = “events”, and for each event we get a list of generated particles and their momentum vectors, types, etc..



Example of a simulated event

Event listing (summary)

| I | particle/jet | KS | KF | orig | p_x | p_y | p_z | E | m |
|-------|--------------|----|----------|------|----------|----------|-----------|----------|-------|
| 1 | !p+ | 21 | 2212 | 0 | 0,000 | 0,000 | 7000,000 | 7000,000 | 0,938 |
| 2 | !p+ | 21 | 2212 | 0 | 0,000 | 0,000 | -7000,000 | 7000,000 | 0,938 |
| ===== | | | | | | | | | |
| 3 | !g! | 21 | 21 | 1 | 0,863 | -0,323 | 1739,862 | 1739,862 | 0,000 |
| 4 | !ubar! | 21 | -2 | 2 | -0,621 | -0,163 | -777,415 | 777,415 | 0,000 |
| 5 | !g! | 21 | 21 | 3 | -2,427 | 5,486 | 1487,857 | 1487,857 | 0,000 |
| 6 | !g! | 21 | 21 | 4 | -62,910 | 63,357 | -463,274 | 471,000 | 0,000 |
| 7 | !~g! | 21 | 1000021 | 0 | 314,363 | 544,843 | 498,897 | 979,000 | 0,000 |
| 8 | !~g! | 21 | 1000021 | 0 | -379,700 | -476,000 | 525,686 | 980,000 | 0,000 |
| 9 | !~chi_1- | 21 | -1000024 | 7 | 130,058 | 112,247 | 129,860 | 263,000 | 0,000 |
| 10 | !sbar! | 21 | -3 | 7 | 259,400 | 187,468 | 83,100 | 330,000 | 0,000 |
| 11 | !c! | 21 | 4 | 7 | -79,403 | 242,409 | 283,026 | 381,000 | 0,000 |
| 12 | !~chi_20! | 21 | 1000023 | 8 | -326,241 | -80,971 | 113,712 | 385,000 | 0,000 |
| 13 | !b! | 21 | 5 | 8 | -51,841 | -294,077 | 389,853 | 491,000 | 0,000 |
| 14 | !bbar! | 21 | -5 | 8 | -0,597 | -99,577 | 21,299 | 101,000 | 0,000 |
| 15 | !~chi_10! | 21 | 1000022 | 9 | 103,352 | 81,316 | 83,457 | 175,000 | 0,000 |
| 16 | !s! | 21 | 3 | 9 | 5,451 | 38,374 | 52,302 | 65,000 | 0,000 |
| 17 | !cbar! | 21 | -4 | 9 | 20,839 | -7,250 | -5,938 | 22,000 | 0,000 |
| 18 | !~chi_10! | 21 | 1000022 | 12 | -136,266 | -72,961 | 53,246 | 181,000 | 0,000 |
| 19 | !nu_mu! | 21 | 14 | 12 | -78,263 | -24,757 | 21,719 | 84,000 | 0,000 |
| 20 | !nu_mubar! | 21 | -14 | 12 | -107,801 | 16,901 | 38,226 | 115,000 | 0,000 |
| ===== | | | | | | | | | |
| 21 | gamma | 1 | 22 | 4 | 2,636 | 1,357 | 0,125 | 2,000 | 0,000 |
| 22 | (~chi_1-) | 11 | -1000024 | 9 | 129,643 | 112,440 | 129,820 | 262,000 | 0,000 |
| 23 | (~chi_20) | 11 | 1000023 | 12 | -322,330 | -80,817 | 113,191 | 382,000 | 0,000 |
| 24 | (~chi_10) | 1 | 1000022 | 15 | 97,944 | 77,819 | 80,917 | 169,000 | 0,000 |
| 25 | (~chi_10) | 1 | 1000022 | 18 | -136,266 | -72,961 | 53,246 | 181,000 | 0,000 |
| 26 | nu_mu | 1 | 14 | 19 | -78,263 | -24,757 | 21,719 | 84,000 | 0,000 |
| 27 | nu_mubar | 1 | -14 | 20 | -107,801 | 16,901 | 38,226 | 115,000 | 0,000 |
| 28 | (Delta++) | 11 | 2224 | 2 | 0,222 | 0,012 | -2734,287 | 2734,287 | 0,000 |

A simulated event

⋮

| | | | | | | | | | |
|-----|---------|----|-------|-----|--------|--------|-----------|----------|-------|
| 397 | pi+ | 1 | 211 | 209 | 0,006 | 0,398 | -308,296 | 308,297 | 0,140 |
| 398 | gamma | 1 | 22 | 211 | 0,407 | 0,087 | -1695,458 | 1695,458 | 0,000 |
| 399 | gamma | 1 | 22 | 211 | 0,113 | -0,029 | -314,822 | 314,822 | 0,000 |
| 400 | (pi0) | 11 | 111 | 212 | 0,021 | 0,122 | -103,709 | 103,709 | 0,135 |
| 401 | (pi0) | 11 | 111 | 212 | 0,084 | -0,068 | -94,276 | 94,276 | 0,135 |
| 402 | (pi0) | 11 | 111 | 212 | 0,267 | -0,052 | -144,673 | 144,674 | 0,135 |
| 403 | gamma | 1 | 22 | 215 | -1,581 | 2,473 | 3,306 | 4,421 | 0,000 |
| 404 | gamma | 1 | 22 | 215 | -1,494 | 2,143 | 3,051 | 4,016 | 0,000 |
| 405 | pi- | 1 | -211 | 216 | 0,007 | 0,738 | 4,015 | 4,085 | 0,140 |
| 406 | pi+ | 1 | 211 | 216 | -0,024 | 0,293 | 0,486 | 0,585 | 0,140 |
| 407 | K+ | 1 | 321 | 218 | 4,382 | -1,412 | -1,799 | 4,968 | 0,494 |
| 408 | pi- | 1 | -211 | 218 | 1,183 | -0,894 | -0,176 | 1,500 | 0,140 |
| 409 | (pi0) | 11 | 111 | 218 | 0,955 | -0,459 | -0,590 | 1,221 | 0,135 |
| 410 | (pi0) | 11 | 111 | 218 | 2,349 | -1,105 | -1,181 | 2,855 | 0,135 |
| 411 | (Kbar0) | 11 | -311 | 219 | 1,441 | -0,247 | -0,472 | 1,615 | 0,498 |
| 412 | pi- | 1 | -211 | 219 | 2,232 | -0,400 | -0,249 | 2,285 | 0,140 |
| 413 | K+ | 1 | 321 | 220 | 1,380 | -0,652 | -0,361 | 1,644 | 0,494 |
| 414 | (pi0) | 11 | 111 | 220 | 1,078 | -0,265 | 0,175 | 1,132 | 0,135 |
| 415 | (K_S0) | 11 | 310 | 222 | 1,841 | 0,111 | 0,894 | 2,109 | 0,498 |
| 416 | K+ | 1 | 321 | 223 | 0,307 | 0,107 | 0,252 | 0,642 | 0,494 |
| 417 | pi- | 1 | -211 | 223 | 0,266 | 0,316 | -0,201 | 0,480 | 0,140 |
| 418 | nbar0 | 1 | -2112 | 226 | 1,335 | 1,641 | 2,078 | 3,111 | 0,940 |
| 419 | (pi0) | 11 | 111 | 226 | 0,899 | 1,046 | 1,311 | 1,908 | 0,135 |
| 420 | pi+ | 1 | 211 | 227 | 0,217 | 1,407 | 1,356 | 1,971 | 0,140 |
| 421 | (pi0) | 11 | 111 | 227 | 1,207 | 2,336 | 2,767 | 3,820 | 0,135 |
| 422 | n0 | 1 | 2112 | 228 | 3,475 | 5,324 | 5,702 | 8,592 | 0,940 |
| 423 | pi- | 1 | -211 | 228 | 1,856 | 2,606 | 2,808 | 4,259 | 0,140 |
| 424 | gamma | 1 | 22 | 229 | -0,012 | 0,247 | 0,421 | 0,489 | 0,000 |
| 425 | gamma | 1 | 22 | 229 | 0,025 | 0,034 | 0,009 | 0,043 | 0,000 |
| 426 | pi+ | 1 | 211 | 230 | 2,718 | 5,229 | 6,403 | 8,703 | 0,140 |
| 427 | (pi0) | 11 | 111 | 230 | 4,109 | 6,747 | 7,597 | 10,961 | 0,135 |
| 428 | pi- | 1 | -211 | 231 | 0,551 | 1,233 | 1,945 | 2,372 | 0,140 |
| 429 | (pi0) | 11 | 111 | 231 | 0,645 | 1,141 | 0,922 | 1,608 | 0,135 |
| 430 | gamma | 1 | 22 | 232 | -0,383 | 1,169 | 1,208 | 1,724 | 0,000 |
| 431 | gamma | 1 | 22 | 232 | -0,201 | 0,070 | 0,060 | 0,221 | 0,000 |

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PYTHIA Monte Carlo
pp → gluino-gluino

Monte Carlo detector simulation

- Takes as input the particle list and momenta from generator.
- Simulates detector response:
 - multiple Coulomb scattering (generate scattering angle), particle decays (generate lifetime), ionization energy loss (generate Δ), electromagnetic, hadronic showers, production of signals, electronics response, ...
- Output = simulated raw data -> input to reconstruction software: track finding, fitting, shower clustering, etc.
- Predict what you should see at “detector level” of events generated by the “event generator”. Compare with “real data”, e.g. can be used to estimate “efficiency” = $[N \text{ events found}] / [N \text{ events generated}]$
- Software: GEANT4